# Crop Identification and Area Estimation over Large Geographic Areas Using LANDSAT MSS Data

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Laboratory for Applications of Remote Sensing
Purdue University
W. Lafayette, Indiana 47906

January 1977
Final Report for the Period of
March 1975 to September 1976

Prepared for GODDARD SPACE FLIGHT CENTER Greenbelt, Maryland 20771

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## CROP IDENTIFICATION AND AREA ESTIMATION OVER LARGE GEOGRAPHIC AREAS USING LANDSAT MSS DATA

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Laboratory for Applications of Remote Sensing

Purdue University

West Lafayette, Indiana 47906'

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#### 16. Abstract

This report describes the results of a study involving the use of computer-aided analysis techniques applied to Landsat MSS data for identification and area estimation of winter wheat in Kansas and corn and soybeans in Indiana. Key elements of the approach included use of aerial photography for classifier training, stratification of Landsat data and extension of training statistics to areas without training data, and classification of a systematic sample of pixels from each county. Major results and conclusions are that (1) Landsat data was\_adequate to accurately identify winter wheat in Kansas, but not corn and soybeans in Indiana; (2) computer-aided analysis techniques can be effectively used to extract crop identification information from Landsat MSS data, and (3) systematic sampling of entire counties made possible by computer classification methods resulted in very precise area estimates at county as well as district and state levels.

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#### PREFACE

This investigation applied to Landsat data the advances and developments of the past decade in analyzing multispectral remote sensing measurements for crop identification and area estimation. Landsat MSS data for Kansas and Indiana were classified using computer-aided analysis techniques to identify and determine the areal extent and distribution of the major crops in the two state test area. It was conclusively demonstrated that Landsat data analyzed by computer methods could be effectively used to produce accurate estimates having extremely small sampling error. Recommendations are made for increasing the spectral, spatial and temporal resolution of data acquired by future satellite systems, along with preprocessing to geometrically correct and register data sets. It is recommended that attention be given to developing more effective methods of scene stratification and obtaining crop yield information from Landsat data.

The rationale and background of the investigation are described in Section 1.0; the objectives follow in Section 2.0. In Sections 3.0 and 4.0 the test areas and experimental

approach and procedures are described. The results of the investigation are presented in Sections 5.0 and 6.0. The significant results and conclusions of the investigation are given in Section 7.0, followed by the recommendations in Section 8.0.

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#### 1.0 INTRODUCTION

As our grain reserves become depleted and world population and demand for food increase, the need to improve the quality of world crop production information becomes ever more critical. Accurate and timely crop production information has been identified at the World Food Conference held in Rome in 1974 [25] and more recently in a National Academy of Science study [20] as a critical part of the solution of the food problem.

During the past decade considerable evidence has developed that multispectral remote sensing from aerospace platforms can provide quantitative data which can be effectively used to identify major crop species and determine their areal extent. Remote sensing techniques may prove to be a more accurate, precise, timely, and/or cost effective method of acquiring crop production information than conventional surveys carried out on the ground. The information gained from this investigation should provide additional data on which to determine the utility of remote sensing.

#### 1.1 Value of Crop Production Information

Most countries forecast and estimate their crop production, but relatively few have reliable methods for gathering the necessary data. The benefits of improved crop information are: (1) accurate estimates result in price stability; (2) timely and accurate forecasts of production allow governments to plan domestic and foreign policies and actions; and (3) accurate forecasts enable optimal utilization of storage, transportation, and processing facilities. Conversely, the socioeconomic costs of not having accurate and timely information available are substantial.

The economic value of increased crop forecast accuracy in the United States was first quantified by Hayami and Peterson [12]. They estimated from their model that a reduction in forecast error for wheat from 3.2% to 2.1% would have annual net social benefits of 70 million dollars at 1968 prices—a figure which would be approximately doubled at 1974-1976 prices. On a world basis the value of improved forecast would be substantially greater. Comparable benefits would be gained by improving the accuracy of estimates for other major crops.

In addition, more frequent information, such as might be provided with remote sensing techniques, would increase the social benefits even without improvements in the crop estimate error [10].

#### 1.2 Conventional Crop Survey Methods

Information gathering is as old as civilization. Census taking by the Egyptian Pharaohs and Roman Emperors are good examples. However, the application of scientific statistical methods to gathering agricultural statistics is only about a hundred years old. But, in spite of many technological advances in the methods used to survey crops, many countries still do not have adequate systems to gather data needed to support satisfactory decision making about food and nutrition.

The system developed in the United States is regarded as being one of the most comprehensive and accurate. In this country the Statistical Reporting Service of the Department of Agriculture (USDA/SRS) has responsibility for collecting and reporting current data on U.S. agriculture. The present program of crop and livestock estimation annually includes over 500 national reports, plus numerous reports issued by individual states. Reports are made for more than 120 crop commodities (including field and seed crops, vegetables, fruits, and nuts) and provide estimates of acreages farmers intend to plant; acreages actually planted and harvested; yield, production and crop disposition; as well as periodic indications of remaining stocks for important crops. Monthly forecasts of production are prepared for major crops throughout the growing season.

Nearly all surveys conducted by SRS are probability surveys based on relatively small samples. Since 1965 a national general purpose survey including 17,000 area segments which are enumerated during May and June each year has been used. The sampling units or area segments are typically about 2.6 square kilometers (about one square mile) in size. This sample is stratified with states and areas within states serving as strata. Crop reporting districts(CRD), groupings of contiguous counties having similar agricultural practices, are generally the intrastate strata. Sample selection within strata follows a systematic approach using a geographically arranged listing of the sampling frame. Trained enumerators visit each segment and interview each farm operator to obtain data on crop acreages, livestock production, production costs, and prices received. About 20% of the questionnaire concerns crop acreage information. Additional information describing the SRS sampling and estimation procedures may be found in references [23] and [7].

The current SRS probability surveys provide independent estimates with known measures of precision (sampling errors). Typical sampling errors for several major crops are shown in Table 1. It should be noted here the SRS surveys are designed to produce accurate, precise estimates at the national level. At the state level where there are generally 300-400 sampling units, the sampling error is

Table 1. Coefficients of variation from June Enumerative
• and Objective Yield Surveys in the United States,
1975.

### Coefficient of Variation (%)

Crop	Acres Planted	Yield	Production
Winter Wheat	1.5	1.0	2.0
Corn .	1.1	0.9	1.7
Soybeans	3.5	1.0	2.1
Cotton	3.5	1.0	3.7

<sup>&</sup>lt;sup>a</sup>From Caudill [7].

greater; coefficients of variation are typically 4-6%.  $_{\text{N}_{2}}$ 

Estimates for counties are not obtained from the June enumerative survey since there are too few segments per county to be reliable. Rather, the estimate of the total acreage of, for example, wheat in the state is obtained and then subdivided among counties. The county allocations are based on a mail survey which may include 50-100 respondents per county and/or the last agricultural census. Variance estimates are not calculated by the SRS for county estimates, but the coefficients of variation are believed to be on the order of 10% or more.

#### 1.3 Development of Remote Sensing Technology for Crop Surveys

To understand the approach used and results from this investigation it\_will be helpful to briefly review the development of remote sensing technology related to crop surveys. This historical perspective will indicate the progress which has been made and the contribution of this investigation.

Remote sensing from satellites is particularly appropriate for crop surveys because of the capability to obtain repetitive coverage of wide areas. The physical basis for remote sensing, data acquisition platforms and sensors, and data analysis techniques are described by Bauer [3] in a review of the potential role of remote sensing in determining the distribution and yield of crops.

Remote sensing as it is known today is an outgrowth of aerial photography. Although the use of aerial photography has been developing for more than a hundred years, remote sensing has been evolving and expanding most rapidly since 1960 as new sensors and interpretation techniques became available.

In 1964, multispectral photography was collected for the first time over agricultural fields, and the potential of the multispectral approach to crop identification was recognized [13]. After this approach was further defined, a crop classification was made from multispectral scanner data in 1967, using pattern recognition methods implemented on a digital computer [17].

One of the first investigations using satellite-acquired imagery to identify crops was performed by Anuta and MacDonald [2]. Apollo-9 multispectral photography was digitized and analyzed using computer-implemented pattern recognition techniques. Agricultural land in the Imperial Valley of California was accurately classified into several individual crops, soil, and water.

The Corn Blight Watch Experiment, conducted in 1971 by NASA, USDA, Purdue University, and the University of Michigan in seven Corn Belt states, provided a prototype remote sensing system [18]. It successfully integrated techniques of sampling, data acquisition, storage, retrieval, processing, analysis, and information dissemination in a quasi-operational

system environment. The results showed that remote sensing could be used to quantitatively recognize corn leaf blight, as well as other agricultural crops and land uses over broad areas.

The supply of remotely sensed data greatly increased with the launch of Landsat-1 (formerly called the Earth Resources Technology Satellite or ERTS-1) in 1972. From an orbit 912 km above the earth, the satellite can complete a full observation of the earth every 18 days. Its multispectral imagery is collected in four visible and infrared wavelength bands over 185 km wide passes over the earth. This newest data source with its synoptic view of earth has opened a whole new dimension to the capability to obtain information about earth resources.

Bauer and Cipra [4] used multivariate pattern recognition methods implemented on a digital computer to classify Landsat-1 data acquired over a three-county area in northern Illinois. The classification of the Landsat data, as measured by an independent sample of test fields, was 85% accurate on a point by point basis (Table 2). Although there were errors in the classification of individual data points, area estimates made over the three-county area were within a few percent of those made by the U.S. Department of Agriculture (Table 3).

Table 2. Classification of corn, soybean, and "other" test fields by computer-aided analysis of Landsat-1 multispectral scanner data for DeKalb County, Illinois.

	Number	Number o	Number of points classified as		Percent
Crop	of points	Corn	Soybeans	"Other"	correctly classified
Corn	3968	3367	357	244 .	85
Soybeans	1113	115	855	133	77
"Other"	295	16	50	234	79
	5376	3498	1262	611	83

<sup>&</sup>lt;sup>a</sup>From Bauer and Cipra [ 4 ].

Table 3. Comparison of area estimates made by U.S. Department of Agriculture and from classification of Landsat-1 multispectral scanner data for DeKalb, Ogle, and Lee Counties, Illinois.

	Percent of total area		
Crop	USDA	LANDSAT	
Corn	40.2	39.6	
Soybeans	18.0	17.8	
Other ·	41.8	42.6	

<sup>&</sup>lt;sup>a</sup>From Bauer and Cipra [ 4 ].

#### 2.0 OBJECTIVES

The long term objective of this investigation is to develop and test procedures utilizing Landsat data to not only identify, but more importantly, determine the areal extent and distribution of earth surface features over large geographic areas. The specific applications selected for this investigation are crop identification and area estimation for two states in the Central United States.

There is high probability that improved crop production information, long recognized as a potential application of remote sensing, can be obtained from Landsat data. The wide area coverage of—Landsat, linked with computer processing, offers a unique opportunity to improve upon the sampling methods now used for making area estimates from ground-based systems. This is particularly true as the size of the area decreases, e.g. state, district, county. Further, the sequential coverage of Landsat should lead to improvements in the timeliness of the estimates. Both of these aspects would result in economic and social benefits.

The specific objectives of this study are:

IJ

- Using Landsat data and computer-implemented pattern recognition, classify the major crops from regions encompassing different climates, soils, and crops.
  - Estimate crop areas for county and state size areas using the crop identification data obtained from the Landsat classifications.
  - Evaluate the accuracy, precision, and timeliness of crop area estimates obtained from Landsat data.

Two important underlying premises to be tested in the investigation are:

- The synoptic view of Landsat provides the opportunity to obtain crop production information over large areas, e.g. states and countries.
- By using computer-implemented data analysis to classify pixels distributed over entire counties, it is also possible to make accurate and precise estimates for local areas, e.g. counties.

The successful accomplishment of the investigation would contribute to the development of earth resources surveys by:

- Leading to operational use of satellite data for obtaining crop area estimates.
- Refining techniques which could also be applied to other problems such as crop yield forecasts, natural resource inventories, and measurement and monitoring of damage caused by floods, drought, insects and disease.
- Developing improved methods of obtaining necessary ground truth.
- Testing statistical sampling models designed specifically for remote sensing applications.
- Providing data for determining needed information on costs and benefits of obtaining information using remote sensing.

#### 3.0 SELECTION AND DESCRIPTION OF TEST AREAS AND CROPS

Kansas and Indiana were selected as the test states for this investigation. Winter wheat in Kansas and corn and soybeans in Indiana were selected as the crops for which area estimates would be made from classifications of Landsat data.

The test areas and crops were selected to sample the range of conditions which are present in the Great Plains and Corn Belt regions of the United States. The selections of test areas and crops were made taking into account the spectral and spatial parameters of the Landsat data and the characteristics of crop production. On the "spectrum of difficulty", wheat identification in Kansas is undoubtedly an easier problem than corn and soybean identification in Indiana. That is, the Landsat data is likely to be more adequate for winter wheat identification in Kansas than for corn and soybean identification in Indiana.

Winter wheat is the first crop to "green-up" in the spring, has the greatest amount of green biomass (except for alfalfa) during the April to mid-June period, and at maturity in late June and early July is the only cover type which is golden-yellow in color. In other words, during much of its

growth cycle it is dissimilar from the other cover types present. Additional factors simplifying the task of wheat identification and area estimation in Kansas is that wheat is grown in relatively large fields, on a large percentage of the agricultural land, and with relatively few other cover types and crops present.

In comparison, corn and soybeans in Indiana are warm season or summer crops which are green at the same time as many other cover types present during the summer in Indiana. Some of the possible "confusion" cover types include trees, pasture, forage crops, and oats. Secondly, field sizes in Indiana are much smaller than in Kansas. This is due to the greater heteorogeneity in soils and the greater number of crops being grown. The smaller field sizes cause a greater fraction of pixels to fall on field boundaries and include more than one cover type. In summary, corn and soybeans in Indiana are more like the classes they are to be discriminated from than is the case with winter wheat in Kansas.

Kansas is the number one wheat producing state in the nation [16]. Its wheat production for 1975 totaled 9.6 million metric tons (351 million bushels), 10% above 1974 and second only to the record 10.5 million metric tons (385 million bushels) produced in 1973. The 1975 crop was seeded on 5.2 million hectares (12.8 million acres), 7% more than a year earlier. Area harvested for grain, at 4.9 million hectares (12.1 million acres), was 4% above the previous year.

Abandonment, at 5.5%, was slightly above recent years but well within normal rates of abandonment. The average yield of 19.5 quintals per hectare (29 bushels per harvested acre) was 1.0 quintal (1.5 bushels) above the 18.5 quintal (27.5 bushel) average in 1974. The distribution of wheat production in the state is shown in Figure 1. The farm value of the 1975 wheat crop in Kansas was 1.2 billion dollars.

Kansas soils were developed under mixed or short prairie grass vegetation. Average precipitation varies from 38 centimeters (15 inches) in the west to 81 centimeters (32 inches) in the east. The climate is continental in most of the state, becoming semi-arid in the west. The distribution and amount of precipitation during the year fit the requirements of winter wheat better than any other crop in much of the state. Other important crops grown include corn, grain sorghum, and alfalfa. The amount of irrigated land is increasing each year. There were 20.2-million hectares (49.9 million acres) of land in farms in 1975; crops were harvested from 12 million hectares (30 million acres).

In 1975 Indiana ranked third among the states in both corn and soybean production [15]. The 2.3 million hectares (5.6 million acres) of corn harvested was a record high. The average corn yield was 59 quintals per hectare (98 bushels per acre). Production at 13.5 million metric tons (552 million bushels) was the second largest crop on record. The area in

WHEAT—Bushels Produced by Counties—1975
Rank of First Ten Counties Shown by Number Within County

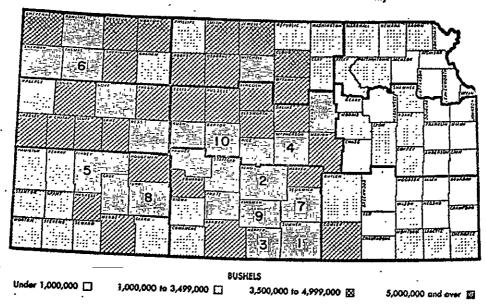


Figure 1. The distribution of 1975 wheat production in Kansas.

soybeans was 1.5 million hectares (3.6 million acres), 7% below the previous year. The 20.7 quintal (33 bushel) average yield was a record high and total production of 3.0 million metric tons (120 million bushels) was the second greatest ever. The distributions of Indiana corn and soybeans are shown in Figure 2.

Indiana includes both glacial and non-glacial soils, with topography ranging from the nearly level prairies of northern and central parts of the state to the rolling and steep lands of the southern areas of the state. Both dark colored soils developed under prairie vegetation and light colored soils developed under forest are present. The climate is typically continental with cold winters, warm summers, and frequent short period fluctuations of temperature, humidity, cloudiness, and wind direction. The well-distributed annual precipitation of 81 to 102 centimeters (32 to 40 inches) favors high agricultural production. Sunshine averages more than 70% of its possible duration for the summer months and summer precipitation occurs mostly during short duration showers or thunderstorms.

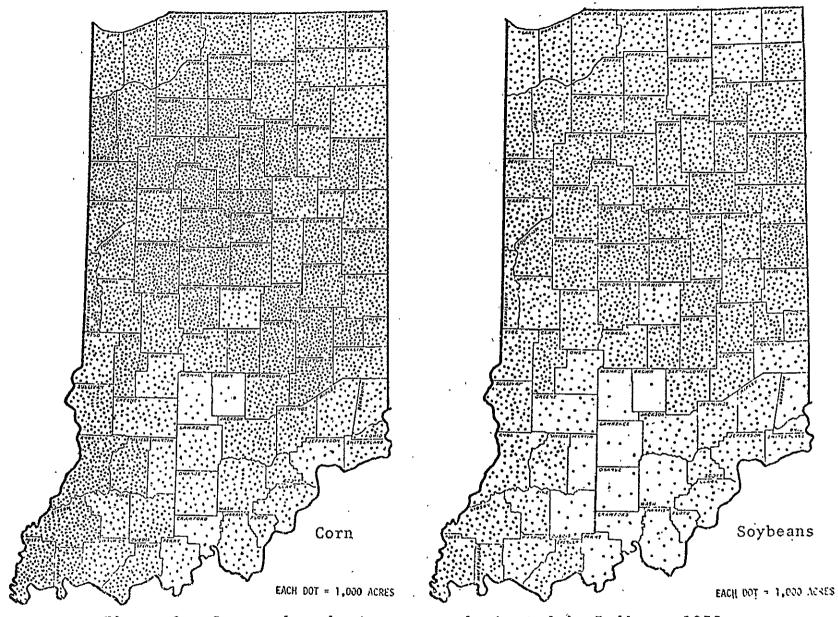


Figure 2. Corn and soybeans acreage harvested in Indiana, 1975.

#### 4.0 EXPERIMENTAL APPROACH AND PROCEDURES

The approach used in the investigation built on procedures developed and utilized in previous research at LARS with the objective of extending them to larger areas. The procedures were developed upon five fundamentals which were determined early in the investigation:

- The classifier would be trained and tested using aerial photography as reference data.
- Counties without reference data would be classified using training statistics from an adjacent county having similar crops and soils and lying in the same Landsat frame.
- Area estimates would be made from a systematic random sample of pixels distributed over the entire\_county.
- Area estimates would be made on a county basis and aggregated to district and state levels.
- Estimates would be adjusted for classification bias.

The implementation of the basic steps is illustrated in Figure 3. The remainder of this section describes in detail the procedures used in the investigation.

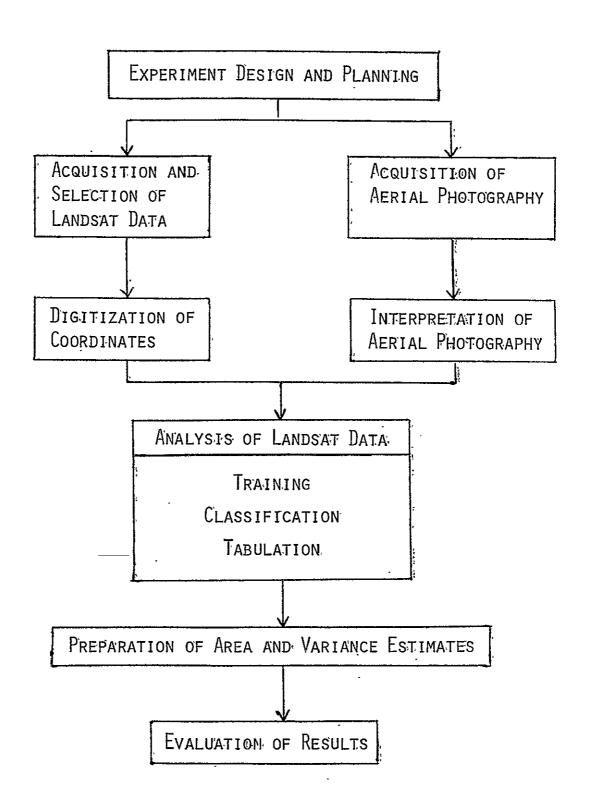


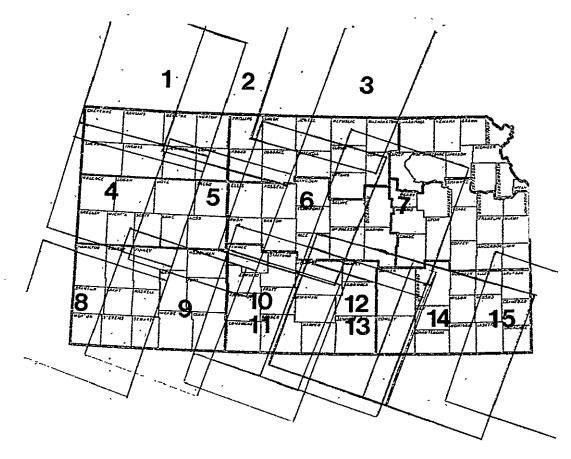
Figure 3. Implementation of experimental approach.

#### 4.1 Acquisition and Selection of Landsat Data

At the beginning of the project a standing order was placed with the EROS Data Center for Landsat-2 photographic imagery over Kansas and Indiana. The imagery was the basis for decisions of the choice of scenes to be used for classification. If a scene was chosen for use, the bulk computer compatible tape was then ordered retrospectively. Landsat-2 was the primary source of multispectral scanner (MSS) data, with Landsat-1 scenes being used only to complete the coverage for the Southwestern Crop Reporting District (CRD) in Kansas.

The selection of a Landsat frame to classify for a given county was based upon the date of the Landsat data, the location of ground truth, and the amount and location of cloud cover. The desired attributes were that the crops of interest were spectrally discriminable at the time of the Landsat pass; aerial photography was available over areas similar in crop stage and soils in the same frame; and both the county to be classified and the training areas were not obscured by clouds or bad data.

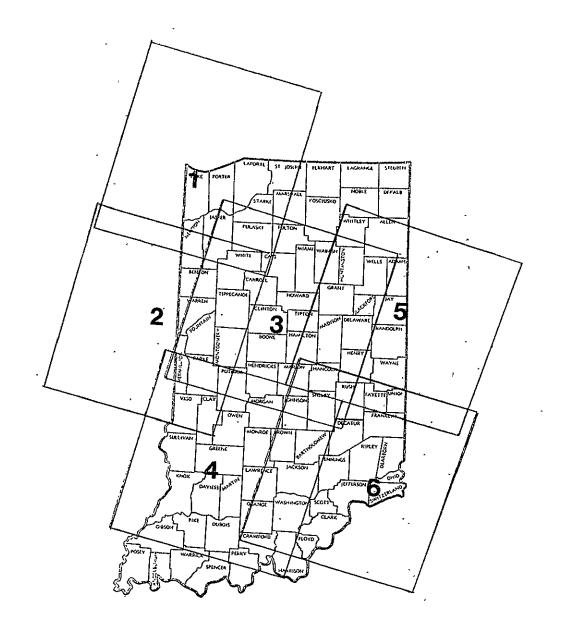
The Landsat frames chosen for the analysis in Kansas and Indiana are shown in Figures 4 and 5, respectively. The amount of cloud cover created a serious problem for obtaining data for much of Indiana and northeastern Kansas. As a result, satisfactory data was not available for the Northeast



Key

	Landsa't Scene ID	LARS Run Number	Date
1	2165-16450	75013800	July 6
2	2146-16392	75005800	June 17
3	2163-16334	75006500	July 4
4	2165-16453	75004600	July 6
5	2146-16395	75005900	June 17
6	2163-16340.	75006600	
7	2144-16282	75005600	July 4 June 15
8	2147-16460	75006200	June 18
9	5032-16310	75007200	May 21
10	2073-16342	75001500	April 15
11	2109-16341	75005000	May 11
12	2072-16284	75000900	April 9
13	2144-16284	75005700	June 15
14.	2107-16225	75004900	May 9
15	2142-16171	75005400	June 13

Figure 4. Landsat Coverage for Kansas.



Key

	Landsat Scene ID	LARS Run Number	Date
1	2228-15515	75009100	September 7
2	2228-15522	75009200	September 7
3	2209-15464	75009000	August 19
4	2173-15480	75008700	July 14
5	2208-15405	75010000	August 18
6	2208-15412	75010100	August 18

Figure 5. Landsat Coverage for Indiana.

and East Central CRDs in Kansas. In Indiana, the only districts that had complete Landsat coverage were the Northwestern, West Central, Central and East Central.

Tables 4 and 5 illustrate the cloud cover problem. The standing order for Landsat-2 photographic imagery requested scenes that contained less than 50% cloud cover. Since a low cloud cover percentage does not necessarily mean that a scene is usable for analysis, the number of usable scenes is specified in Tables 4 and 5. For example, a frame could be half in Indiana and half in Illinois. If the frame has 10-20% cloud cover but the clouds cover the Indiana portion of the frame, it is unusable. Or, if there are three or four large cloud patches which occur as long streaks across the frame, the frame is unusable even though the cloud cover may have only been 20%. The magnitude of the cloud cover problem is indicated in the tallies of data acquired and data used which show that only 21 out of 93 frames in Kansas and only eight out of 40 in Indiana were usable.

In Kansas, there was April data available to cover the entire south central CRD and data in May and June to provide duplicate coverage for ten of the thirteen counties. It was decided to analyze these ten counties twice and compare the results. Figure 4 indicates which counties were analyzed twice and which frames and dates were used. In the statistical analysis of the results for Kansas, both dates, were used for most of the statistical tests. However, the tables

Table 4. Summary of acquisition and usability of Landsat-2 data for Kansas, April 1 - July 17, 1975.

Month	No. Frames Acquired by NASA/GSFC	No. Frames Received from EROS Data Center*	No. Usable Frames
April	29	8	6
May	28	9	2
June	18	15	9
July	18	9	4 —
Total	93	41	21

<sup>\*</sup>Standing order for all frames with < 50% cloud cover.

Table 5. Summary of acquisition and usability of Landsat-2 data for Indiana, July 1 - September 7, 1975.

Month	No. Frames Acquired by NASA/GSFC	No. Frames Received from EROS Data Center*	No. Usable Frames
July	14	11	2
August	16	7	4
September	10	6 <del>-</del>	<u>2</u>
Total	40	24 .	8

<sup>\*</sup>Standing order for all frames with < 50% cloud cover.

in sections 5.2 to 5.3 display figures only for the second date for these ten counties since the second date was closer to the time the wheat was harvested. The estimates made at harvest time are more important since the SRS estimates for area harvested were used for comparison of results.

#### 4.2 Acquisition of Aerial Photography

A critical part of the entire investigation involved the reference or "ground truth" data set to be utilized in conjunction with the computer-aided analysis of the Landsat MSS data. Reference data was required for training the classifier and to test the accuracy of classification. Detailed crop type maps do not exist because the crop grown in an individual field generally changes each year. And, indeed some field boundaries are changed from year to year. Therefore, current reference data sets had to be acquired to support the planned Landsat data analysis.

In many previous agricultural remote sensing experiments, reference data were obtained by on-the-ground identification and recording of crop type and other information by the researchers or local USDA personnel. But, the amount of data which can be obtained in this way is restricted by the time and personnel available and generally can be done for only a few relatively small areas. Resources were not available to implement such an effort, even using sampling, for two

entire states.

During the CITARS project conducted by NASA/JSC, LARS, and ERIM, this type of ground observations was supplemented by interpreting aerial color infrared photography acquired concurrently and over the same area as the ground observations [5]. The accuracies of crop identification by photointerpretation routinely exceeded 95% and the data were successfully used for training and test purposes. It was therefore decided to take this approach one step further and make aerial photography the primary reference data source to identify and locate samples of wheat, corn, soybeans, and other cover types in the Landsat data.

After studying soil, climatology, and land use maps, flightlines were selected throughout each state to sample the variation in soils, land use, and crops. The flightlines were oriented north-south following major highways in Kansas and Indiana so that the aerial photography and Landsat data could be coordinated easily.

A 70 mm Hulcher two-camera system was used with color infrared and color transparency film. The average ground speed was 275 km per hour and photographs were taken, with both cameras, at intervals of 38 seconds, producing a continuous strip of imagery with an overlap of 25-30%. The average altitude for each flight mission was 3,000 meters. The approximate scale of the photography was 1:80,000. Each frame

of aerial photography included an area roughly four kilometers square (2.5  $\times$  2.5 square miles). Examples of the photography are shown in Figures 10 and 11.

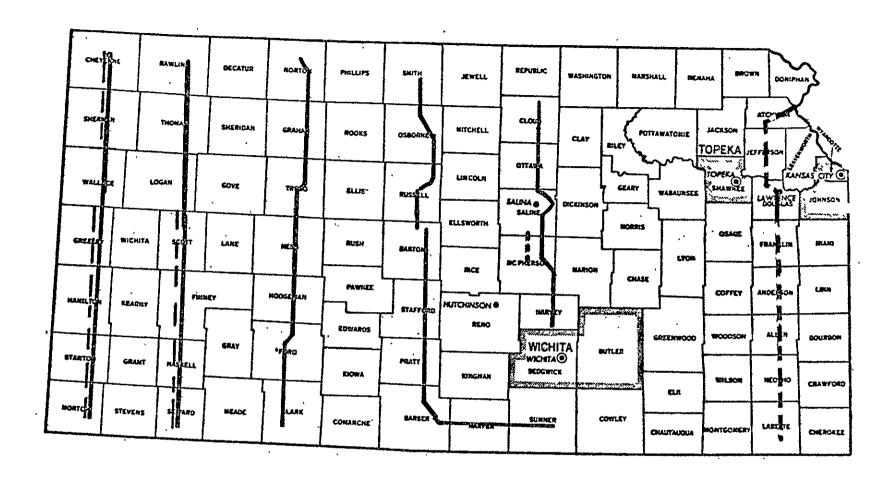
In Kansas, aerial photography was acquired on April 29-30 and June 26-27. Both dates were quite adequate for differentiating wheat from all other cover types. The June mission covered the eastern counties (and some western counties) while the April one covered the rest of the state (Figure 6).

The flightlines and dates of aerial photography acquisition for Indiana are shown in Figure 7. The May photography, when used concurrently with the July or August photography, helped to differentiate corn and soybeans from all other fields.

## 4.3 Digitization of Coordinates

The Landsat coordinates for county boundaries were needed in order to make county crop estimates. In addition, three to eight points were needed along the flightline in a county in order for the analyst to match a computer map of Landsat data to the aerial photography. To find coordinates, the following procedure was used:

- 1. Determine which counties are contained in the Landsat scene.
- 2. Locate 25-30 checkpoints in the Landsat scene.
- 3. Digitize these checkpoints on a 1:250,000 USGS map.



APRIL 20, 1975

JUNE 26-27, 1975

Figure 6. Kansas aerial photography flightlines and dates of photography acquisition.

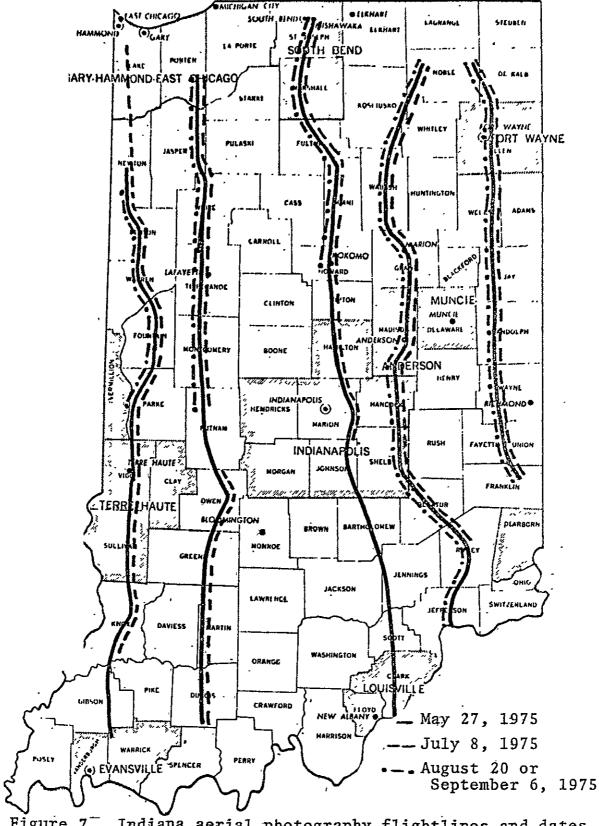


Figure 7. Indiana aerial photography flightlines and dates.

- 4. Digitize points defining county boundaries.
- 5. For each county that has aerial photography, digitize three to eight points along the flight-line.
- 6. Use a bivariate quadratic regression routine to fit coordinates of the checkpoints from the Landsat scene to the corresponding coordinates on the USGS maps. Then calculate Landsat coordinates for points defining county boundaries and checkpoints along the flightline.
- 7. Record the Landsat coordinates for county boundaries, and mark the Landsat coordinates for flightline points on the county maps.

In the following paragraphs each of the steps is described further.

The outlines of the state and all the county boundaries are displayed on a digital display device. Using the latitude and longitude for the Landsat scene center, the outline of the scene can be superimposed. A photograph taken of this image aids in determining which counties are covered.

In order to locate checkpoints, the data was displayed one channel at a time, in 16 gray levels. Twenty-five to 30 checkpoints were found, generally at the intersection of two highways, and the Landsat coordinates of these points were recorded.

The (x,y) coordinates of the checkpoints found in the Landsat scene, the points defining the county boundaries, and additional checkpoints along the flightlines are obtained from USGS 1:250,000 scale maps. A regression routine was used to fit the Landsat checkpoints to the checkpoints

digitized from the USGS maps. The Landsat coordinates of the county boundaries and additional points along the flight-lines were then listed and recorded on maps (Figures 8 and 9). The Landsat coordinates of the county boundaries were later used for tabulating county classification results. The coordinates of the points along the flightlines were used by the analysts to locate the flightlines in the Landsat data.

#### 4.4 Interpretation of Aerial Photography

Large scale aerial photography was used as reference data following the assumption that the crops of interest could be readily and accurately identified. Standard photointerpretation techniques were used to identify fields of wheat and nonwheat in Kansas and fields of corn, soybeans, and "other" in Indiana. The coordinates of the identified fields were then located in Landsat data. Wheat was relatively easy to identify in Kansas; corn and soybeans were more difficult to identify in Indiana. Fields which were not positively identified were not included as either training or test fields. Problems in photointerpretation, therefore, resulted in smaller training sets rather than inaccurate identification. Two general problems, clouds or haze and improper film exposure, were occasionally encountered, but did not seriously affect the photointerpretation process.

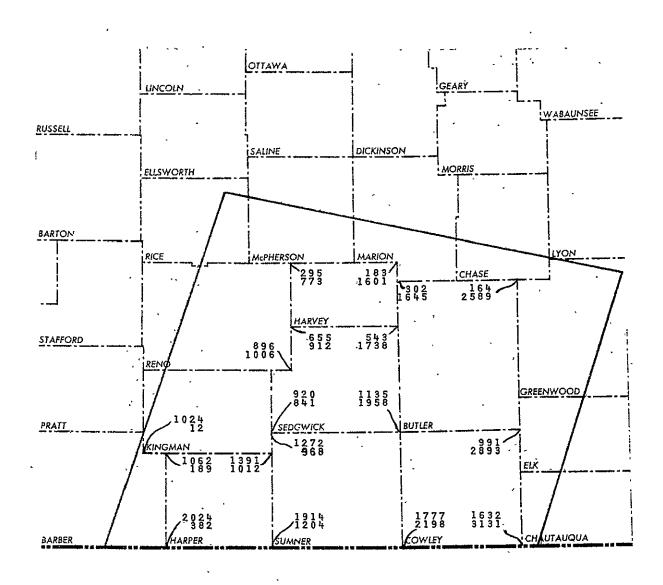
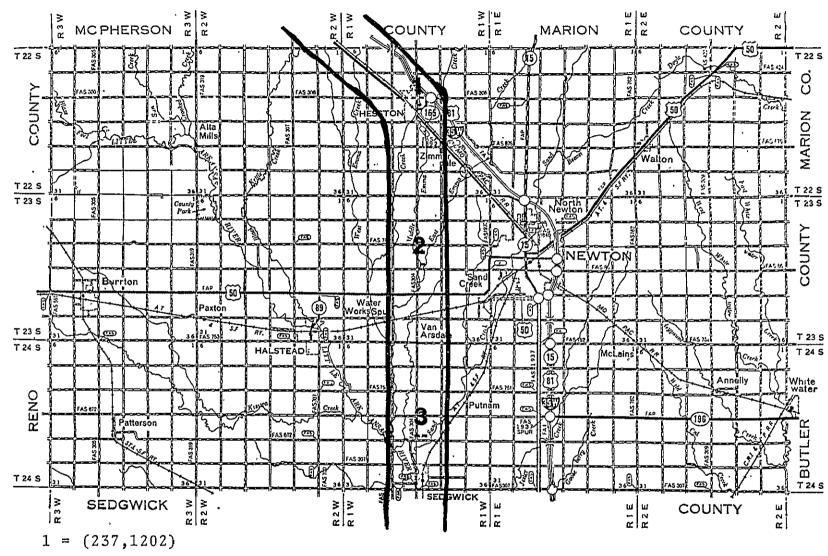


Figure 8. Example of Landsat coordinates of county boundaries.



<sup>2 = (402, 1250)</sup> 

3 = (539, 1297)

Figure 9. County map showing aerial flightline and Landsat coordinates of points along it (Harvey County, Kansas).

Examples of the aerial photography over Kansas and Indiana are shown in Figures 10 and 11, respectively. These figures illustrate scale, quality, and appearance of major cover types. The difference in the number and size of fields in a section of land in the two states is also illustrated.

#### 4.4.1 Kansas Wheat

Photography acquired on April 30, 1975, was used as reference data for all of Kansas except the Southeast CRD. this date the wheat fields had nearly total ground cover and were light green compared to alfalfa or clover and wheat during May. Clover and alfalfa were the only other crops achieving full ground cover and a bright green color at this time in the season. Confusion of wheat with these crops was occasionally a problem, but generally clover and alfalfa were brighter red on the color infrared film and could be discriminated from wheat. The planting patterns in wheat fields also helped in its identification. Pastures could usually be easily separated from wheat fields in the infrared photography. Color infrared photography was used exclusively for this date.

Photography of June 26-27, 1975, was used for a limited area in the southeast part of the state. By this date, winter wheat was mature and harvest was ready to begin. Thus, with the straw dead, the wheat fields are golden yellow, a color which readily separates them from any other major feature present at this time. Primarily the Ektachrome color positive





Figure 10. Examples of color infrared and color aerial photography acquired over Finney County, Kansas on April 20 and June 27, 1975, respectively.



Figure 11. Example of color infrared photography acquired over Wayne County, Indiana on August 20, 1975.

images were used for the interpretation at this date, since the wheat fields could be easily identified on it.

## 4.4.2 Indiana Corn and Soybeans

Almost complete coverage of the Indiana flightlines was achieved on May 27, 1975, but corn had not yet emerged and soybeans may not even have been planted at this time. Photography from this date, however, was useful in separating corn and soybean fields from other fields since corn and soybeans are the primary crops appearing as bare soil at this time.

The quality of the photography taken in July over Indiana was generally poor; there was a hazy overcast and the film was often overexposed. On the infrared film, corn fields appeared deep red and were confused with pasture. This photography was used only in conjunction with photography from another date.

During the period from August 20 to September 6, 1975, corn fields are tasseled, thus their green color as viewed from the air is not as intense. These fields are therefore easily separated from the soybean fields, which are at a full leaf stage, and have a uniform deep green color. Corn fields also exhibit more texture than most other cover types. This was the optimum period for obtaining photographic data over Indiana during 1975, and it was more extensively used as reference data than any of the other time periods. Only the color infrared images were used since soybean fields appeared as a bright red, and corn fields were of a less intense red

or brownish color.

#### 4.5 Analysis of Landsat Data

The Landsat data analysis techniques used in the investigation utilized the LARSYS Version 3 multispectral data analysis system. LARSYS is the software system, an integrated set of computer programs, for analyzing remote sensing data developed by Purdue/LARS during the past decade. The pattern recognition concept utilized in LARSYS represents a powerful and quantitative methodology for accommodating the multivariate nature of remote sensing data. While the LARSYS approach takes full advantage of modern computer technology for data processing, man is an indispensable part of the analysis process. Thus, the techniques are better described as "computer-assisted" rather than "automatic". The processing functions of LARSYS are shown in Figure 12. Its theoretical basis and details of the algorithm implementation are described in references [24] and [22], respectively.

In utilizing the LARSYS software for analyzing multispectral scanner data, one normally follows a procedure that
involves: (1) defining a group of spectral classes (training
classes); (2) specifying these to a statistical algorithm
which calculates a set of defined statistical parameters;
(3) utilizing the calculated statistics to "train" a pattern
recognition algorithm; (4) classifying each data point within

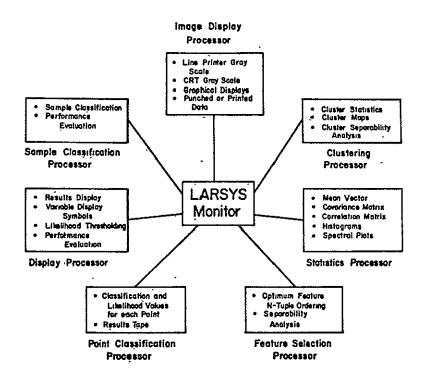


Figure 12. Analysis functions of the LARSYS software system.



the data set of interest (such as part of a Landsat frame) into one of the training classes; and finally '(5) displaying the classification results in either map or tabular format (or both), according to the specifications of the application.

During the past few years, experience at LARS has shown that there are many possible refinements in the methodology utilized by the analyst for obtaining training classes, while the rest of the procedure does not vary much from one analysis task to another. The most common techniques for defining training classes involve the so-called "supervised" approach, and the "unsupervised" or "clustering" approach.

In the "supervised" approach, the analyst selects fields of known cover types and specifies these to the computer as training fields, using a system of (x,y) coordinates. The statistics are obtained for all categories of cover type in each area to be classified. The data are then classified and the results evaluated. Because the analyst had defined specific areas of known cover types to the computer, such classifications are referred to as "supervised".

The second method uses a clustering algorithm which divides the entire area of interest into a number of spectrally different classes. The number of spectral classes into which the data will be divided must be specified by the analyst. The spectral classes defined by the clustering algorithm are then used to classify the data, but at this point the analyst does not know what cover type is defined by each of the

spectral classes. After the classification is completed, the analyst will identify the cover type represented by each spectral class using available reference data or cover type maps. Because the analyst does not need to define particular portions of the data for use as training fields, but must only specify to the computer the number of spectral classes into which the data is to be divided, a classification using this procedure is referred to as "unsupervised".

Additionally, several variations of these basic methods for defining training classes are possible. One is to select training areas of known cover type (a supervised approach up to this point), but then utilize the clustering algorithm to refine the data into unimodal spectral classes for each cover type. This is called a "modified supervised" approach and is the approach which was used in this investigation.

The remainder of this section describes the analysis methodology and additional details of the training procedure. An overview of the steps in the analysis sequence is shown in Figure 13.

# 4.5.1 <u>Selection of Training Data</u>

The accuracy of classification results is highly dependent upon the training data. Selection of training areas was based on two factors; first, the amount and quality of reference data (aerial photography) available, and second, the presence of a representative sample of cover types of the area(s) to be classified. To insure that the best

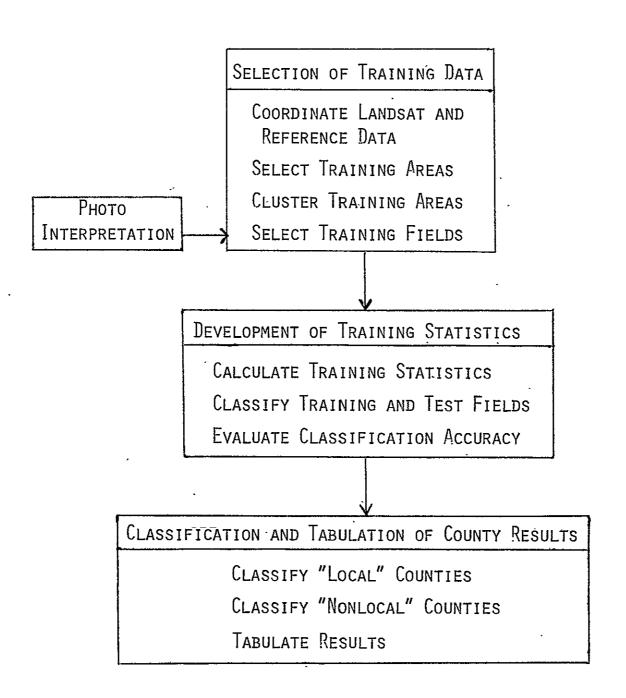


Figure 13. Flowchart of procedures used in analysis of Landsat data.

classification accuracy is obtained, a sample of every spectral class of each cover type should be included in one or more of the training areas. This provides a reasonably representative training set to the classification algorithm.

The analyst's first task was to gather and coordinate the information available about the county or counties to be analyzed. The Landsat scene had been selected (see Sec. 4.1) and the Landsat coordinates for each county boundary had been found (see Sec. 4.3). In addition, county maps had been prepared showing the Landsat coordinates of the checkpoints along the aerial photography flightline (Figure 10). The frame numbers of the aerial photography for each county were marked on the map. From this information, the analyst could determine the areas in the Landsat data corresponding to frames of aerial photography and then select the areas to be used for training the classifier.

Training areas of 100 lines and 100 columns (approximately 8 x 5.5 km) of Landsat data were selected in areas corresponding to aerial photography. For smaller counties, especially in Indiana, three to five training areas were chosen covering the entire flightline. In Kansas, four to six areas were selected with at least one in both the northern and southern portions of the county in order to adequately represent the variation present in the county.

To facilitate locating agricultural fields in the Landsat data, a spectral class map was produced by clustering each

training area. The clustering algorithm implemented in LARSYS finds natural groupings in the spectral data utilizing all four wavebands. Generally six to eight classes were sufficient to provide an image on which the crop fields were readily identifiable. This approach was found to be more satisfactory than working with gray scale maps of a single spectral band.

Examples of cluster maps are shown in Figures 14 and 15; the color infrared photographs of the same areas were shown in Figures 10 and 11. The cluster maps were matched with the corresponding frames of aerial photography, and roads, towns, and field boundaries were sketched on the cluster maps.

Fields were marked on the cluster maps and their cover type identified from the aerial photography. During the photointerpretation process, the analyst became familiar with the variation in wheat, corn, soybeans, and other fields.

Training fields had to meet three criteria. First, the cover type of the fields selected for training had to be positively identified by the photo-interpreter. Secondly, the fields themselves must be of only one cover type; for example, if a ditch ran through the field, the analyst would avoid the ditch and select samples on either side of it. Thirdly, the training fields must adequately represent the variation present in the cover types throughout the area to be classified; to insure this, the fields were geographically dispersed throughout the flightline. The Landsat coordinates

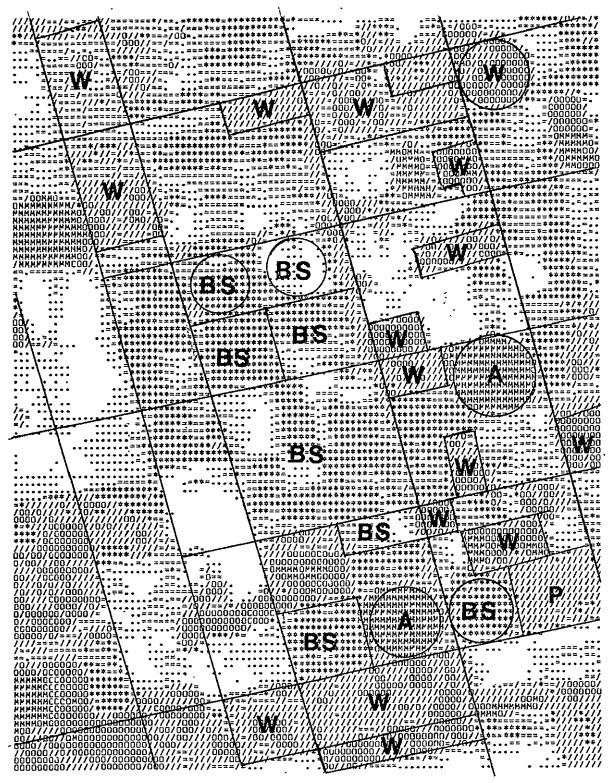


Figure 14. Example of cluster map used for location and identification of fields in Finney County, Kansas. (W = wheat, A = alfalfa, BS = bare soil, P = pasture)

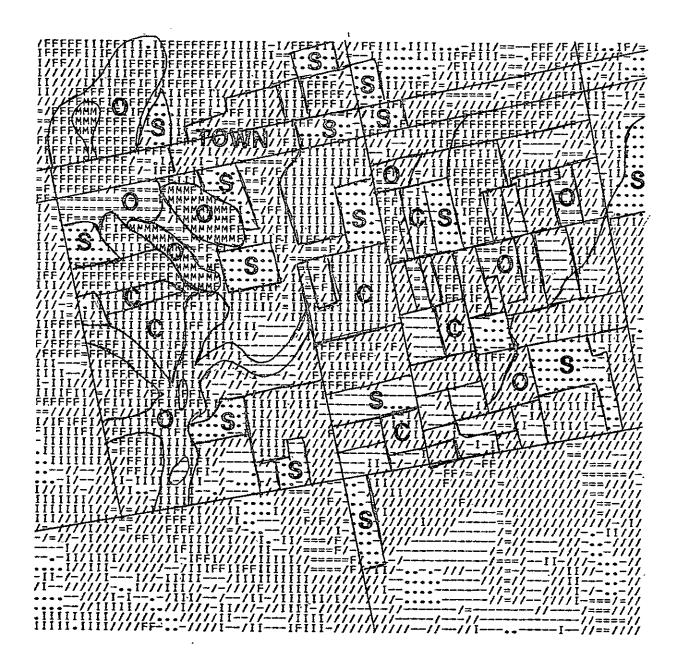


Figure 15. Example of cluster map used for location and identification of fields in Wayne County, Indiana. (C = corn, S = soybeans, O = other)

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of field center (non-boundary) pixels were then obtained and field description cards prepared.

If there were any reservoirs or rivers in the county, training samples were obtained for water. If there were no bodies of water in the flightline, the analyst obtained an additional cluster map which would include water bodies.

Training samples for water were then selected from this area.

As a general rule at least 25 wheat samples and 25 other samples were chosen in Kansas. In Indiana, fields were much smaller and homogeneous samples were difficult to find due to the large proportion of boundary pixels. In general, more than 25 samples each of corn, soybeans, and other were chosen, but the samples were small compared to those for Kansas.

The number of samples used for training the classifier in Kansas and Indiana is shown in Tables 6 and 7, respectively. The median number of fields used for training in Kansas was 66 and the median number of pixels used was 2600. In Indiana, the corresponding figures are 163 fields and 2750 pixels.

# 4.5.2 <u>Development of Training Statistics</u>

The training fields for each major cover type have been selected, but the spectral characteristics of each class have not been calculated. Each major cover type must be divided into its spectral subclasses, each of which must be a unimodal distribution to satisfy the assumptions of the maximum likelihood Gaussian classifier and is characterized by its mean vector and covariance matrix. Confusion between the

Table 6. Number of fields and pixels used for training and testing the classifier in Kansas.

	Training Samples		Test Samples	
	No.	No.	No.	No.
County	Fields	Pixels	Fields	Pixels
Northwest District				
Cheyenne ·	47	1587		
Graĥam	59	1225		
Norton	30	600		
Sherman	76	2609	75	2289
West Central District				
Greeley	82	3090	81	2672
Ness	82	2400		<del>-</del>
Trego	50	2955	51	2345
Wallace	67	4139		
Southwest District		• •		
Finney	127	2917	•	
Ford	119	3320	121	2763
Hamilton	117	7161	96	5785
Haskell	77	2118	50	0,05
Hodgeman	82	5105	83	4927
Seward	43	1001	0.5	7527
Stanton	98	6337	132	2884
North Central District	50	0557	102	2004
Cloud	77	1174		
Osborne	39	1446		
Ottawa	56	3215		
Smith	97	2924		
Central District	<i>3</i>	4264		
Barton	55	2928		
McPherson	5 <i>7</i>	2562		
Russell	42	1257		
Saline	50	1847	41	994
South Central District	30	104/	4 T	334
Barber	58	1942	25	2147
	56 69		45	414/
Harvey		2202	71	7177
Pratt Stafford	69 62	2850		3433
		2586	31.	2522
Sumner	49	2244		
Southeast District	126	4225	171	43.40
Allen-Neosho	126	4225	131	4149

Table 7. Number of fields and pixels used for training the classifier in Indiana.

	Training Samples		
County	No. Fields		
Northwest District			
Benton	144	3271`	
Lake	163	3424	
LaPorte	167	3976	
Newton	145 ·	2684	
Pulaski-Starke	192	4475	
White	224	3002	
West Central District			
Fountain-Parke	337	4419	
Montgomery	223	3715	
Owen	82	1595	
Tippecanoe	92	1685	
Vigo	120	2543	
Warren	63 .	1269	
Central District			
Decatur	155	2748	
Grant	163	1690	
Hamilton-Howard-Tipton	284	4145	
Johnson-Shelby	174	2825	
Madison	158	1888	
Eașt Central District			
Fayette	110	1868	
Jay	166	1862	
Rando1ph	277	3035	
Wayne	203	2617	

spectral subclasses of different cover types must be minimized to decrease the error in classification. The adequacy of the training statistics should be evaluated before carrying out large area classifications.

In order to satisfy the first of these three requirements, the cluster function was again used to obtain subclasses for the major cover types of wheat and nonwheat in Kansas and corn, soybeans, and other in Indiana. This time, instead of one large rectangular area, the field center samples of each of the major cover types were clustered separately to find natural groupings or spectral classes within the cover types.

Statistics were calculated to represent each spectral class and the transformed divergence between each pair of classes was calculated. The saturating transformed divergence, a number between 0 and 2000, provides a measure of the distance between classes in multi-dimensional space. High values indicate class pairs which are more separable and which, if grouped, would yield a bimodal distribution. Class pairs with small divergence values are spectrally similar and may be confused with each other during classification. If classes of different cover types were spectrally similar, the analyst inspected the fields involved by checking the location and type of field on both the cluster map and the aerial photography. If an error in field identification or location had been made, the class in error was deleted. If no error occurred, the confusion classes were left in the training statistics since deleting

one or both of them would have biased the classification results.

Test field classification results, if available, or training field results were used to evaluate the adequacy of the training statistics before the county was classified in order to allow for additional training if required. For many counties in Kansas, there were enough sample fields available that both a training and a test set could be developed. A statistical test showed that the proportion estimates calculated using training field performance matrices were not significantly different in accuracy from estimates calculated using test field performance matrices. In Indiana, where the field sizes were small compared to Kansas, the number of usable samples was much smaller, and selecting test fields from the sample fields would have greatly reduced the size of the training set.

# 4.5.3 Classification and Tabulation of County Results

The final training statistics were used to classify a systematic random sample of the Landsat pixels within each county (Figure 16). In a systematic random sample, the first sample is chosen randomly and the remainder are determined by a constant sampling interval. Systematic random sampling was convenient and has the advantages of high precision and excellent geographic stratification [9].

For about 60 counties in Kansas and a few in Indiana, every other line and column was classified, a one-fourth

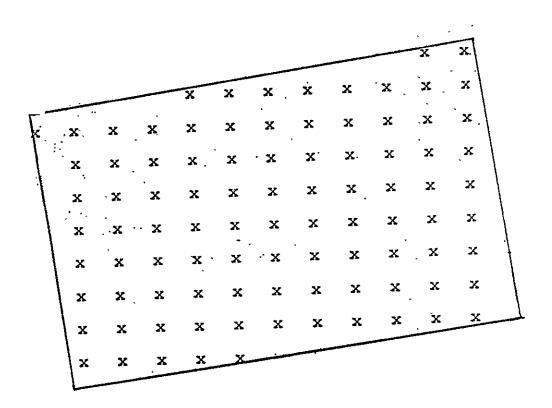


Figure 16. Schematic of a systematic random sample of Landsat pixels classified within a county boundary.

sample. However, every fourth line and column, a onesixteenth sample, was used for the remainder of the counties. Tests showed that there was no significant difference in results obtained between these two sample sizes.

When a county was classified with a training set at least partially trained with fields from that county, the classification is labelled "local". A "nonlocal" classification is one in which the training set does not contain any training fields from the county classified. The training set used to perform a nonlocal classification came from a county in the same Landsat frame having similar soils and land use. Figure 17 is a map of Kansas showing geographically the local and nonlocal classifications and the source of training data for nonlocal classifications. Similar information for the counties classified in Indiana is given in Figure 18. Tables Al and A2 in the appendix summarize the Landsat frame, date of data, and source of training statistics for all counties classified in Kansas and Indiana.

The number of points of each major cover type and the total number of points in the county were tabulated. These points fall within an irregular polygon in the Landsat data which corresponds to the county boundaries. Using the coordinates of cities and large towns which had been obtained earlier, the number of points of each major cover type in the urban areas were tabulated and subtracted from the county totals. These adjusted totals form the base of the area and



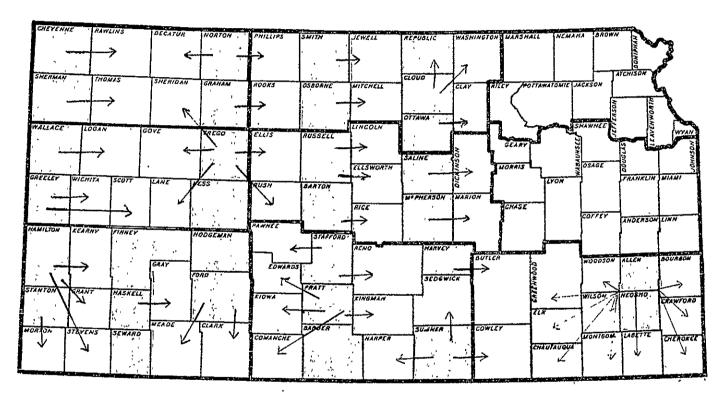


Figure 17. Local and nonlocal classifications in Kansas. Arrows point from the source of training statistics to the area classified; shaded areas denote local recognition counties.

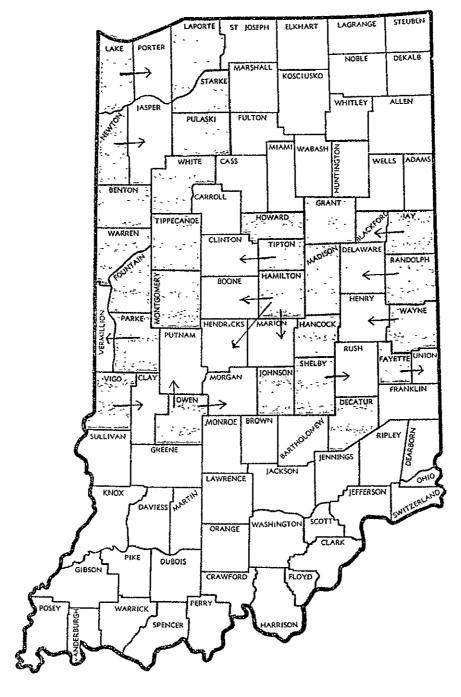


Figure 18. Local and nonlocal classifications in Indiana. Arrows point from the source of training statistics to the area classified; shaded areas denote local recognition counties.

proportion estimates for the county.

#### 4.6 Preparation of Area and Variance Estimates

Following classification, crop area and proportion estimates were made. Estimates of the areal extent or proportion of a cover type were desired for county, crop reporting district, and state levels. The county was the smallest unit for which an estimate was wanted, so estimates of the cover types of interest were made for each county and then aggregated to the district and state levels. Steps in the area estimation procedure included: (1) calculation of the area and proportion estimates, (2) correction of the estimates for classification bias, and (3) calculation of variance estimates. For counties in which Landsat classifications were not performed, a regression procedure utilizing historical data and current Landsat estimates was used.

#### 4.6.1 Area and Proportion Estimates

The Landsat estimated proportion of the  $i\frac{th}{}$  crop in the  $j\frac{th}{}$  county was calculated using the equation

$$\hat{p}_{ij} = \frac{n_{ij}}{n_{j}}$$

where  $n_{ij}$  is the number of pixels classified as crop i and  $n_{j}$  is the total number of pixels in the sample. The estimated hectares of crop i in the  $j\frac{th}{}$  county can be calculated in two equivalent ways:

$$\hat{h}_{ij} = \hat{p}_{ij}h_{j}$$

where  $\hat{p}_{ij}$  is defined as above and  $h_j$  is the number of hectares in the county, or

$$\hat{h}_{ij} = k n_{ij}$$

where  $n_{ij}$  is as above and k is the area in hectares of a pixel (approximately 0.45).

Area and proportion estimates for the crop reporting districts and the entire state are aggregated from the county estimates. The area estimate of crop i for a CRD is found by  $\Sigma$   $\hat{h}_{ij}$ , summing the area estimates from all the counties in the CRD. The proportion of crop i in a CRD is found by  $\Sigma \hat{h}_{ij}$  where the summations are taken over all the counties in the CRD and  $\hat{h}_{ij}$  and  $h_{j}$  are as defined above. Area and proportion estimates for entire states are found similarly.

# 4.6.2 Correction\_for Classification Bias

Experience has shown that it is inevitable that some pixels are incorrectly identified by the maximum likelihood classifier. The primary source of these errors is overlapping density functions for two or more classes. For example, some corn looks like soybeans and/or some soybeans are spectrally similar to corn. Classification errors of this type cause the resulting area estimates to be biased. However, if the error rates are known the area estimates can be adjusted or unbiased after the classification has been performed. This technique was first used in the 1971 Corn Blight Watch

Experiment [18] and later in a Landsat-1 investigation by LARS [4].

An estimate of the classification error rates is the matrix of training or test field classification performance,

$$E = \begin{pmatrix} e_{11} & e_{12} \\ e_{21} & e_{22} \end{pmatrix}$$

where  $e_{ij}$  is the proportion of samples of type i classified as type j. If P is the vector of true proportions of the cover types and  $\hat{P}$  the proportions estimated from the Landsat data, then

$$\hat{P} = E^{t}P$$
.

Since  $\hat{P}$  and E are known from the classification, but P, the vector of true proportions, is not known,

$$P = (E^t)^{-1} \hat{p}$$

is solved. The example of Figure 19 shows how this is done.

It is possible for this method to give a negative value for the proportion of a cover type. Since it is unrealistic for an estimate of a proportion or probability to be negative, an alternative problem was considered when this occurred:

$$\begin{array}{c|c}
\min \\
0 \le p_{1} \le 1
\end{array} \quad \left| \quad P - \left(E^{t}\right)^{-1} \cdot \hat{P} \quad \right| \quad .$$

for all  $p_i$ , elements of the vector P. This is equivalent to minimizing the Euclidean distance (denoted by  $|\cdot|$ ) between the true proportion and the Landsat corrected estimate. The vector of proportion estimates after bias correction is

$$E = \begin{pmatrix} .85 & .15 \\ .18 & .82 \end{pmatrix}$$

$$E^{T} = \begin{pmatrix} .85 & .18 \\ .15 & .82 \end{pmatrix}$$

$$(E^{T})^{-1} = \begin{pmatrix} 1.2239 & -.2687 \\ -.2239 & 1.2687 \end{pmatrix}$$

$$P = \begin{pmatrix} 38.9 \\ 61.1 \end{pmatrix}$$

$$P = \begin{pmatrix} 1.2239 & -.2687 \\ -.2239 & 1.2687 \end{pmatrix} \begin{pmatrix} 38.9 \\ 61.1 \end{pmatrix} = \begin{pmatrix} 31.2 \\ 68.8 \end{pmatrix}$$

$$SRS \qquad LANDSAT \qquad LANDSAT \qquad CORRECTED \qquad CORRECT$$

Figure 19. A numerical example of classification bias correction (Cloud County, Kansas).

38.9

31.2

31.6

denoted by  $\hat{P}$ . The discussion of bias correction generalizes to n cover types of interest with E being an n x n matrix and the vectors having n components.

The corrected estimate will be unbiased if the error matrix found from the test or training field performance is the true error matrix. It may not be truly unbiased because of photointerpretation difficulties or because the flightline might not be representative of the entire area classified.

### 4.6.3 <u>Calculation of Variance Estimates</u>

In addition to knowing the accuracy of an estimate, it is desirable to know the precision, or variance, of the estimate. The variances of the proportion and area estimates were obtained as follows. Since each pixel is classified as crop i or not, the binomial distribution can be used to obtain the variance of the bias-corrected proportion estimates. For the  $j\frac{th}{c}$  county, an estimate of the variance is given by

$$v\left(\hat{\hat{p}}_{ij}\right) = \frac{\hat{\hat{p}}_{ij}\left(1-\hat{\hat{p}}_{ij}\right)}{n-1}\left(1-f_{j}\right)$$

where  $f_j$  is the county sampling fraction [8]. For individual county estimates, the sampling fraction can be ignored (though it is not negligible) to give a conservative estimate of the variance. As

$$\hat{h}_{ij} = \hat{p}_{ij} h_{j}$$

the variance of the area estimate  $\hat{\hat{h}}_{ij}$  can be calculated by

$$v\left(\hat{h}_{ij}\right) = h_{j}^{2} v\left(\hat{p}_{ij}\right)$$

where  $h_{i}$  is the total number of hectares in the county.

In calculating the proportion estimate from the sample the assumption is made that each pixel would be classified as a particular crop or not classified as that crop, which leads to a multinomial or binomial model of the classified data. The binomial distribution can be used to estimate the total number of wheat pixels and the percentage of wheat in the area. Theoretical estimates of the sampling error are then available [8]. It is also assumed that there is no cyclic pattern in the data to bias the estimate from a sample taken systematically. To test these assumptions, a sampling study was performed early in this project.

The study examined the sampling error produced for a given sampling fraction against the theoretical error given by using binomial distribution theory. In order to measure just the effect of sampling, the error introduced in classification was ignored by comparing the various samples to a 100% sample. The results are based on classifications of Rice and Morton Counties, Kansas, and were substantiated by further tests in Benton and Wayne Counties, Indiana.

In the Kansas sampling study, estimates of both the total number of wheat resolution elements and the percentage of wheat in the area were calculated for sampling fractions of 50, 33.3, 25, 11.1, 10, 6.25, 4, and 2.8 percent. These

samples were taken systematically. For example, an 11.1% sample of the area was obtained by tabulating the classification with both a line and column interval of three. Nine 11.1% samples were selected with a different starting point for each sample. The theoretical variance of these sample estimates was calculated from the binomial distribution and compared to the variance among the repeated estimates of the same sample size. For example, the theoretical variance of an 11.1% sample was calculated and then compared to the variance of the nine sample estimates.

The results of the study (Table 8) showed that in all cases the two variances were not significantly different, indicating that the theoretical estimate of the sampling error based on the binomial distribution can be used as the estimate of the variance of the proportion estimate. The Morton results show a cyclic effect due to "six line scan" noise. In practice, Landsat data with such a noise problem was avoided.

Wayne and Benton Counties in Indiana were used to test the applicability of the Kansas results to Indiana. The results were consistent with those of Kansas.

The variance for a crop reporting district can be obtained in two ways. The variance can be calculated as though a systematic random sample were taken throughout the district or it can be calculated considering each county as a stratum. The estimated variance for crop i in the stratified case would be given by:

Table 8. Theoretical and computed sampling errors of wheat proportion estimates for different sample sizes in two counties in Kansas.

	Štanda	ard Error (%)
% Sample	Theoretical	Computed
Rice County		
50.0 33.3 25.0 11.1 10.0 6.25 4.0 2.8	0.0902 0.1277 0.1563 0.2555 0.2717 0.3509 0.4453 0.5358	0.0361, 0.1126* 0.1018, 0.1597 0.0992 0.1824 0.1752, 0.1937 0.2812 0.2797 0.4890
Morton County		
50.0 33.3 25.0 11.1 10.0 6.25 4.0 2.8	0.0867 0.1226 0.1501 0.2455 0.2599 0.3372 .0.4241 0.5152	0.1293, 0.9233 0.0430, 1.0067 0.7637 0.8799 0.3358, 0.6939 0.6948 0.3405 2.6950

<sup>\*</sup> 50.0%, 33.3% and 10% systematic samples can be taken in two ways. For example, a 50% sample can be either every other line or every other column.

$$\Sigma W_j^2 = \frac{\hat{p}_{ij} \left(1 - \hat{p}_{ij}\right)}{n_j} \left(1 - f_j\right)$$

where the summation is taken over all counties in the crop reporting district [8].

In essence, it matters little what formula is used to calculate the variance estimates whether conservative or not, because the estimates are very small in either case. The distribution in Indiana is actually given by the multinomial, but the variances can be calculated by considering each crop separately with the binomial assumptions.

# 4.6.4 Estimation for Counties Without Landsat Data

An alternative approach for crop estimation must be taken when adequate data for Landsat classification is not available for an area. One approach to this problem lies in formulating a regression equation from which a crop prediction can be made.

Regression is valid as a predictor only for the population from which it is derived. This predictor will not be valid for a county which has historical crop acreage or county size falling outside the range of values used in the derivation of the regression equation. For these counties, the 1974 USDA/SRS area estimates were used as the 1975 estimates. Revised estimates from Kansas and preliminary estimates from Indiana were used.

For Kansas, the regression model used to predict the area in hectares of wheat in a given county was:

$$\hat{y} = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \beta_3 x_3$$

where  $x_1$  is the 1974 USDA/SRS wheat acreage for the county,  $x_2$  is the 1973 USDA/SRS wheat acreage for the county, and  $x_3$  is the total number of acres in the county. The coefficients  $\beta_0$ ,  $\beta_1$ ,  $\beta_2$ , and  $\beta_3$  are estimated by using the available Landsat estimates as y values. A pseudo-Landsat estimate is made by applying these coefficients to the x values of the counties to be estimated.

Only historical data could be used in the regression in order to simulate real-time estimation. It was felt that wheat data before 1973 should not be considered because major increases in the wheat acreage planted occurred beginning in 1973. The area of the county was also included as a factor which might contribute to the amount of wheat grown.

For Indiana, similar regression models were used to predict the area in corn and soybeans. Again, the variables considered as predictors were the number of acres in the county and the USDA/SRS estimates of acres harvested in 1973 and 1974 for corn or soybeans. The regression model used was:

$$\hat{y}_{i} = \beta_{0} + \beta_{1} x_{1i} + \beta_{2} x_{2i} + \beta_{3} x_{3}$$

where  $\hat{y}_i$  denotes the area in hectares of crop i,  $x_{1i}$  is the 1974 USDA/SRS estimate of acreage in crop i for the county,  $x_{2i}$  is the 1973 USDA/SRS estimate of acreage in crop i in the county, and  $x_3$  is the total number of acres in the county.

#### 4.7 Evaluation of Results

Once an adequate training set has been defined, it is not difficult to classify large geographic areas using computer analysis techniques. However, unless the accuracy of such computer classification results can be verified, little has been accomplished by simply classifying the data over various areas of interest.

In this investigation two quantitative evaluation techniques were used to judge the accuracy of crop classifications and area estimates. One evaluation involved statistical sampling of individual areas of known cover types (designated as test fields). This offers an effective method of examining inclusive and exclusive classification errors for the various crops or cover types. Such techniques, however, must be used with caution, and must be carefully designed to provide statistical reliability of the results. In general, areas need to be selected in such a way that the number of resolution elements in the test areas for each cover type are approximately in proportion to the amount of that cover type present in the area.

A second quantitative technique for evaluating classification accuracy is comparison of area estimates from the computer classification and area estimates obtained by some conventional method. Ideally, crop area measurements from large contiguous areas would be used for comparison.

Realistically, it is not possible to acquire a large amount of such data. Therefore estimates of the crop areas or proportions must be used. The USDA/SRS annually publishes estimates of the acreage of major crops for counties, districts, and states. Estimates or measurements for a smaller unit such as a township are generally not available.

In addition to evaluating the classification accuracy, several factors which might have affected accuracy were examined.

# 4.7.1 Assessment of Training and Test Field Classification Accuracy

Test fields are frequently used to evaluate the accuracy of the Landsat classifications. Areas with a known cover type which were not used for training are chosen as test fields. These are then classified and the accuracy of the classifier determined by the proportions of pixels which are correctly identified. If these fields have been randomly selected and their classification accuracy is high, then the classification of the entire area should be accurate.

In this project test fields were chosen in a manner similar to training fields. Some of the fields identified from the aerial infrared photography were randomly selected as test fields. The method of random selection depended upon the analyst and included systematic sampling, stratified random sampling, and simple random sampling. However, in some counties all the available fields were used for

training, leaving none for test. In these cases, training field performance was evaluated to determine the accuracy of the classifier, since a statistical test of counties with both test and training fields showed that using training fields to evaluate classification accuracy was not significantly different from using test fields.

# 4.7.2 <u>Statistical Comparison of Landsat and USDA/SRS</u> <u>Estimates</u>.

The standard of comparison for Landsat estimates was the USDA/SRS estimate of acres harvested. SRS estimates were used primarily because of their availability on a state, crop reporting district, and county basis for 1975. There is a national agricultural census which also provides these estimates, but it is performed only every five years and was not taken in 1975. Acres harvested were used rather than acres seeded because: (1) the acquisition of Landsat data used in this analysis was closer to harvest time than to seeding time and (2) the harvested acreages are used for estimating total production. Estimates of both the proportion of total land area and of the area in hectares of a crop were considered as variables.

The purpose of USDA/SRS crop surveys is, primarily, to make national estimates and, secondly, state estimates. The state estimates are considered to be unbiased and to have small coefficients of variation, generally not exceeding about 5% for major crops [23]. The SRS does publish county and

crop reporting district estimates, but coefficients of variation are not calculated for these estimates. It is expected that the county and CRD estimates will not be as accurate as the state and national estimates, and that the coefficients of variation will be larger at the county level. The SRS county estimates then are not the ideal standards for comparison, but must be used due to lack of any more reliable data.

The method used to arrive at county estimates varies from state to state. In Indiana, county estimates are made on the basis of mail surveys. About 12,000 questionnaires are mailed to get a response of at least 4,000. This should guarantee at least 50 responses per county on which to base The mail survey results are adjusted for the the estimates. difference from the June enumerative survey (E. L. Park, State Statistician, Indiana, personal communication). Kansas, however, uses information from three different surveys to calculate county estimates. The first is the annual State Farm Census which is supposed to be an enumeration of all farming operations in the state, but which contains some incompleteness. Mail surveys from June and late summer are combined with the census data to form a composite area estimate for each county. These are then adjusted for various factors and scaled to add to the state estimate (M. E. Johnson, State Statistician, Kansas, personal communication).

The levels for testing Landsat against SRS estimates were determined according to the problem at hand. In choosing a significance level, a large  $\alpha$  is chosen to minimize the chance of claiming the hypothesis of equality is true when it is really false; a small value of  $\alpha$  is chosen to minimize the chance of rejecting the hypothesis of equality when it is actually true. To ascertain whether SRS and Landsat estimates were close, the two estimates were obtained and the hypothesis of their equality, the null hypothesis, was tested. Statistical tests are not designed to prove that the null hypothesis is true, although in this case that is what we did want to conclude. In order to be reasonably certain that the SRS and Landsat estimates are the same, the probability of accepting the hypothesis of equality, when it was in fact false, was made very small. This was achieved by choosing a large value of α such as 0.25.

## 4.7.3 Analysis of Factors Affecting Classification Accuracy

In order to perform statistical tests on the Landsat estimates, normality and homogeneity of the data must be considered. Standard tests for homogeneity were not useful here because they consider the variance of the sample variances, which in this case was zero because the variance  $\sigma^2$  is determined rather than estimated by the large sample size used in Landsat estimation. Instead, the range was used to determine if the variances were homogeneous for tests on proportions.

Variances are stable only for proportion estimates in the 0.30-0.70 range [1]. Since some values of the Landsat proportion estimates fell outside this range, a transformation was required. For this range, p was transformed by  $\sqrt{p}$  [1].

The nonhomogeneity of the data affects the statistical test results by introducing a bias into the test statistic, in this case either an F-statistic or a t-statistic. The bias of the F-statistic for the Kansas proportion variances was calculated and found to be 1.29 [6]. Thus, when testing a hypothesis with a significance level of  $\alpha$  = 0.05, the hypothesis is really being tested with  $\alpha$  = 0.09, and will be rejected too often. For this amount of bias, p should be transformed.

The bias of the test statistic for Kansas area estimate variances was found to be 1.17. Thus when testing a hypothesis with a significance level of  $\alpha=0.05$ , the hypothesis would really be tested with  $\alpha=0.07$ . This is not as biased as is the case with the proportion variances, though the null hypothesis would be rejected slightly too often. Testing was performed on these variables without transformation. With larger sample sizes, homogeneity tends to be a minimal problem. For Indiana, the proportion estimates were transformed and the hectare estimates were not, following the same pattern as for Kansas.

Numerous tests were made to identify and assess factors which might affect the accuracy of the area and proportion

estimates. Those factors tested included: date of the Landsat coverage, date of the aerial photography (Indiana only), effect of the data analyst (Kansas only), the effect of local versus nonlocal recognition, and the effect of geographic location (crop reporting districts).

For Kansas, two types of tests were made for testing the effect of date. The first was a paired comparison of 10 counties which had been classified twice using two different Landsat frames. The second type of test, done in both Kansas and Indiana, used all counties which were classified and tested for a difference due to groups of dates. A limitation of this test is that date effects may be confounded with other factors such as geographic location.

Tests for the effect of aerial photography date were not done in Kansas because essentially only one date was used. For Indiana, all counties were included in the analysis and tests were performed in the same manner and with the same limitations as the tests for the effect of date of Landsat data.

In tests for the data analyst and local vs. nonlocal recognition effects, all available data were utilized. In tests to determine the accuracy of a CRD or state, duplicate observations were not permitted. Of these duplicates, the estimate derived from the Landsat pass closest to harvest was used without reference to which one was closer to the SRS estimate.

#### 5.0 WHEAT IDENTIFICATION AND AREA ESTIMATION IN KANSAS

In this section the results of the Landsat data analysis for winter wheat identification and area estimation in Kansas are presented and evaluated. The material includes a discussion of factors affecting classification accuracy, comparisons and evaluations of training and test field classification performance, and comparisons of USDA/SRS estimates to Landsatderived estimates of the area and proportion of wheat. Finally, the accuracy and precision of the Landsat estimates are discussed.

## 5.1 Analysis of Factors Affecting Classification Accuracy

Although an assessment of factors affecting classification performance was not a primary objective, several analyses to assess factors which might have influenced classification results were performed in order to more fully understand and interpret the results. The variables tested included: Landsat acquisition date, data analyst, local vs. nonlocal classifications, and the interaction of date and locality. The results of these tests are presented in this section.

#### 5.1.1 Effect of Landsat Acquisition Date

Ten of the 13 counties in the South Central Crop Reporting District were classified twice, using data from two different Landsat passes. All counties were classified using April data and then reclassified using either May or June data (Table 9). Since these were the only counties for which multitemporal data were available, they were used to explore the effect of dates on classification performance. The "goodness" of an estimate was considered to be its closeness to the SRS estimate. Paired t-tests showed that there was no significant difference ( $\alpha$  = 0.25) in the accuracy due to the date of Landsat coverage. The inference of these tests is not strong due to the small sample size, so a further study on the effect of dates with larger samples was performed.

A second analysis, including all counties in the seven districts classified, was performed to determine if there was an effect due to the date of the Landsat data acquisition, ignoring other factors. Five groups of dates were considered: early April, early May, late May, mid-June, and early July. An analysis of variance showed that neither the proportion nor area estimates were significantly affected by Landsat data acquisition period. These results indicate that date was not a major factor influencing the classification performance and that all counties regardless of the date of Landsat data

Table 9. Comparison of wheat estimates from April and May or June Landsat data acquisitions to USDA/SRS harvested estimates, South Central Crop Reporting District, Kansas.

		USDA/SRS Harvested			Landsat Classification		erence m SRS
County	Date	Hectares	Proportion	Hectares,	Proportion	Hectares	Proportion
•		(000)	(%)	(000)	(%)	(000)	(%)
Barber	April	69.1	23.3	23.1	7.8	-46.0	-15.5
	May	69.1	23.3	89.4	30.1	20.3	6.8
Comanche	April	43.4	20.9	31.1	15.0	-12.3	- 5.9
	May	43.4	20.9	46.3	22.3	3.0	1.4
Edwards	April	53.1	33.4	58.0	36.4	4.9	3.1
	May	53.1	33.4	46.6	29.3	- 6.5	- 4.1
Harper	April	116.3	56.0	110.8	53.4	- 5.5	- 2.6
	June	116.3	56.0	117.8	56.8	1.5	0.7
Harvey	April	55.0	39.3	55.3	39.5	0.3	0.2
	June	55.0	39.3	42.2	30.2	-12.8	- 9.1
Kingman	April	97.0	43.3	113.7	50.8	16.7	7.5
	May	97.0	43.3	124.8	55.8	27.9	12.4
Kiowa	April	51.3	27.5	43.3	23.2	- 8.0	- 4.3
	May	51.3	27.5	45.6	24.4	- 5.6	- 3.0
Pratt	April May	82.6 82.6	43.7 43.7	91.3 80.5	48.3 42.6	8.8	4.6 - 1.1
Sedgwick	April	105.3	40.7	71.0	27.5	-34.3	-13.3
	June	105.3	40.7	117.3	45.4	12.0	4.6
Sumner	April June	196.9 196.9	64.3 64.3	217.0 195.8	70.9 63.9	$\begin{array}{c} 20.1 \\ -1.1 \end{array}$	6.6 - 0.4

acquisition can + considered together. The results also mean that a best date for Landsat coverage cannot be recom-, mended from this study.

### 5.1.2 Effect of Data Analyst

Since there was no significant date effect, the effect of analysts on the classification performance could be considered. This was a nested design with counties appearing within analysts. Three analyses were run: (1) all counties (2) all local counties, and (3) all nonlocal counties. Each result showed that the analyst effect was nonsignificant at any reasonable  $\alpha$  level when considering either proportion or area estimates. Since all analysts used similar methods, no inferences can be made about methodology; but it can be concluded that individual analysts did not introduce a bias in the results.

## 5.1.3 Effect of—Local vs. Nonlocal Recognition

One of the major problems encountered in the LACIE has been to develop a means for successfully extending training statistics from a training segment to "recognition" segments. In our investigation a different methodology involving stratification of counties into groups having similar characteristics and developing training statistics from throughout the training county was used. To determine if this method was satisfactory for classifying several counties the effect of local vs. nonlocal classification was tested. For proportion



estimates, the difference became apparent at the 20% significance level. For area estimates, however, the difference was significant for any  $\alpha$  larger than 0.10. Our conclusion is that there was some difference in performance between local and nonlocal counties; the amount of wheat was overestimated in local counties and underestimated in nonlocal counties; but, on the average, nonlocal recognition counties were closer to SRS estimates than the local recognition counties. It can probably be concluded that this factor did not have a strong influence on the overall results.

### 5.1.4 Effect of Interaction Between Dates and Locality

In the South Central Crop Reporting District, there appeared to be an interaction between date of the Landsat coverage and locality. Since the sample size was too small to draw any inference, a plot was made to examine this effect for the entire state. The interaction that was present in the South Central district analysis was not present over the entire state, although other factors which may have affected the accuracy were ignored. There is no good test on the significance of this interaction since variance estimates from the SRS are not available.

### 5.2 Landsat Classification Results

The Landsat classification results include the training

and test field performances; estimates of the area and proportion of wheat for the state, districts, and counties; comparisons of the Landsat estimates to USDA/SRS estimates; and evaluation of the accuracy and precision of the Landsat estimates. In addition regression estimates of wheat area and proportion in two districts for which Landsat data was not available are presented.

#### 5.2.1 Classification Accuracy

Classification accuracy was determined by the test field or training field performance matrices. The training field classification performance for all local recognition counties is given in Table 10. The test field performance is given in Table 11 for those counties which had test fields. The accuracy of the classification as assessed by training fields is not significantly different from that found by measuring test field performance. The overall classification performances are generally 85% or higher, an indication that the classification should result in accurate area estimates.

Since the classification performance of test (or training) fields was used to correct for classification bias in the area estimates, a plot was made of the absolute value of the bias correction of the Landsat results and the overall classification accuracy to show the relation between them (Figure 20). The simple correlation between these two variables is r = -0.80. The amount the Landsat estimates were adjusted

Table 10. Classification accuracy of training fields in Kansas.

	CLASSIFI	CATION AC	ACCURACY (%)		
COUNTY	WHEAT	OTHER	OVERALL		
CHEYENNE	87.8	99.0	91.8		
GRAHAM	84.3	87.2	86.1		
NORTON	93.7	87.0	89.5		
SHERMAN	70.3	97.5	89.5		
CLOUP	85.1	81.9	83.0		
OSBORNE	95.4	98.6	97.4		
OTTAWA	99.3	99.5	99.3		
SMITH	88.3	87.0	87.2		
GREELEY	82.7	93.8	90.0		
NESS	95.7	89.8	91.3		
TREGO	76.8	77.1	77.1		
WALLACE	51.7	97.7	90.0		
BARTON	95.3	83.7	87.8		
MCPHERSON	99.5	98.8	99.1		
RUSSELL	95.0	92.2	93.5		
SALINE	72.3	92.7	82.5		
FINNEY	97.0	94.5	95.4		
FORD	94.9	98.8	97.4		
HAMILTON	<b>75.</b> 3	55.5	61.9		
HASKELL	96.4	98.8	97.8		
HODGEMAN	86.3	79.3	81.3		
SEWARD	97.8	98.2	98.0		
STANTON	66.8	62.9	63.€		
BARBER	96.3	99.7	98.1		
HARVEY	98.1	93.7	95.∤5		
PRATT	99.8	94.8	97.0		
STAFFORD	94.4	98.5	96.4		
SUMNER	93.4	95.3	94.3		
ALLEN	94.2	94.5	94.4		

Table 11. Classification accuracy of test fields in Kansas.

	CLASSÍÍ	ICATION .	ACCŪRACY (%)
COUNTY	WHEAT	OTHER	OVERALL
SHERMAN	75,4	89.0	85.0
GREELEY	84.8	93.0	89.9
TREGO	86.7	81.1	82.4
SALINE	83.5	94.5	87.5
FORD	93.7	97.0	95.7
HAMILTON	94.2	78.4	82.5
HODGEMÂN	89.4	77.7	80.9
STANTON	62.5	79.1	75.5
BARBER	92.7	88.8	90.4
HARVEY	93.6	98.2	95.6
PRATT	92.7	95.6	93.8
STAFFORD	99.5	93.4	96.0
SUMNER	92.6	89.2	91.2
ALLEN	95.3	89.7	90.7

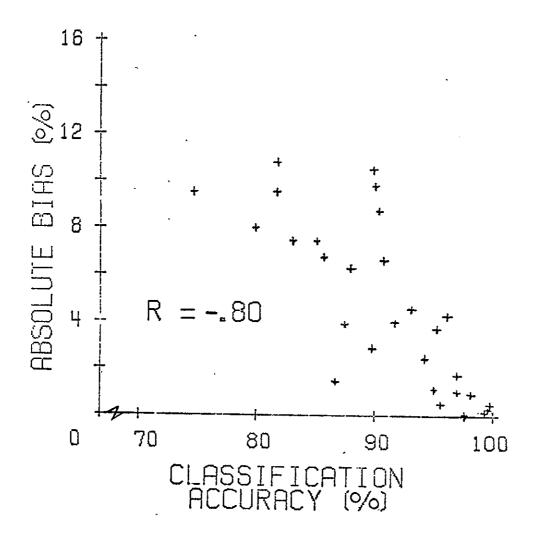


Figure 20. The relationship of the magnitude of the calculated bias correction to overall classification accuracy.

depends primarily upon the classification accuracy, but also on the estimated proportion of wheat in the county. The graph clearly shows that high classification performance is desirable to reduce the need for classification bias correction. High classification performance for each individual cover type is also a desirable attribute.

#### 5.2.2 Classification Bias Correction

To evaluate the consistency and usefulness of the bias correction, a subset of Kansas counties was examined. This was not a random sample of Kansas counties as the first completed counties were used, but it was considered to be representative enough and large enough to determine: (1) if the accuracy achieved by the estimates which used training field performance matrices to calculate the bias is different from that achieved when test field performance matrices are used, (2) if error matrices can be extended to nonlocal recognition counties, and (3) whether correction for the bias increases the accuracy of the estimates by decreasing the difference from the SRS estimates.

To determine if the accuracy achieved by the estimates for which training field performance matrices were used to calculate the bias is different from that achieved when test field performance matrices were used, the variable considered was the difference between Landsat and SRS estimates. The test performed was a two-sample t-test for difference in the

means between those counties for which training fields were used and those counties for which test fields were used to calculate the biases. The results were nonsignificant at the 25% significance level. It can be concluded that when test field performance is not available, the bias can be calculated by using the error performance matrix from the training fields.

Nonlocal recognition counties present another problem because there is no reference data from which a classification performance matrix can be obtained. Since statistics for the classification were extended from another county, it also seemed reasonable to extend the error matrix from the same county. To determine the validity of this extension, differences of Landsat estimates from SRS estimates for local counties were tested against the differences from SRS for nonlocal counties. This was accomplished by t-tests and the results showed that there was no difference ( $\alpha = 0.25$ ) between the closeness of Landsat estimates to SRS for corrected local counties and for corrected nonlocal counties. It, therefore, seemed reasonable to calculate the bias correction for nonlocal recognition counties by the extension of an error matrix.

Two t-tests were used for quantitative evaluation of the bias correction. For local recognition counties, the corrected estimates for proportions and areas did not differ from the SRS estimates at the 25% significance level. On the other hand, the uncorrected estimates did differ from SRS estimates at the 25% level, indicating that correction for the bias brought Landsat estimates closer to the SRS hectares harvested. Hence, all the local recognition counties were corrected for bias by the method previously described.

For the nonlocal recognition counties, the bias correction also brought the Landsat estimates closer to the SRS estimates. There was a significant difference ( $\alpha=0.001$ ) from SRS in both proportion and area of wheat for the uncorrected estimates while the corrected estimates were not significantly different from the SRS estimates even at  $\alpha=0.25$ . Therefore, all nonlocal county estimates were also corrected for classification bias.

In summary, we concluded that correcting for the bias is worthwhile since the difference of the corrected Landsat estimates from the SRS estimates is nonsignificant. Correction for the bias seems to be consistent between counties having test performance matrices and counties having only training performance matrices and is also consistent in extending error matrices to nonlocal counties. The same results were obtained for this part of the analysis regardless of whether the variable considered was proportion or area of wheat.



#### 5.3 Wheat Area and Proportion Estimates

The estimates of hectares and proportions from the Landsat classifications on a county-by-county basis are presented in Table 12. Estimates for both proportion and area of wheat are given as the uncorrected and bias-corrected values. The values used in the statistical analysis were always the bias-corrected estimates.

# 5.3.1 Correlation of Landsat and USDA/SRS Estimates of Area and Proportion of Winter Wheat

The SRS estimates for proportion and area of wheat harvested are presented in Table 13 along with the corresponding Landsat estimates and their differences. The proportion and area estimates obtained from the Landsat classification are highly correlated with the USDA/SRS estimates. The correlation between Landsat and SRS wheat harvested proportions is  $r = 0.77 \pm 0.05$  (Figure 21), while the correlation between Landsat and SRS wheat area estimates is  $r = 0.80 \pm 0.04$  for harvested estimates (Figure 22). The correlation values are presented in standard error form which represents approximately a 68% confidence interval. These intervals are not exactly symmetric, but the furthest boundary has been presented here for simplicity [11].

## 5.3.2 Accuracy of Landsat Estimates

The accuracy of Landsat estimates of the area and proportion of wheat can be assessed at three levels: state,

Table 12. Uncorrected and bias-corrected Landsat estimates of hectares and proportions of wheat in Kansas.

	LAND UNCORRI ESTIM	ECTED	CORF	IDSAT RECTED - MATES
COUNTY	HECTARES	PROPORTION	HECTARES	PROPORTION
	(000)	(%)	(000)	(%)
NORTHWEST DISTRICT CHEYENNE DECATUR GRAHAM NORTON RAWLINS SHERIDAN SHERMAN THOMAS TOTAL	93.5 55.7 59.6 70.1 69.0 79.7 46.8 45.6	35.1 33.8 325.8 304.7 17.1 16.5 25.8	82.6 31.4 44.8 50.3 76.2 53.1 25.8 22.6	31.0 13.5 19.4 22.1 27.3 23.0 9.4 8.2
	72.00	,	,	2.02
NORTH CENTRAL DISTRICT CLAY CLOUD JEWELL MITCHELL OSBORNE OTTAWA PHILLIPS REPUBLIC ROOKS SMITH WASHINGTON	37.57 71.8 83.4 83.4 784.9 681.1 70.1	22.399.19 22.8.199.6.0 143.9.6.0 135.49 136.49 136.49 136.49	55077596231 803.45577596231	21.2 31.2 46.7 34.6 72.8.7 28.2 31.4 231.4 318.3
TOTAL	688.2	29.9	575.0	25.0
WEST CENTRAL DISTRICT GOVE GREELEY LANE LOGAN NESS SCOTT TREGO WALLACE WICHITA	786 53.6 74590 865 865 865 8368	27.0 41.3 41.0 16.2 32.0 32.1 36.6 31.2	33.15 890.5 78.5 771.2 650.3 661.3 58.4	11:9 44:1 38:2 25:4 34:8 25:0 31:1
TOTAL	610.7	29.5	578.6	28.0
CENTRAL DISTRICT  BARTON DICKINSON ELLIS ELLSWORTH LINCOLN MCPHERSON MARION RICE RUSH RUSSELL SALINE	120-6 84-9 117-3 61-35 104-2 69-5 105-3 126-6 75-6	8339929 538023448679 538023448679 548679 548679	107.45 108.23 108.35 108.52 10	934698 4468 • • 698 477 • • 698 2847 • • 695 774 • • 4
TOTAL	994.9	42.8	956.4	41.2



Table 12. (continued)

	LAND UNCORR ESTIM	SAT ECTED ATES	CORP	IDSAT . RECTED MATES
COUNTY	. HECTARES	PROPORTION	HECTARES	PROPORTION
,,	(000)	(%)	.( 000 )	(%)
SOUTHWEST DISTRICT				•
CLARK FINNEY	30.5 148.5	12.0 44.0	25.9 143.1 71.7	10.2 42.4 25.5
FORD GRANT	73.4 39.0	76.1	9.8.	25.5 6.6
GRAY HAMILTON	59.4 138.5	26.5 26.4 53.9	60 - I	26.7 44.5
HASKELL HODGEMAN	30.0	20.6 51.4	114.3	20.6 43.4
KEARNEY	114.5 48.5 20.7 55.2	22.0	96.7 0.8	7 <b>2.</b> 4
MEADE MORTON	55.2	8 • 2 2 9 • 4	14.4 37.9 34.2	20:2
SEWARD STANTON-	36.2 63.8	21.9 36.4	34•2 47•3	20.2 20.7 20.7 27.0
STEVENS	61.6	32.6	28.3	13.0
TOTAL	920.7	30.0	715•4	23.3 <sup>c</sup>
SOUTH CENTRAL DISTRICT				
BARBER COMANCHE	88.5 44.0	29.8 21.2	89.4 46.3	30 · 1 22 · 3
EDWARDS HARPER	44.4 114.3	29 • 8 21 • 2 27 • 9 55 • 1	46.6 117.8	29.3 56.8
HARVEY KINGMAN	47.7 118.8	34 · 1 53 · 1	42.2 124.8	30.2 55.8
KIOWA Pawnee	43.4 77.3	23.2 39.8	45.6	24.4 35.4
PRATT RENG	76.8 123.3	40.6 37.9	68.7 80.5	42.6 33.3
SEDGWICK STAFFORD	116.6	45.1	108.3 117.3 75.0	45.4
SUMNER	83.9 187.8	40.8 61.3	195.8	36.5 63.9
TOTAL	1166.8	40.2	1158.3	40.0
SOUTHEAST DISTRICT	5# A			
ALLEN BOURBON	25.9 25.5 38.6	19.8 15.4	14.9 10.2	6.2
BUTLER ÇHAUTAUQUA	38.6 23.5	$\frac{10.3}{14.1}$	15.8 0.0	4.2 0.0
CHEROKEE COWLEY	34.3 53.3	22.5 18.1	22.1 43.0	11.4 6.2 4.2 0.0 14.5 14.6
CRÁWFÚRD ELK	334.3 234.3 3534.9 277.	16.1 16.7	10.8	7.0° 0.0
GREENWOOD LABETTE		20.i 20.3	0.0 20.4	0.0 12.0
MÔNTGCMERY NEOSHO	34.5 57.2 24.2	34.0	23.2 10.4	13.8
WILSON WOODSON	57.6 55.7	15.9 38.7 43.7	33.5	22.5 .29.2
TOTAL	542.9	42.7 20.3	38.1 242.4	.29•2 9•1
	- · <del>- ·</del>	<u> </u>		· • •
STATE TOTAL	5444.2	31.4	4612.9	26.6

Table 13. Comparison of USDA/SRS wheat harvested estimates and bias-corrected Landsat estimates of area and proportion of wheat in Kansas.

	USDA HARV	/SRS ESTED	LANDSAT CLASSIFICATION		DIFFERENCE FROM SRS	
COUNTY		PROPORTION			HECTARES	PROPORTION
	(000)	(%)	.(000).	(%)	(000)	(%)
NORTHWEST DISTRED CHEYENNE CHEYENNE DECATUR GRAHAM NORTON RAWLINS SHER IDAN SHERMAN THOMAS	61.0 48.6 44.2 42.3 60.3 50.2 73.1 90.4	22.9 20.9 19.1 18.5 21.6 21.7 26.7	8140.21 8140.21 8140.21 8140.21 8140.21	313.05 133.4 122.3 223.4 23.4 8.2	21.7 -17.2 0.6 8.0 15.9 3.0 -47.4 -67.8	8 • 1 -7 • • 4 -7 • • 5 -7 • • 5 -17 • • 3 -124 • •
TOTAL	470-1	23.3	386.8	19.2	-83.3	<del>-4.</del> 1
NORTH CENTRAL DE CLAY CLOUD JEWELL MITCHELL OSBORNE OTTAWA PHILLIPS REPUBLIC ROOKS SMITH WASHINGTON	STRICT 45.0 58.1 56.4 717.2 56.3 35.8 47.1 53.6 41.0	26.8 21.0 24.0 24.0 24.0 24.0 24.0 24.0 24.0 24	36.5 19.0 186.7 80.5 172.6 152.6 152.6 152.6 152.6 152.6 152.6 152.6 152.6 152.6 152.6 152.6 152.6 152.6 152.6 152.6 152.6 154	217 312 81 467 287 287 284 243	-8.5 -0.6 -37.4 15.6 22.8 -12.7 -18.0 5.5 18.7 1.1	139488870165
TOTAL	578.0	25.1	575.0	25.0	-3.0	-0.1
WEST CENTRAL DIS GOVE GREELEY LANE LOGAN NESS SCOTT TREGO WALLACE WICHITA	5TRICT 56.52 755.1 64.0 74.7 58.2 35.0 56.1	255.65.07 225.61.13 2222214.89	33.5 890.9 78.5 71.2 650.3 61.3 58.4	11.9 142.6 225.4 225.8 225.8 225.8 225.8 226.8	-23.4 17.8 14.5 -3.5 -7.5 10.5 26.3 2.4	451229513 -8835-34-45111
TOTAL	521.6	25+2	578.6	28.0	57.0	2.8
CENTRAL DISTRICT BARTON DICKINSON ELLIS ELLSWORTH LINCOLN MCPHERSON MARION RICE RUSH RUSSELL SALINE	72.38 5523.86 5539.61 55395.89 658.97 746.0	76516020984 223883624984 4322244323	1072 912 1082 535 1035 1035 1035 1035 1348 9348 828	933469860584 716.69860584 716.69860584 716.69860584	119.50 10.35 10.63 10.34 159.19	5.7. 22.9.63 1.9.60 1.00 9.00
TOTAL	769.7	33.1	956.4	41.2	186.7	8.1

Table 13. (continued)

	USDA HARV	/SRS ESTED	CLASSI	NDSAT FICATION	DIFFE FROM	RENCE SRS.
СОЏИТҮ	HECTARES	PROPORTION	HECTARES	PROPORTION	HECTARES	PROPORTION
	(000)	(%)	(000)	(%)	(000)	(%)
SOUTHWEST DISTRIC CLARK FINNEY FORD GRANT GRAY HAMILTON HASKELL HODGEMAN KEARNEY MEADE MORTON SEWARD STEVENS		17.4.9 127.4.61 234.7 2314.7 2314.7 244.4 214.8		10.456.7566.456.44566.4403.447.2700.000000000000000000000000000000	- 188.994 - 188.994 - 230.06 - 151.51 -	35505 -4.8.5 -4.8.5 -1.8.5
TOTAL	783.9	25.6	715.4	23.3	-68.5	-2.3
SOUTH CENTRAL DIS BARBER COMANCHE EDWARDS HARPER HARVEY KINGMAN KIOWA PAWNEE PRATT RENO SEDGWICK STAFFORD SUMNER	STR I CT 69 . 1 . 43 . 4 . 53 . 1 156 . 3 . 557 3 . 571 . 56 . 146 . 3 . 766 . 9	39.40335970733 223553425970733 42763.4450	89.4 46.6 117.8 124.8 45.6 68.7 108.3 117.3 195.0 195.0	1338284463459 0296	203.05 -6.55 -12.58 -27.68 -27.68 -28.00 -21.61	6.8 1.4 -4.1 0.7 -9.1 12.4 -3.0 -1.4 -1.7 4.68 -0.4
TOTAL	1164.5	40.2	1128-3	40.0	-0.2	0.2
SOUTHEAST DISTRI ALLEN BOURBON BUTLER CHAUTAUQUA CHEROKEE COWLEY CRAWFORD ELK GREENWOOD LABETTE MONTGOMERY NEOSHO WILSON WOODSON	11.4 17.5 42.3 88.9 18.3 10.9 20.8 20.8 21.1 21.7 267.1	8.75335 4.335 11.52.80 12.753337359 12.75337359 14.50 10.0	14.9 15.8 15.8 02.1 10.8 10.0 20.4 23.8 10.0 20.4 23.3 38.1 242.4	11.4 6.2 9.0 14.5 14.6 7.0 0.0 12.8 13.8 22.5 29.2	3.57 -26.6 -8.9 -21.3 -0.0 -6.9 -6.9 -0.4 -3.7 12.0 30.3	2.6 1.6 1.6 1.7 1.6 1.7 1.6 1.7 1.6 1.7 1.6 1.7 1.7 1.6 1.7 1.6 1.7 1.7 1.7 1.7 1.7 1.7 1.7 1.7 1.7 1.7
STATE TOTAL	-					

# PROPORTION OF WHEAT

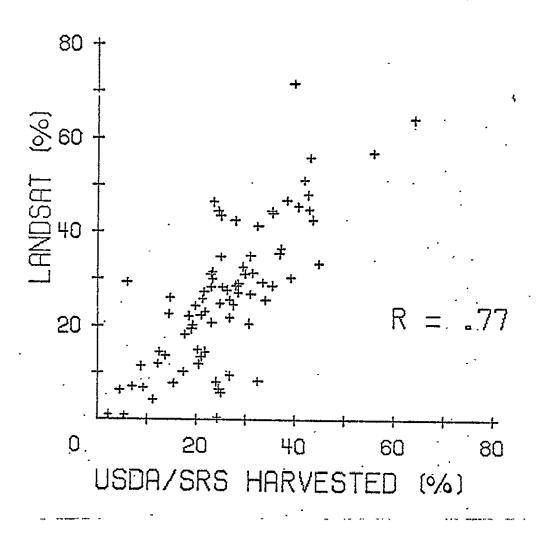


Figure 21. The correlation of Landsat and USDA/SRS estimates of the proportion of winter wheat in Kansas counties.

# AREA OF WHEAT

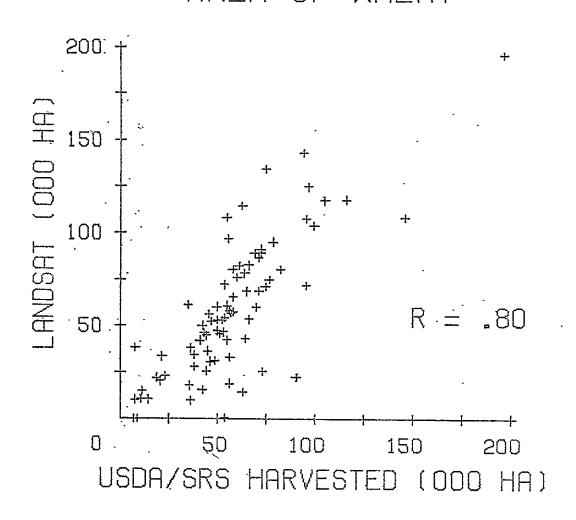


Figure 22. The correlation of Landsat and USDA/SRS estimates of the area of winter wheat in Kansas counties.

district, and county. A summary of the results at these three levels, including comparisons with the corresponding SRS estimates, is shown in Table 14. It should be noted that in comparing Landsat to SRS figures that the SRS figures are also estimates (and, thus subject to sampling error). The accuracy of the SRS estimates is greatest at the state level and least at the county level.

In tests of the accuracy of Landsat estimates at the state level, a large  $\alpha$  was used to reduce the possibility of claiming that Landsat estimates were the same as SRS estimates when, in fact, they were not. T-tests were performed to determine if there was a significant difference between Landsat and SRS estimates. At the 25% significance level, there was no difference in the proportion or area of wheat.

At the crop reporting district level there was no significant difference in Landsat and SRS estimates of proportion or area of wheat except in the Central CRD. In the Central CRD, wheat was overestimated for every county in relation to the SRS estimates, creating a bias in the CRD estimate. However, all the county estimates were close to the SRS estimates except for two counties which accounted for most of the difference. The Central CRD is not the "worst" CRD when considering relative difference or average absolute difference from SRS as a measure of comparison between crop reporting districts (Table 15). On the whole,

Table 14. Summary of USDA/SRS and Landsat estimates of area and proportion of wheat in Kansas.

.	Area			Proportion			
Region	USDA/SRS	Landsat	Difference	USDA/SRS	Landsat	Difference	
	(00	0 Hectares)			(%)		
State	4555	4613	58	26.2	26.6	0.4	
District		,	,				
Northwest	470	387	- 83	23.3	19.2	-4.1	
North Central	578	575.	- 3	25.1	25.0	-0.1	
West Central	522	579	57	25.2	28.0	2.8	
Centra1	770	956	187	33.1	41.2	8.1	
Southwest	784	715	- 68	25.6	23.3	-2.3	
South Central	1164	1158	- 6	40.2	40.0	-0.2	
Southeast	267	242	- 25	10.0	9.1	-0.9	
Counties							
(Median)	55.0	53.4	0.6	24.85	26.25	0.4	

Table 15. Relative difference and average absolute difference between Landsat and SRS estimates for districts and state.

District	Landsat Estimate	· Difference from SRS	Relative Difference	Average Absolute Difference
,	(000 Ha)	(000 Ha)	. (%)	(000 Ha)
Northwest	386.8	- 83.3	-21.5	22.7
North Central	575.0	- 3.0	- 0.5	13.8
West Central	578.6	57.0	9.9	12.3
Central	956.4	186.7	19.5	17.0
Southwest	715.4	- 68.5	- 9.6	25.4
South Central	1158.3	- 6.2	<b>-</b> 0⋅5	10.4
Southeast	242.4	- 24.7	-10.2	9.2
State	4612.9	58.0	1.3	• -

Landsat estimates were fairly close to SRS proportion and area estimates on a crop reporting district basis.

No statistical tests could be performed for differences from SRS estimates on a county-by-county basis because SRS does not calculate county variance estimates. Similarly, confidence limits cannot be placed around the SRS estimates. However, if the standard deviation of the SRS proportion estimates is assumed to be at least 10% at the county level, then 89% of the Landsat estimates were within a 90% confidence interval. For further comparison of Landsat and SRS county estimates, 49% of the counties were within ±5% (absolute difference) of SRS, 81% were within ±10%, and 88% were within +15%.

#### 5.3.3 Precision of Landsat Estimates

The second measure of the quality of an estimate is its precision which refers to the size of the deviations from its expected value obtained by repeated application of the sampling procedure. Using statistical theory, however, it is not necessary to repeatedly sample the population to determine the variance of an estimate.

The Landsat estimates are of a binomial nature since each point was classified as wheat or other. The variance of  $\hat{p}$  for a single county was calculated as:

$$v(\hat{p}) = \frac{\hat{p}(1-\hat{p})}{n-1}(1-f)$$

where  $\hat{p}$  is the proportion estimate after correction for the bias, n is the number of pixels classified in the county, and  $f=\frac{n}{N}$  where N is the total number of pixels in the county. The standard deviations for the districts and state were calculated considering the sample as stratified, but were approximately the same size as when calculated under the assumption of a systematic random sample throughout the CRD or state.

The standard deviations and coefficients of variation of the Landsat estimates are shown in Table 16. It can readily be seen that the standard deviations and the coefficients of variation (CV) are extremely small even at the county level. The CV of the SRS estimate of wheat acreage in the state of Kansas is 4%, compared to the CV of 0.06% for the Landsat estimate. The median CV of the Landsat county estimates is 0.60% which is smaller even than the 1.5% CV of the SRS national estimate of wheat acreage. Clearly the combined technologies of Landsat MSS data and computeraided classification methods provides a means to make very precise crop area estimates.

Table 16. Estimates of the standard deviations and coefficients of variation of Landsat estimates of wheat in Kansas,

	· · · · · · · · · · · · · · · · · · ·				· · · · · · · · · · · · · · · · · · ·
	Area Es	timate	Propor Estin		
County.	. Hectares	Standard Deviation		tandard eviatión	Coefficient of Variation
	(000 Ha)	(Ĥa)		(%)	(%)
Northwest Dist	rict				
Cheyenne Decatur Graham Norton Rawlins Sheridan Sherman Thomas	82.6 31.4 44.8 50.3 76.2 53.1 25.8 22.6	280.02 432.59 519.21 527.01 611.92 235.82 184.11 375.80	31.0 13.5 19.4 22.1 27.3 23.0 9.4 8.2	.1052 .1857 .2249 .2311 .2191 .1019 .0674 .1356	.33 1.38 1.16 1.05 .80 .44 .72 1.65
North Central Clay	District 36.5	448.79	21.7	.2668	1.23
Cloud Jewell Mitchell Osborne Ottawa Phillips Republic Rooks Smith Washington	57.5 19.0 86.7 80.7 53.5 17.9 ——52.6 72.2 56.3 42.1	566.41 359.92 567.23 604.48 233.98 354.56 517.03 689.56 561.17 621.13	31.2 8.1 46.7 34.7 28.6 7.7 28.2 31.4 24.3 18.3	.2004 .3074 .1532 .3058 .2598 .1249 .1523 .2775 .2997 .2425 .2691	.99 1.89
Total	575.0	1721.33	25.0	.0747	. 30
West Central I Gove Greeley Lane Logan Ness Scott Trego Wallace Wichita	33.1 89.5 60.9 78.5 71.2 65.4 60.3 61.3 58.4	199.98 265.57 289.98 278.04 271.56 243.08 249.10 249.47 236.34	11.9 44.1 32.6 28.2 25.4 34.9 25.8 26.0 31.1	.0714 .1309 .1555 .1000 .0969 .1297 .1067 .1057	.60 .30 .48 .35 .38 .37 .41
Total	578.6	763.55	28.0	.0369	.13

Table 16. (continued)

	Aŗea Es	timate	Propo Esti	rtion mate	Coefficient	
County	Hectares	Standard Deviation		tandard eviation	of Variation	
	(000 Ha)	(Ha)		(%)	(%)	
entral Distr	ict					
Barton	107.4	269.37	47.9		.25	
Dickinson	91.5	274.76	41.3	.1240	.30	
Ellis	108.2	284.36	46.4	.1219	.26	
<b>Ellsworth</b>	53.3	503.91	28.6	.2708	.95	
Lincoln	54.5	522.31	28.9	.2777	.96	
McPherson	103.9	283.67	44.8	.1223	. 27	
Marion	68.5	263.38	27.6	.1060	.38	
Rice	95.2	562.69	51.0	.3012	.59	
Rush	134.2	232.65	71.5	.1240	.17	
Russell	56.8	537.75	24.8	.2351	.95	
Saline	82.9	256.30	44.4.	.1374	.31	
Total	956.4	1277.74	41.2	.0550	.13	
Southwest Dis	strict				•	
Clark	25.9	182.06	10.2	.0714	· <b>.</b> 70	
Finney	143.1	783.49	42.4	.2323	. 55	
Ford	71.7	269.07	25.5	.0959	.38	
Grant	9.8	110.96	6.6	.0754	1.14	
Gray	60.1	552.52	26.7	.2454	.92	
Hamilton	114.3	308.61	44.5	.1200	.27	
Haskell	30.9	412.53	20.6	.2750	1.33	
Hodgeman	96.7	275.23	43.4	.1235	.28	
Kearney	0.8	43.31	0.4	.0196	4.90	
Meade	14.4	306.19	5.7	.1210	2.12	
Morton	37.9	205.85	20.2	.1096	.54	
Seward	34.2	433.69	20.7	.2619	1.27	
Stanton	47.3	217.81	27.0	.1244	.46	
Stevens	28.3	182.13	15.0	.0964	.64	
Total	715.4	1336.91	23.3	.0436	.19	

Table 16. (continued)

	Area Es	Propo: Estir		055:	
. County ,	Hectares	Standard Deviation		tandard eviation	Coefficient of Variation
	(000 Ha)	(Ha)		(%)	(%)
South Central Barber Comanche Edwards Harper Harvey Kingman Kiowa Pawnee Pratt Reno Sedgwick Stafford Sumner	89.4 46.3 46.6 117.8 42.2 124.8 45.6 68.7 80.5 108.3 117.3 75.0 195.8	291.83 219.97 213.44 265.85 209.98 278.11 216.33 244.64 252.87 312.23 297.32 295.20 311.55	30.1 22.3 29.3 56.8 30.2 55.8 24.4 35.4 42.6 33.3 45.4 36.5 63.9	.0983 .1061 .1341 .1281 .1501 .1243 .1160 .1261 .1339 .0960 .1150 .1435 .1018	.33 .48 .46 .23 .50 .22 .48 .36 .31 .29 .25 .39
Total	1158.3	954.06	40.0	.0329	.08
Southeast Dis- Allen Bourbon Butler Chautauqua Cherokee Cowley Crawford Elk Greenwood Labette Montgomery Neosho Wilson Woodson Total	14.9 10.2 15.8 0.0 22.1 43.0 10.8 0.0 0.0 20.4	138.02 113.60 147.35 0.00 162.31 224.81 122.77 0.00 0.00 156.22 166.20 115.64 187.84 194.02 532.05	11.4 6.2 4.2 0.0 14.5 14.6 7.0 0.0 12.0 13.8 6.8 22.5 29.2 9.1	.1055 .0686 .0394 .0000 .1067 .0764 .0792 .0000 .0922 .0988 .0760 .1263 .1486	.93 1.11 .94 .00 .74 .52 1.13 .00 .00 .77 .72 1.12 .56 .51 .22
State Total	4612.9	3089.32	26.6	.0178	.07

# 5.4 Regression Estimation for Wheat in Areas without Landsat Coverage

Usable Landsat data was not available for the Northeast and East Central Crop Reporting Districts; thus those districts were not analyzed. Since estimates of area and proportion of wheat in the counties were required, a prediction equation was formulated using the 80 counties which had been classified with Landsat data. The Landsat wheat estimates were written as a function of historical wheat production in the two previous years and acres in the county. The prediction equation derived by this procedure was:

$$\hat{y} = 10274.97 + 0.66 x_1 - 0.26 x_2 - 0.02 x_3$$

where  $\mathbf{x}_1$  is the acreage of wheat grown in a county in 1974,  $\mathbf{x}_2$  is the acreage of wheat grown in a county in 1973,  $\mathbf{x}_3$  is the number of acres\_in the county, and  $\hat{\mathbf{y}}$  is the "pseudo-Landsat" estimate in hectares. The  $R^2$  value for the regression equation was 0.65.

Regression is good for prediction only when the x values corresponding to the estimate to be predicted fall within the range of the x values used in deriving the equation. If this held true for a given county, the estimate was made from the prediction equation. If this did not hold true, the USDA/SRS wheat estimate from the previous year was used. The estimates are presented in Table 17.

Table 17. Regression estimates of area and proportion of winter wheat in counties for which usable Landsat data was not available.

<sup>\*</sup>Historical estimates used.

The estimates obtained were tested for differences from SRS estimates of proportion and area of wheat harvested on a crop reporting district basis. There were significant differences from SRS in both area and proportion estimates in both crop reporting districts. Estimation from regression consistently underestimated wheat as did the historical estimates. Regression seems a reasonable alternative if Landsat estimation cannot be done for a given county, but a significant decrease in the accuracy of the estimates is likely to occur.

## 6.0 CORN AND SOYBEAN IDENTIFICATION AND AREA ESTIMATION IN INDIANA

The second state selected for analysis was Indiana; corn and soybeans, the two major grain crops in the state, were selected for study. This section includes the results of the Landsat data classifications and analyses. As for Kansas, the material presented includes a discussion of the factors affecting classification performance, comparisons of USDA/SRS and Landsat estimates of the area and proportions of the crops of interest, and evaluations of the accuracy and precision of the Landsat estimates.

### 6.1 Analysis of Factors Affecting Classification Accuracy

The effects of several factors likely to influence the accuracy of the Landsat area and proportion estimates were investigated. These included: Landsat acquisition date, aerial photography acquisition date, and local vs. nonlocal training and classification. There are, of course, many additional factors such as field size, number of crops and cover types

present, uniformity of soils, and production practices, which may have also influenced the results, but which were beyond the scope of this investigation to pursue.

### 6.1.1 Effect of Landsat Acquisition Date

To study the effect of the date of Landsat coverage on the accuracy of the estimates, pairwise comparisons were made among three groups of dates (July, August, and September) without considering the effect of other factors. Different counties were in each group since all counties in Indiana were classified only once. The accuracy of an estimate was considered to be its closeness to the SRS estimate.

The estimates of the proportion and area of corn were significantly further from the SRS estimates ( $\alpha \geq 0.02$ ) using September Landsat data than either July or August data. For soybean proportion and area estimation, the effect of Landsat acquisition date was not significant.

Estimates made from July and August Landsat data were not significantly different in accuracy for either corn or soybeans; thus, either date could be recommended. However, the August estimates of both corn and soybeans were closer in average difference to the SRS estimates than were the July estimates. Similar results were obtained in the CITARS experiments in which corn and soybeans in six Indiana and Illinois test sites were classified throughout the growing season [5].

### 6.1.2 Effect of Aerial Photography Acquisition Date

Three groups of dates (July, August, and September) also existed for the aerial photography acquisition dates. Although the groups are the same as for the study of Landsat acquisition date, the counties within each group were not always the same since photographic acquisition was not necessarily coordinated with Landsat data acquisition. Considering performance as a function of photography acquisition date only for corn estimation, both July and August estimates were significantly closer to the SRS estimates than September estimates were. For soybean estimation, August estimates were significantly closer to the SRS estimates than were the July estimates, while not significantly closer than September estimates.

Even though there was not a significant difference in the accuracy of July and August estimates for corn or of August and September estimates for soybeans, the August estimates were closer to the SRS estimates in both cases. The best time for aerial infrared photography acquisition appears to be August, coinciding with the optimal time period for the Landsat data acquisition. In some cases, multidate photography proved useful for identifying corn and soybeans when individual acquisition dates were not acquired at a good time for interpretation.

#### 6.1.3 Effect of Local vs. Nonlocal Classification

The significance of the effect of local versus nonlocal classification depended upon the crop being estimated. Corn estimates were significantly better in nonlocal counties than in local recognition counties; an explanation of this unexpected result has not been identified. Soybean classification accuracy was not significantly affected by local versus nonlocal classification although local counties were closer to SRS estimates on the average.

#### 6.2 Landsat Classification Results

The Indiana results include training field classification performances, estimates of the area and proportions of corn and soybeans for 43 counties in four districts, comparisons of the Landsat and USDA/SRS estimates, evaluation of the accuracy and precision of the estimates, and regression estimates for counties for which Landsat data were not analyzed.

### 6.2.1 Classification Accuracy

Classification accuracy was determined for Indiana by the training field performance matrices. No test fields were used in Indiana since it was felt that additional training data would be more valuable than having test fields; comparison of classification accuracies of training and test fields in Kansas showed them to be not significantly different. The training

field classification performance for all local recognition counties is given in Table 18.

The training field classification performances are typically 75 to 85 percent. Although these accuracies are about 10 percent lower than obtained in Kansas, they would generally be considered adequate for making satisfactory area estimates provided a consistent bias was not present. As shall be shown in subsequent sections, the area and proportion estimates, particularly on a county basis, are not as accurate as might have been predicted from the training field classification performances. This is believed to be caused by a combination of two factors. First, the training performances are for "pure" pixels from the centers of fields; the area estimates, however, are made from samples including "mixed" or field boundary pixels. The proportion of pure pixels for Indiana fields which average only about 10 hectares in size is typically no more than 50 percent. Secondly, we encountered some difficulty in accurately identifying all fields as corn, soybeans, or other. Since positive identification of a field was required in order to use it for training, a significant number of fields representing several spectral classes was omitted from training. This would cause the training field classification performance to be biased upward.

### 6.2.2 Classification Bias Correction

Training field performance matrices were used to calculate

Table 18. Classification accuracy of training fields in Indiana.

		Classification	n Accuracy	(%)
County	Corn	Soybeans	Other	Overal
Benton	87.0	98.1	72.2	83.7
Lake	79.6	89.4	91.5	85.7
LaPorte	85.0	97.0	88.8	89.1
Newton	86.2	97.1	70.0	84.1
Pulaski	92.3	98 <b>.</b> 2'	85.8	91.6
Starke	92.3	98.2	. 85.• 8	91.6
White .	90.9	89.8	78.7	87.5
Fountain	88.6	91.9	79.18	86.1
Montgomery	84.6	89.8	81.2	85.6
Owen	87.2	64.0	94.2	84.1
Parke	88.6	91.9	79.8	86.1
Tippecanoe	98.3	90.9	86.9	92.5
Vigo	61.8	60.4	89.6	75.9
Warren	.95.3	94.4	92.2	93.9
Decatur	79.4	98.1	79.1	85.3
Grant	91.8	98.5	72.7 <sup>,</sup>	89.2
Hamilton	71.6	98.0	76.6	81.1
Hancock	85.1	99.1	84.8	90.4
Howard	71.6	98.0	76.6	81.1
Johnson	90.3	93.7	94.8	92.5
Madison	88.4	97.6	73.3	88.8
Shelby	90.3	93.7	94.8	92.5
Tipton	71.6	98.0	76.6	81.1
Fayette	90.5	90.9	85.1	88.5
Jay	73.5	88.5	81.5	83.6
Randolph	84.4	95.5	75.9	87.8
Wayne	88.1	94.7	823	88.3

the bias in the absence of test fields; the Kansas analysis had demonstrated this was feasible. Also following the results from the Kansas analysis, error matrices were extended to nonlocal recognition counties.

All crop estimates were corrected for the bias because this operation brought them closer to SRS estimates on the average. For soybeans, there was no significant difference at any reasonable  $\alpha$  level in the accuracy of corrected and uncorrected estimates. For corn estimates, however, corrected estimates were closer to SRS at the 20% significance level.

#### 6.3 Corn and Soybean Area and Proportion Estimates

Tables 19 and 20 present the results of the Landsat classifications on a county-by-county basis. Estimates for both proportion and area of corn and soybeans are given as the uncorrected and bias-corrected values. The values used in the statistical analysis were always the bias-corrected estimates.

## 6.3.1 Correlation of Landsat and USDA/SRS Estimates of Area and Proportion of Corn and Soybeans

Plots of the Landsat vs. SRS county estimates of corn and soybean area and proportions, along with correlation estimates, are shown in Figures 23-26. The two estimates are not as highly correlated as the Kansas estimates; three counties, however, accounted for much of the lack of correlation of the corn estimates. The Landsat estimates for corn are

Table 19. Uncorrected and bias-corrected Landsat estimates of hectares and proportions of corn in Indiana.

	Unco	rrected	Bias-Co	orrected
County	Hectares	Proportion	Hectares	Proportion
	(000)	(%)	(000)	(%)
Northwest Dist	rict			
Benton	53.5	50.5	53.6	50.6
Jasper	36.8	25.3	92.0	63.3
Lake	56.1	42.1	62.7	47.1
LaPorte	60.8	38.6	64.7	41.1
Newton	63.2	59.3	63.0	59.2
Porter	47.2	42.9	53.1	48.2
Pulaski	54.0	48.1	54.1	48.2
Starke	38.8	48.2	38.1	47.3
White	66.6	51.7	63.4	49.2
Tota1	477.0	44.2	544.7	50.4
West Central D:	istrict			
Clay	17.1	18.1	18.0	19.1
Fountain	45.9	44.6	42.2	41.0
Montgomery	60.8	46.3	62.2	47.4
Owen	23.2	23.3	19.2	19.2
Parke	50.1	42.9	44.4	38.0
Putnam	39.8	31.5	36.2	28.6
Tippecanoe	56.7	43.7	53.0	40.8
Vermillion	34.4	50.5	33.5	49.2
Vigo	20.2	18.8	21.7	20.2
Warren	38.0	39.9	35.9	37.6
Total	386.2	36.0	366.3	34.2

Table 19. (continued)

	Unco	rrected	Bias-Corrected		
County	Hectares	Proportion	Hectares`	Proportion	
	(000)	(%)	(000)	(%)	
Central Distri	ct				
Bartholomew	20.3	19.5	3.4	. 3.3	
Boone	19.6	17.7	5.6	5.1	
Clinton	17.1	16.2	2.4	2.3	
Decatur	38.5	40.2	37.3 31.0	38.9 28.4	
Grant	42.3 35.8	38.8 34.5	38.0	36.6	
Hamilton Hancock	29.6	34.5 37.5	30.6	38.7	
Hendricks	41.6	38.5	48.2	44.6	
Howard	31.8	41.9	39.5	52.0	
Johnson	32.1	39.3	32.6	39.9	
Madison	51.3	43.7	46.7	39.8	
Marion	28.5	27.4	15.1	14.5	
Morgan	19.3	18.3	15.3	14.5	
Rush	38.6	36.4	38.8	36.6	
She1by	51.6	48.7	54.0	51.0	
Tipton	26.8	39.7	33.7	49.9	
Tota1	524.8	33.2	472.2	29.9	
East Central D	istrict	,			
Blackford	13.2	30.4	15.2	35.2	
Delaware	41.8	40.5	43.9	42.6	
Fayette	15.3	27.5	13.3	23.8	
Henry	25.9	25.0	23.8	23.0	
Jay	27.3	27.3	30.9	30.9	
Rando1ph	46.8	39.5	49.0	41.4	
Union	13.9	31.9	12.4	28.4 21.9	
Wayne	26.5	25.3	23.0		
Total	210.7	31.3	211.5	31.4	
State	1598.7	36.3	1594.7	36.2	

Table 20. Uncorrected and bias-corrected Landsat estimates of hectares and proportions of soybeans in Indiana.

_	Unco	rrected	Bias-Corrected		
County	Hectares	Proportion	Hectares	Proportion	
	(000)	(%)	(000)	(%)	
Northwest Disti	rict		•		
Benton	22.6	21.3	20.3	19.2	
Jasper	22.8	15.7	22.4	15.4	
Lake	. 24.0	18.0	22.1	16.6	
LaPorte	32.9	20.9	32.9	20.9	
Newton	13.5	12.7	12.4	11.6	
Porter	22.6	20.5	21.4	19.4	
Pulaski	32.3	28.8	32.6	29.1	
Starke	18.3	22.7	18.5	22.9	
White	27.4	21.3	26.4	20.5	
Total .	216.4	20.0	209.0	19.3	
West Central Di	istrict				
Clay	19.4	20.6	26.0	27.6	
Fountain	12.7	_ 12.3	11.6	11.3	
Montgomery	23.1	17.6	24.4	18.6	
Owen	12.5	12.5	15.6	15.6	
Parke	11.1	9.5	9.3	8.0	
Putnam	16.9	13.4	21.1	16.7	
Tippecanoe	23.9	18.4	23.4	18.0	
Vermillion `	8.0	11.8	7.5	11.0	
Vigo	22.2	20.6	29.6	27.5	
Warren	11.5	12.1	12.2	12.8	
Total	161.3	15.0	180.7	16.9	

Table 20. (continued)

	Unco	rrected	Bias	-Corrected
County	Hectares	Proportion	Hectares	Proportion
·	(000)	(%)	(000)	(%)
Central Distri	, ct			
Bartholomew	15.7	15.1	15.7	15.1
Boone	38.4	34.7	38.6	34.9
Clinton	37.0	35.1	37.2	35.3
Decatur	15.5	16.2	15.6	16.3
Grant	22.8	20.9	21.1	19.3
Hamilton	29.7	28.6	29.3	28.2
Hancock	23.1	29.2	21.8	27.6
Hendricks Howard	30.7	28.4	30.1	27.9
nowara Johnson	22.5 33.3	29.6	22.0	29.0
Madison	30.4	40.8 25.9	34.9	42.8
Marion	12.3	11.8	28.1 11.7	23.9 11.2
Morgan	9.8	-9.3	11.3	10.7
Rush	29.8	28.1	30.9	29.2
She1by	32.2	30.4	33.4	31.5
Tipton	23.5	34.8	23.3	34.4
Totál	406.7	25.7	405.0	25.6
East Central Di	istrict			
Blackford	12.7	29.3	11.6	26.7
Delaware	37.3	36.2	33.0	32.0
Fayette	12.4	22.2	12.3	22.1
Henry	28.6	27.6	24.3	23.4
Jay	34.6	34.6	33.3	33.3
Randolph	43.7	36.9	38.8	32.8
Union	6.7	15.3	6.2	14.3
Wayne	16.5	15.7	10.0	9.5
Total	192.5	28.6	169.5	25.2
State	97.6.9	22:2	964.2	21.9

## PROPORTION OF CORN.

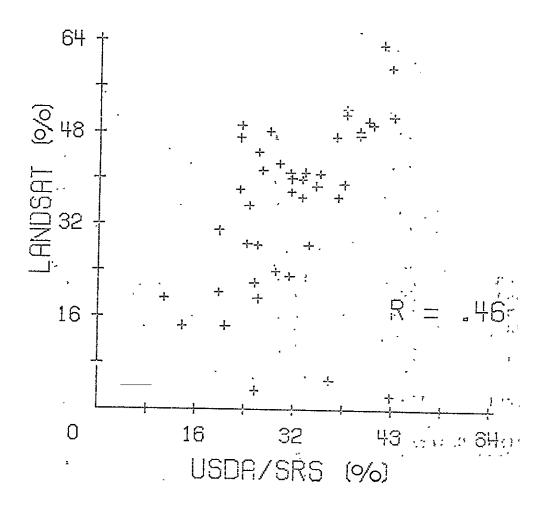


Figure 23. The correlation of Landsat and USDA/SRS estimates of the proportion of corn in Indiana counties.

## AREA OF CORN

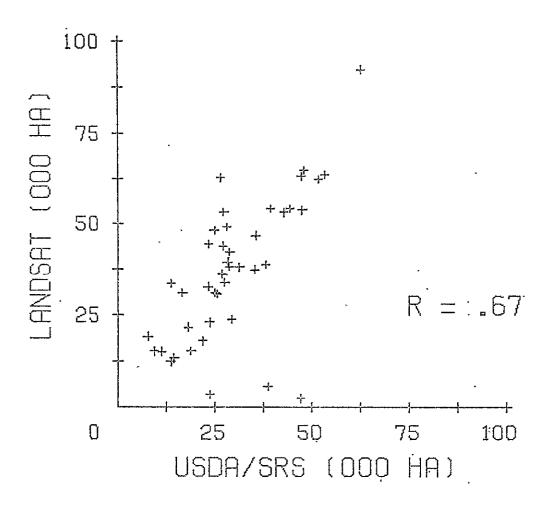


Figure 24. The correlation of Landsat and USDA/SRS estimates of the area of corn in Indiana counties.

## PROPORTION OF SOYBEANS

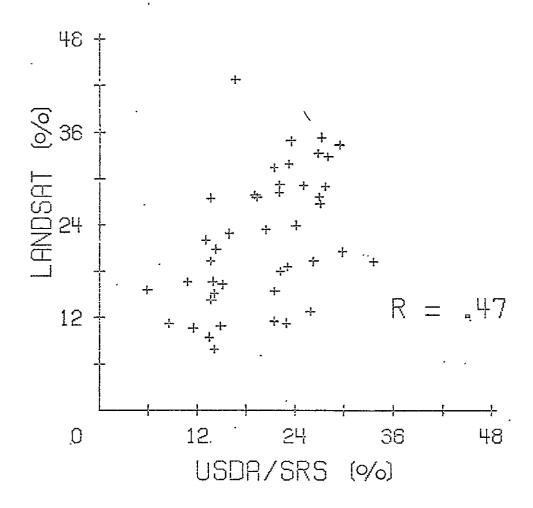


Figure 25. The correlation of Landsat and USDA/SRS estimates of the proportion of soybeans in Indiana counties.

## AREA OF SOYBEANS

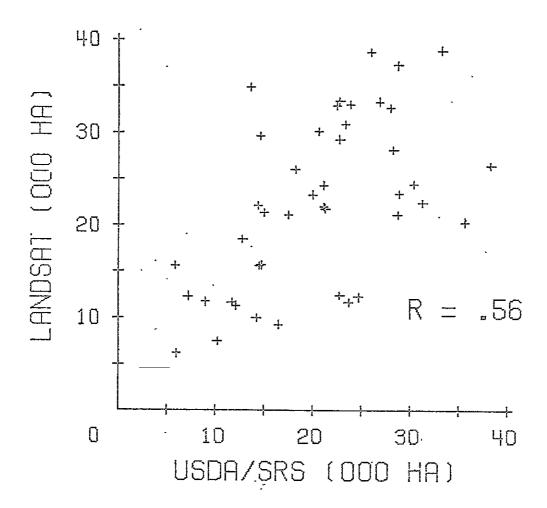


Figure 26. The correlation of Landsat and USDA/SRS estimates of the area of soybeans in Indiana counties.

consistently greater than the SRS estimates. On the other hand, the Landsat soybean estimates do not appear biased, but are clearly more variable than either the corn or Kansas wheat estimates.

More quantitative comparisons of the Landsat and SRS estimates at the county, as well as the district and "state" levels, are shown in Tables 21 and 22.

#### 6.3.2 Accuracy of Estimates

Only four of Indiana's crop reporting districts were estimated using Landsat classification methods. These four districts together make up a "pseudo" state estimate which was tested against an SRS "pseudo" state estimate. The Landsat corn proportion and area estimates were significantly different from the SRS estimates. The soybean estimates were closer to SRS estimates, but the differences became significant at the 20% level for both proportion and area estimates.

Assuming that the SRS estimates were unbiased in these crop reporting districts, the estimates derived from the Landsat classification were not as accurate as the SRS estimates.

Tests were also performed for differences from SRS estimates on a crop reporting district basis. In the Northwest and West Central Districts, corn estimates were significantly different from SRS, while soybean estimates were not significantly different. In the Central District, the reverse was found: corn estimates were not significantly different from

Table 21. Comparison of USDA/SRS corn estimates and biascorrected Landsat estimates of area and proportion of corn in Indiana.

	Proportion (%)	ı	Hectares (000)			
County	SRS	Landsat	Diff.	SRS	Landsat	Diff.
Northwest Dist	~; ~+					
	_	50.6	5.7	47.6	53.6	6.0
Benton	44.9		20.1	62.8	92.0	29.2
Jasper	43.2	63.3			62.7	36.1
Lake	20.0	47.1	27.1	26.6		
LaPorte	30.6	41.1	10.5	48.1	64.7	16.6
Newton	44.6	59.2	14.6	47.4	.63.0	15.6
Porter	24.8	48.2	23.4	27.3	53.1	25.8
Pulaski	39.4	48.2	8.8	44.2	54.1	9.8
Starke	35.6	47.3	11.7	28.7		9.4
White	41.6	49.2	7.6	53.5	63.4	9.8
Total	35.8	50.4	14.6	386.2	544°.7	158.5
West Central D	distric	:t			•	
Clay	23.1	19.1	- 4.0	21.8	18.0	- 3.8
Fountain	28.1	41.0	12.9	29.0	42.2	13.2
Montgomery	39.5	47.4	7.9	51.8		10.4
Owen	7.8	19.2	11.4	7.7	19.2	11.5
Parke	20.0	38.0	18.0	23.4	44.4	21.0
Putnam	21.3	28.6	7.3	26.9		9.3
Tippecanoe	33.0	40.8	7.8	42.8		10.2
Vermillion	20.1	49.2	29.1	13.7	33.5	19.8
Vigo	16.8	20.2	3.4	18.1	21.7	3.6
Warren	28.4	37.6	9.2	27.0	35.9	8.8
	•					
Total	24.4	34.2	9.8	262.2	. 366.3	104.1

Table 21. (continued)

•		Proportion (%)	on		Hectares (000)		
County	SRS	Landsat	Diff.	SRS	Landsat	Diff.	
Central Distri Bartholomev Boone Clinton Decatur Grant Hamilton Hancock Hendricks Howard Johnson Madison Marion Morgan Rush Shelby Tipton Total	ict	3.3 5.1 2.3 38.9 28.4 36.6 38.7 44.6 52.0 39.9 39.8 14.5 14.5 36.6 51.0 49.9 29.9	-19.5 -29.8 -42.5 2.0 5.4 6.4 6.2 21.6 14.7 11.4 9.6 3.7 -3.4 0.6 13.8 9.1 -0.1	23.7 38.6 47.2 35.3 25.1 31.4 25.7 24.9 28.3 23.3 35.5 11.3 18.9 38.1 39.4 27.6 474.3	3.4 5.6 2.4 37.3 31.0 38.0 30.6 48.2 39.5 32.6 46.7 15.1 15.3 38.8 54.0 33.7 472.2	-20.3 -33.0 -44.8 1.9 5.8 6.6 4.9 23.3 11.1 9.3 11.2 3.8 -3.6 0.7 14.7 6.1 -2.1	
East Central I Blackford Delaware Fayette Henry Jay Randolph Union Wayne Total State		35.2 42.6 23.8 23.0 30.9 41.4 28.4 21.9 31.4	13.7 16.2 - 2.2 - 5.3 14.2 17.7 - 2.9 - 0.6 7.3	9.3 27.2 14.5 29.3 16.7 28.1 13.6 23.6 162.3	15.2 43.9 13.3 23.8 30.9 49.0 12.4 23.0 211.5	5.9 16.7 - 1.2 - 5.5 14.2 21.0 - 1.2 - 0.6 49.2	

Table 22. Comparison of USDA/SRS soybean estimates and biascorrected Landsat estimates of area and proportion of soybeans in Indiana.

		Proportion (%)			Hectares (000)		
County	SRS	Landsat	Diff.	SRS	Landsat	Diff.	
Northwest Dist Benton Jasper Lake LaPorte Newton Porter Pulaski Starke White	rict 33.6 21.5 10.8 14.3 21.4 13.6 25.0 15.9 29.8	19.2 15.4 16.6 20.9 11.6 19.4 29.1 22.9 20.5	-14.4 - 6.1 5.8 6.6 - 9.8 5.8 4.1 7.0 - 9.3	35.6 31.3 14.4 22.5 22.8 15.0 28.0 12.8 38.3	20.3 22.4 22.1 32.9 12.4 21.4 32.6 18.5 26.4	-15.2 - 8.9 7.7 10.4 -10.4 6.3 4.6 5.7 -11.9	
Total	20.4	19.3	- 1.1	220.7	209.0	-11.7	
West Central D Clay Fountain Montgomery Owen Parke Putnam Tippecanoe Vermillion Vigo Warren	19.5 23.0 23.1 5.9 14.1 13.9 22.2 14.9 13.6 25.9	27.6 11.3 18.6 15.6 8.0 16.7 18.0 11.0 27.5	8.1 -11.7 - 4.5 9.7 - 6.1 2.8 - 4.2 - 3.9 13.9 -13.1	18.4 23.7 30.4 5.9 16.5 17.5 28.9 10.2 14.6 24.7	26.0 11.6 24.4 15.6 9.3 21.1 23.4 7.5 29.6 12.2	7.6 -12.1 - 5.9 9.7 - 7.1 3.6 - 5.5 - 2.7 15.0 -12.5	
Tótal	17.8	16.9	- 0.9	190.8	180.7	-10.1	

Table 22. (continued)

		Proportion (%)			Hectares (000)	
County	SRS	Landsat	Diff.	SRS	Landsat	Diff.
Central Distri	.c't	15.1 34.9 35.3 16.3 19.3 28.2 27.6 27.9 29.0 42.8 23.9 11.2 10.7 29.2 31.5 34.4 25.6	1.0 11.4 8.0 1.2 -7.0 6.2 0.6 8.8 1.2 26.1 -0.2 2.6 -0.9 7.1 10.0 4.9 4.9	14.7 26.0 28.8 14.4 28.7 22.8 21.3 20.6 21.1 13.6 28.3 9.0 12.2 23.4 22.8 20.0 327.7	15.7 38.6 37.2 15.6 21.1 29.3 21.8 30.1 22.0 34.9 28.1 11.7 11.3 30.9 33.4 23.3 405.0	1.1 12.6 8.4 1.2 7.7 6.5 0.5 9.5 0.9 21.3 - 0.3 2.7 - 0.9 7.5 10.6 3.3 77.3
East Central I Blackford Delaware Fayette Henry Jay Randolph Union Wayne Total	0istrict 27.1 23.2 13.0 20.4 26.9 28.1 13.7 13.5 21.5	26.7 32.0 22.1 23.4 33.3 32.8 14.3 9.5 25.2	- 0.4 8.8 9.1 3.0 6.4 4.7 0.6 - 4.0 3.7	11.7 23.9 7.2 21.1 26.9 33.3 6.0 14.2 144.3	11.6 33.0 12.3 24.3 33.3 38.8 6.2 10.0 169.5	0.2 9.1 5.1 3.1 6.4 5.5 0.3 - 4.2 25.2

SRS while soybean estimates were different. In the East Central District, both corn and soybean estimates differed significantly from SRS estimates at the 25% level.

In conclusion, compared to SRS, the Landsat estimates of corn area and proportion were consistently overestimated. This is attributed in part to the spectral similarity of corn to other cover types, particularly trees, as well as to factors mentioned earlier such as boundary pixels. Because the corn estimates, although biased, were correlated with the SRS estimates, a regression technique such as described by Wigton [26] might be effectively used if sufficient "ground truth" data were available to determine the magnitude of the bias. On the other hand, the large variation present in soybean estimates would make it infeasible to attempt such a correction. When aggregated, however, the soybean estimates were reasonably close to the SRS estimates.

One further factor, perhaps accounting for some of the differences in the Landsat and SRS estimates, is that the SRS county and district estimates used for comparison are preliminary and may be revised before the final estimates are published in 1977. This possibility was identified when 1974 estimates were examined for use in regression equations to predict crop areas in counties for which Landsat data were not analyzed.

In November 1976, revised 1974 county estimates of corn

and soybean acreages were published by SRS. At first glance, these estimates seemed to be different from the preliminary estimates. For prediction of crop acreages where historical data was used (either as an estimate or in a regression) the preliminary figures were used to simulate real-time estimation. However, in a test on a few counties, a regression equation using the revised estimates appeared to give better prediction for 1975.

The Landsat estimates for corn and soybeans did differ from the available SRS estimates which were preliminary.

Looking at the changes in the 1974 estimates, it seems possible that the SRS revised estimates may be enough different from the estimates used for comparison that the Landsat estimates may not differ (at least not so much) when compared to the revised figures. It is unfortunate, however, that the revised 1975 estimates will not be available until late in 1977.

To evaluate the difference between the preliminary and revised estimates on a county basis, the relative difference of the preliminary estimate from the revised estimate was calculated. These are presented for each crop and each county in Table 23. Relative differences were as great as 33.3%. This extreme figure occurred in a county with a very small corn and soybean production, but other large relative differences of 10 to 20% occurred where these crops were more important. The differences in hectares of the preliminary from the revised

estimates are also given in Table 23. Some estimates have changed by as much as 4000 hectares.

### 6.3.3 Precision of Estimates

The variance of the corn and soybean estimates can be calculated from the binomial assumptions. If  $\hat{p}_{c}$  represents the bias-corrected estimate of proportion corn in a county and  $\hat{p}_{s}$  represents the bias-corrected estimate of proportion soybeans in a county, then

$$v(\hat{p}_{c}) = \frac{\hat{p}_{c}(1-\hat{p}_{c})}{n-1}$$
 (1-f) and  $v(\hat{p}_{s}) = \frac{\hat{p}_{s}(1-\hat{p}_{s})}{n-1}$  (1-f),

where n is the number of pixels classified in the county and  $f^{=\frac{n}{N}} \text{ where N is the total number of pixels in the county.}$ 

The SRS sampling error is not known, but the sampling error of Landsat estimates is very small in comparison as it is very small absolutely. Sample standard deviations and coefficients of variation for Landsat estimates are presented in Tables 24 and 25. The standard deviations for the crop reporting districts and for the state were calculated considering the sample as stratified with each county considered a stratum. As in Kansas, the sampling error of the state, district, and county crop area estimates is very small.

Table 23. Differences of USDA/SRS preliminary 1974 estimates from revised estimates.

			·		
	Relative Difference (%)		i	rence n ares	
County	Corn	Soybeans	Corn	Soybeans	
Northwest District Benton Jasper Lake LaPorte Newton Porter Pulaski Starke White	-4.7 -5.0 -4.2 -0.1 -5.1 -1.0 .1.0 .0.4 -2.6	6.0 4.4 6.0 -3.8 -3.5 -3.1 4.7 9.8 4.0	-2145.7 -3238.9 -1133.6 -40.5 -2388.7 -283.4 404.9 121.5 -1376.5	2267.2 1457.5 931.2 -890.7 -850.2 -485.8 1417.0 1295.5 1578.9	
North Central District Carroll Cass Elkhart Fulton Kosciusko Marshall Miami St. Joseph Wabash	-0.9 -2.8 5.8 -1.0 -2.9 3.8 3.2 2.7 -0.9	2.5 6.4 -3.2 5.1 -4.0 -5.4 -6.2 -6.9 -7.6	-404.9 -1052.6 1619.4 -283.4 -1174.1 1295.5 1012.1 769.2 -283.4	566.8 1417.0 -445.3 931.2 -850.2 -1012.1 -1214.6 -1012.1 -1700.4	
Northeast District Adams Allen DeKalb Huntington LaGrange Noble Steuben Wells Whitley	2.4 -3.2 6.4 -1.0 -1.0 -0.9 6.0 2.1 -0.9	-8.1 -2.3 13.3 5.0 -7.6 -3.2 13.6 0.7 7.3	566.8 -1012.1 1093.1 -242.9 -202.4 -242.9 1012.1 566.8 -202.4	-2267.2 -890.7 2510.1 1417.0 -485.8 -404.9 850.2 242.9 1336.0	

Table 23. (continued)

		ative		rence
	Difference (%)		in Hectares	
County	Corn	Soybeans	Corn	Soybeans
West Central District		•		
Clay	-9.2	-15.4	-1740.9 1336.0	-2955.5
Fountain	4.5	-1.9		
Montgomery	-1.0	-7.7	-485.8	
Owen	17.1	6.9	1295.5	
Parke	4.4	5.4	1012.1	. 931.2
Putnam	-6.8 -1.0	0.6	-1619.4	121.5
Tippecanoe Vermillion	24.2	$-4.9 \\ 11.6$	-404.9 3279.4	
Vigo	6.2	0.7	1052.6	
Warren	6.4	0.6	1781.4	161.9
Wal I cli	0.4		1701.4	101.5
Central District		•		
Bartholomew	1.8	-1.5	445.3	-242.9
Boone	10.3	-4.0	3684.2	
Clinton	-0.9	-0.6	-404.9	
Decatur	2.5	0.7	890.7	121.5
Grant	0.6	-6.7	161.9	-1943.3
Hamilton	-1.0	-8.2	-283.4	
Hancock	-0.9	-0.7	-242.9	
Hendricks	2.7	-3.3	647.8	-769.2
Howard	-7.1	10.1	-1862.3	2186.2
Johnson	5.9 -4.6	-0.8 -13.4	1376.5 -1619.4	-121.5 -4048.6
Madison Marion	2.4	5.0	283.4	485.8
Morgan	-0.9	9.7	-161.9	1295.5
Rush	1.1	0.7	445.3	161.9
She1by	-4.8	0.7	-1902.8	161.9
Tiptón	5.3	8.0	1498.0	1781.4
	-			
East Central District	· ·			
Blackford	3.3	0.6	323.9	81.0
Delaware -	-0.9	-3.0	-242.9	-769.2
Fayette	-0.9	0.5	-121.5	40.5
Henry	-8.4	-2.7	-2469.6	-607.3
Jay	14.0	2.1	2388.7	607.3
Randolph	1.8	2.9	526.3	1052.6
Union Wayne	-0.9 2.5	12.4	-121.5	850.2
MMVUM	4.5	10.4	566.8	1740.9

Table 23. (continued).

	Diff	ative erence %)	Difference in Hectares	
County	Corn	Soybeans	Corn	Soybeans
Southwest District				
Daviess	3.1	-2.0	931.2	-283.4
Dubois	2.8	0.7	607.3	40.5
Gibson	-1.0	6.2	-404.9	1376.5
Greene	-2.2	-6.5	-404.9	-688.3
Knox	7.9	-1.3	3967.6	-283.4
Martin	-1.1	22.2	-81.0	404.9
Pike	-0.8	9.9	-121.5	890.7
Posey	4.1	4.6	1295.5	971.7
Spencer	-10.6	3.0	-1578.9	526.3
Sullivan	2.7	7.3	607.3	1336.0
Vanderburgh	8.2	-1.7	1093.1 -526.3	-202.4 -1700.4
Warrick	-3.8	-13.9	- 3-20 • 3	-1700.4
South Central District				1410
Brown	0.0	-33,3	0.0	-161.9
Crawford	0.0	8.3	0.0	81.0
Floyd	0.0	30.0	0.0	242.9
Harrison	-16.9	1.0	-1457.5	40.5
Jackson	4.0	12.2	971.7	1700.4
Lawrence	-0.9	24.1	-81.0 -40.5	850.2 202.4
Monroe	-1.1 -12.6	10.6 $1.2$	-1174.1	40.5
Orange	-6.6	1.4	-242.9	40.5
Perry	-23.6	0.7	-4048.6	40.5
Washington	-23.0	0.7	-4040.0	40.3
Southeast District			-242.9	40.5
Clark	-3.3	0.6	-890.7	-445.3
Dearborn	- <b>→ 18.2</b>	-15.7	-1295.5	445.3
Franklin	-7.7	5.9	-202.4	-890.7
Jefferson	-2.9	-11.4	1498.0	890.7
Jennings	11.5	8.4	-40.5	-121.5
Ohio	-1.8	-17.6	-121.5	1700.4
Ripley	-0.9	12.0	-40.5	1255.1
Scott	-0.8	25.0	-40.5	0.0
Switzerland	-1.4	0.0		

Table 24. Estimates of the standard deviations and coefficients of variation of Landsat estimates of corn in Indiana.

AREA ESTIMATE			PROPORTION ESTIMATE		
	<del></del>		FRUFURI	······································	<del></del>
	HECTARES	STANDARD DEVIATION	(%)	STANDARD DEVIATION	COEFFICIENT OF VARIATION
	(AH 000)	(HA)		(%)	(%)
,					
NORTHWEST DISTRIC	ST 53.6	195.97	50.6	0.1849	0.37
JASPER LAKE:	92.0 62.7	195.97 499.30 477.08	5631122223 471.4598.479.4879.4	0.1849 0.3435 0.3588 0.3230 0.4392 0.3885 0.4371 0.1616	0 • 37 0 • 54 0 • 76 0 • 79 0 • 74
LAPORTE NEWTON	64.7 63.0	510.05 40.05	41.1 59.2	0.3238	0.74 0.74
PORTER PULASKI	53.1 54.1	428.55 435.87	48.2	0.3885	0.81 0.81 0.92
STARKE WHITE	5964.77 643.01 643.4 6554.1 6554.1 6554.1	208.11	49.2		0.92 0.33
TOTAL	544.7	1239.02	50.4	0.1147	0.23
WEST CENTRAL DIS	TRICT		10.1	0.0470	1.30
CLAY FOUNTAIN	TRICT 0220 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0	233.04 471.05 471.05 592.06 4772.06 201.06 2	19.1 41.0 419.0 138.0 49.2 37.6	0.41188 0.318859 0.3158667 0.3159667 0.3159420 0.3159420 0.31696	1.070 0.078 1935 1935 1935 1935 1935 1935 1935
MÖNTGÖMERY OWEN PARKE	19.2	379.55	19.2 38.0	0.3805	1.98 1.33
PIITNAM	36.2 53.0	451.09 200.56	28.6 40.8	0.3567 0.1545	1.25 0.38
TIPPECANOE VERMILLION VIGO	33.5 21.7	342.09 342.62	49.2 20.2	0.5020 0.3186	1.02 1.58
WARREN				0.2056	0.33
TOTAL .	366.3	1211.80	34.2	0.1120	0.55
CENTRAL DISTRICT BARTHOLOMEW. BOONE	3.4	153.59	3.3	0.1474	4.47
BOONE CLINTON	3.4 5.6	191.23 127.60	5.1 2.3	0.1728 0.1210	3.39 5.26
DECATUR	37.3 31.0	397.20 177.28	38.9 28.4	0.4147 0.1625	0 · 6 7
GRANT HAMILTON HANCOCK	38.0 30.6	405.14 154.32	36.6 38.7	0.3899 0.1953	0.50
HENDRÍCKS HOWARD	48.2 39.5	432.73 361.32	52.0 52.0	0.4759	0.92
JOHNSON MADISON	36.7 46.7	191.20	39.8 14.5	0.1629	432050509914869 43510100011021
MADISON MARION MORGAN RUSH	15.3 38.8	9300842325050 • 2622132505050 5177• • • • • • • • • • • • • • • • • • •	3139467609855609 3508868409944619 3033453311354	0.14759359 0.14721429359 0.14689505939 0.1468950593957 0.447446073754	1.96 1.03 0.78
SHELBY TIPTON	30006256713807 30006256713807 30333433411353	421.18 341.94	51.0 49.9	0.3974 0.5056	0.78 1.01
TOTAL	472.2	1289.24	29.9	0.0816	0.27
EAST CENTRAL DIS	STRICT			-	
BLACKFORD DELAWARE	15 2	260.39 720.23	35.2 42.6	0.6018 0.6984	1.71 1.64
FÄŸETTE HENRY	13.3 23.8	401.80 354.60	23.8 23.0	0.7213 0.3421	3.03 1.49 0.56
JAY RANDOLPH	43.89 13.89 23.99 49.0	174.15 202.96	35.0.9 423.0.9 423.0.9 423.0.9 423.0.9 423.0.9	0.1741 0.1714 0.4406	0.41 1.55
UNION WAYNE	12.4 23.0	191.81 160.47	21.9	0.1529	6:76
TOTAL	211.5	1003.60	31.4	0.1492	0.48
STATE TOTAL	1594.7	2383.23	36.2	0.0541	0.15

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Table 25. Estimates of the standard deviations and coefficients of variation of Landsat estimates of soybeans in Indiana.

	AREA EST	IMATE	PROPORTION	ESTIMATE	
-	HECTARES	STANDARD DEVIATIO	) (%) )N	STANDARD DEVIATION	COEFFICIENT OF VARIATION
	(000 HA)	(HA)		(%)	(%)
NORTHWEST DISTRICT	, 20 3	154.39	19.2	0.1457	<b>9.</b> 76
JASPER LAKE	341944654 0000000000000000000000000000000000	154.39 373.92 355.62	19.2 15.4 16.6	0.1457 0.2570 0.2676 0.2676 0.28681 0.35332 0.3679 0.1305	06618791114.0.04
LAPORTE NEWTON PORTER	32.9 12.4	421.51 304.63 339.14 396.22 296.46 168.05	20.5 19.1 29.1 29.1 22.9 20.5	0.2676 0.2861	1.28 2.47 1.50
PULASKI	21.4 32.6	396.22	29.1	0.3532	î.Ží 1.61
STARKE WHITE					
TOTAL	209.0	974.36	19.3	0.0902	0.47
WEST CENTRAL DISTR	RIČŢ	245 02	. 27 6	0 2020	1 02°°
WEST CENTRAL DISTR CLAY FOUNTAIN MONTGOMERY	0646314562 1215913792 2212192	934 934 935 652 934 166 166 166 166 166 166 166 16	27.63 118.66 118.00 118.00 118.00 117.8	0.235 0.235 0.235 0.235 0.235 0.235 0.235 0.235 0.235 0.235 0.235 0.235 0.235	10.4405 10.555467 10.5767
MUNITOMERY OWEN PARKE PUTNAM	15.6	349.66 331.06	15.6. 8.0	0.3505 0.2833	2.25 3.54
PUTNĂM TIPPECANOE	21.1 23.4	372.32 156.78	16.7 18.0	0.2944 0.1208	1.76 0.67
VĒRMĪLLĪON VIGO WARREN	7.5 29.6	214.10 381.05	11.0 27.5	000077	0.67 2.86 1.29 1.11
		135.20 940.49	12.8 16.9	0.1418	0.52
TOTAL	180.7	740.47	10.3	U G U Q F T.	0.52
CENTRAL DISTRICT BARTHOLOMEW BOONE CLINTON	15.7	307.484	15.1	0.2955 0.3745	1.96
BOONE CLINTON	38.6 37.2	414.32	34.9. 35.3	0.3745 0.3857	1.07 1.09 1.93 0.74
DECATUR GRANT HAMILTON	7626138109173 5875191024811 1731222222222222222222222222222222222	8379355445759763 844005810886017 744005810886017 844313133331323	19999000000000000000000000000000000000	7572222 7572222 7814249 733146762319 73314649 74576329 74576319 745763 745763 745763 745763 745763 745763 745763 745763 745763	1.93 0.74
HANCOCK	21.8	141.64	27.6	0.3642 0.1792 0.3614	1.29 0.65 1.30
HENDŘÍČKS HOWARD JOHNSON	22.0	328.17	29.ú 42.8	0.4323 0.4519	34969632270 
MADISON MARION	.28.1 11.7	166.59 380.17	23.9 11.2	0.4519 0.1419 0.3650 0.2490	0.59 3.26
MORGAN RUSH	11.3 30.9	261.96 377.63	10.7	0.2490 0.3563	1.55 5.33
SHELBY TIPTON	30.9 33.4 23.3	377.63 391.37 324.87	29.2 31.5 34.4	0.3563 0.3693 0.4804	1.40
TOTAL	405.0	1320.84	25.6	0.0836	0.33
EAST CENTRAL DISTR	RICT				
BLACKFORD DELAWARE	11.6 33.0	241.19 679.42	26.7 32.0	0.5574 0.6588	2.09 2.06
FAYETTE HENRY	12.3 24.3	241.19 679.42 391.48 356.74 177.62	22.1 23.4	0.5574 0.6588 0.7022 0.34776 0.1631	3.18 1.47
JAY RANDOLPH	-603338800 10324338600	193.46	70149895 600099049 000099049	0.1634	.96873094 2031.455994
UNION WAYNE	10.5	148.90 113.77	9.5	0.3421 0.1084	
TOTAL	169.5	951.14	25.2	0.1414	0.56
STATE TOTAL	964.2	2118.91	21.9	0.0481	0.22

6.4 Regression Estimation for Corn and Soybeans in Areas Without Landsat Coverage

Landsat data was not analyzed due primarily to cloudiness for five districts in Indiana: North Central, Northeast, Southwest, South Central, and Southeast. Since estimates of the area and proportion of corn and soybeans in these counties were required, a prediction equation was developed for each crop using the 43 counties which had been classified with Landsat data. The Landsat estimates were written as a function of historical crop production in the two previous years, and acres in the county. These equations were then used to predict area and proportion estimates for corn and soybeans in the counties which did not have Landsat coverage.

To estimate the area of corn, the counties classified in Indiana were divided into three groups according to the USDA/SRS 1974 preliminary estimates of acreage of corn (Table 26). The rationale for dividing the counties into groups was to make the variances more homogeneous within groups. A prediction equation was formulated for each of the groups using the variables: acres in the county, the 1973 SRS revised estimate and the 1974 SRS preliminary estimates of acres of corn harvested in the county. The counties in which the area of corn was to be predicted fell into one of these three groups according to the same criterion; however, if the number of acres in the county or the 1973 or 1974 corn acreage estimate fell outside the

Table 26. Groupings used for regression estimation and the number of counties per group.

Group	Counties with Landsat data	Counties to be predicted	USDA/SRS 1974 preliminary acreage estimates
	For	Corn Estimat	ion
1	10	8	<50,000 acres
2	21	13	50-90,000 açres
3	.12	3	>90,000 acres
	For S	Soybean Estim	ation
1	12	12	<40,000 acres
2	14	14	40-60,000.acres
3	17	2	>60,000 acres

appropriate range, historical estimation was used. For 26 counties, historical estimates were used.

The prediction equations found are given as follows: for the first group,

$$\hat{y} = 3.98 + 0.01 x_1 - 0.46 x_2 + 0.81 x_3$$
 (R<sup>2</sup> = 0.31); for the second group,

$$\hat{y} = -19.33 + 0.10 x_1 + 1.22 x_2 - 0.67 x_3$$
 (R<sup>2</sup> = 0.30); for the third group,

$$\hat{y} = -69.36 + 0.17 x_1 - 1.80 x_2 + 2.33 x_3 (R^2 = 0.49)$$

where  $x_1$  is the number of thousands of acres in the county,  $x_2$  is the acreage of corn grown in a county in 1973 in thousands, and  $x_3$  is the acreage of corn grown in a county in 1974 in thousands. The "pseudo" Landsat estimate,  $\hat{y}$ , is given in thousands of hectares.

For soybean\_estimation, the counties were again divided into three groups, but this time the groupings were based upon the USDA/SRS 1974 preliminary soybean estimates (Table 26). For 21 counties, historical estimation was used. The prediction equations found are given as follows: for the first group,

$$\hat{y} = -2.08 + 0.02 x_1 + 0.25 x_2 + 0.17 x_3$$
 (R<sup>2</sup> = 0.32); for the second group,

$$\hat{y}$$
 = -6.71 + 0.04  $x_1$  + 0.33  $x_2$  ( $R^2$  = 0.20)  
(the variable  $x_3$  did not add sufficient information to enter the regression);

and for the third group,

$$\hat{y} = 29.87 - 0.03 x_1 - 0.19 x_2 + 0.27 x_3 (R^2 = 0.02);$$

where  $x_1$  is the number of thousands of acres in the county,  $x_2$  is the acreage of soybeans grown in a given county in 1973 in thousands, and  $x_3$  is the acreage of soybeans grown in a county in 1974 in thousands. The "pseudo" Landsat estimate,  $\hat{y}$ , is given in thousands of hectares. Estimates were then made using these six equations and historical data (Tables 27 and 28).

The estimates made by the prediction equations were generally not of as high an accuracy as the SRS estimates. Estimates of corn area and proportion were not significantly different from SRS estimates at the 25% level in the Northeast and Southeast Districts. In all other districts, however, and for soybean area and proportion estimates in all districts, the regression estimates were significantly different from those obtained by SRS.

Table 27. Regression estimates of area and proportion of corn in counties for which usable Landsat data was not available.

			Hectare (000)	es	P	roporti (%)	ron
County	*	SRS	Reg.	Diff.	SRS	Reg.	Diff.
North Central	Die	trict	,				
Carroll	H	44.2	43.4	- 0.8	45.6	44.8	- 0.8
Cass	H	38.7	37.0	- 1.7	36.0	34.4	- 1.6
E1khart	2	29.8	42.2	12.4	24.6	34.8	10.2
Fulton		31.5	36.6	5.1	33.1	38.5	5,4
Koscuisko	2 3 2 2 2 2		37.7	- 6.0	32.3		- 4.4
Marshall	2		44.1	8.6	30.9		7.4
Miami	2	33.2	36.8	3.6	33.7	37.4	3.7
St. Joseph	2	28.9	37.5	8.6	23.9		7.1
Wabash	2	33.4	43.7	10.3	30.6	40.1	9.5
Tota1		318.9	359.0	40.1.	31.9	35.9	4.0
Northeast Dist	tric	t	•				
Adams	2	23.0	23.4	0.4	25.7	26.2	0.5
Allen	Н	34.6	30.6	- 4.0	19.9	176	- 2.3
DeKa1b	1	18.6	22.6	4.0	19.7	23.9	4.2
Huntington	2	23.5	28.4	4.9	23.3	28.1	4.8
Lagrange	$\mathrm{H}^-$	<b>25.</b> 5	20.8	- 4.7	26.0	21.2	- 4.8
Noble	2	27.1	30.8	3.7	25.5	29.0	3.5
Steuben	1	17.5	23.1	5.6	21.8	28.8	7.0
Wells	2	25.9	27.5	.1.6	27.2	28.8	1.6
Whitley	H	22.6	21.3	- 1.3	26.0	24.5	- 1.5
Total		218.3	228.5	10.2 .	23.6	24.7	1.1

Table 27. (continued)

County	*	SRŚ	(000) Reg.	D • C C		-	
	6		6 -	Diff.	SRS	Reg.	Diff.
Southwest Distraction Daviess Dubois Gibson Greene Knox Martin Pike Posey Spencer Sullivan Vanderburgh Warrick	1 2 H 3 H 1 2 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1	30.7 23.2 43.1 21.2 52.0 8.7 15.1 33.2 18.8 23.9 13.8 14.9	39.4 22.3 42.0 18.3 86.7 7.5 19.5 38.5 17.4 39.2 20.2 19.1	8.7 - 0.9 - 1.1 - 2.9 34.7 - 1.2 4.4 5.3 - 1.4 15.3 6.4 4.2	14.9 38.8 9.7 17.4 31.0 18.3 20.2 22.1 14.7	35.3 19.9 32.5 12.9 64.7 8.4 22.5 35.9 17.0 33.1 32.4 18.9	7.8 - 0.8 - 0.8 - 2.0 25.9 - 1.3 5.1 4.9 - 1.3 12.9 10.3 4.2
Tota1		298.6	370.1	71.5	23.0	28.5	5.5
South Central  Brown Crawford Floyd Harrison Jackson Lawrence Monroe Orange Perry Washington Total	Dis H H H H H H H H	trict 1.2 2.1 1.4 8.3 27.0 9.7 3.7 10.1 4.4 18.1 86.0	1.2 1.9 1.3 7.2 25.3 9.2 3.6 8.2 3.4 13.1	0.0 - 0.2 - 0.1 - 1.1 - 1.7 - 0.5 - 0.1 - 1.9 - 1.0 - 5.0 -11.6	1.4 2.6 3.6 6.7 20.0 8.2 3.7 9.6 4.4 13.5	1.4 2.4 3.4 5.8 18.8 7.7 3.6 7.8 3.4 9.8	0.0 - 0.2 - 0.2 - 0.9 - 1.2 - 0.5 - 0.1 - 1.8 - 1.0 - 3.7 - 1.1

Table 27. (continued)

			Hectare (000)	es	. P	roport: (%)	ion
County	*	SRS	Reg.	Diff.	SRS	Reg.	Diff.
Southeast Dis	tric	t					
Clark	H	7.4	7.1	- 0.3	7.4	7.1	- 0.3
Dearborn	Η	5.2	4.0	- 1.2	6.6	5.0	- 1.6
Franklin	1	16.8	20.9	4.1	16.5	20.5	4.0
Jefferson	H	7.7	6.9	- 0.2	8.1	7.3	- 0.8
Jennings	1	12.5	21.6	9:1	12.8	22.1	9.3
Ohio	H	. 2.0	2.2	0.2	8.9	9.8	0.9
Ripley	Н	12.8	12.9	0.1	11.2	11.3	0.1
Scott	Н	4.9	4.7	- 0.2	9.8	9.4	- 0.4
Switzerlan	d H	3.1	2.8	- 0.3	5.4	4.9	- 0.5
Total		72.4	83.1	10.7	10.1	11.6	1.5

<sup>\*</sup>Method of Estimation: H-historical; 1, 2, and 3 refer to the groups defined\_in Table 26.

Table 28. Regression estimates of area and proportion of soybeans in counties for which usable Landsat data was not available.

North Central Distractions Carroll 2 Cass 2 Elkhart 1 Fulton 2 Koscuisko 2 Marshall 2 Miami 2 St. Joseph 1	SRS Reg. rict 21.7 24.8 20.5 23.5	Diff.	SRS`	Reg.	Diff.
Carroll 2 Cass 2 Elkhart 1 Fulton 2 Koscuisko 2 Marshall 2 Miami 2 St. Joseph 1	21.7 24.8	3.1	00.4		
Carroll 2 Cass 2 Elkhart 1 Fulton 2 Koscuisko 2 Marshall 2 Miami 2 St. Joseph 1	21.7 24.8	3.1	00 /		
Wabash 2	14.0 21.0 16.9 20.3 21.1 24.4 17.3 21.0 18.3 20.7 14.3 20.5 21.8 23.0	3.0 7.0 3.4 3.3 3.7 2.4 6.2 1.2	22.4 19.1 11.5 17.8 15.6 15.0 18.6 11.8 20.0	21.9 17.3	5.1
Total T	165.9~199.2	33.3	16.6	19.9	3.3
Northeast District Adams H Allen H DeKalb 2 Huntington 3 Lagrange H Noble 1 Steuben H Wells 3 Whitley 2	26.7 25.8 34.8 37.3 16.8 20.5 27.8 28.9 5.5 5.9 12.0 17.6 5.3 7.1 31.9 29.7 17.2 18.1	- 0.9 2.5 3.7 1.1 0.4 5.6 1.8 - 2.2 0.9	29.9 20.0 17.8 27.5 5.6 11.3 6.6 33.5 19.8	28.9 21.5 21.7 28.6 6.0 16.6 8.8 31.2 20.8	- 1.0 1.5 3.9 1.1 0.4 5.3 2.2 - 2.3 1.0

Table 28. (continued)

`			Hectare (000)	 ∋s	P	roporti (%)	on
County	*	SRS	Reg.	Diff.	SRS	Reg.	Diff.
Southwest Dist Daviess Dubois Gibson Greene Knox Martin Pike Posey Spencer Sullivan Vanderburgh	1 H 2 1 2 H 1 2 2	12.5 5.3 20.0 10.0 20.0 1.7 8.2 19.1 17.0 16.4 10.8 11.7	19.1 5.8 25.9 17.0 25.3 2.2 13.0 23.2 20.1 22.3 14.7 15.6	6.6 0.5 5.9 7.0 5.3 0.5 4.8 4.1 3.1 5.9 3.9	13.8 17.3	17.1 5.2 20.0 12.0 18.9 2.5 15.0 21.6 19.6 18.8 23.5 15.4	5.9 0.5 4.5 5.0 4.0 0.6 5.6 3.8 3.0 5.0 6.2 3.8
Total	•	152.7		51.5	11.8	15.7	3.9
South Central Brown Crawford Floyd Harrison Jackson Lawrence Monroe Orange Perry Washington	Dist	0.4 0.9 0.8 4.0 13.4 3.4 1.8 3.0 2.7	0.3 1.1 1.1 4.2 23.2 4.4 2.1 3.3 3.0 5.7	- 0.1 0.2 0.3 0.2 9.8 1.0 0.3 0.3	0.5 1.1 2.1 3.2 9.9 2.9 1.8 2.9 2.7	0.4 1.4 2.9 3.4 17.2 3.7 2.1 3.1 3.0 4.3	- 0.1 0.3 0.8 0.2 7.3 0.8 0.3 0.2 0.3
Total		35.8	48.4	12.6	3.5	4.8	1.3

Table 28. (continued)

			Hectar (000)	es	P	roport: (%)	ion
County	*	SRS	Reg.	Diff.	SRS	Reg.	Diff.
outheast Dis	trict						
Clark	Н	6.0	6.3	0.3	6.0	6.3	0.3
Dearborn	H	2.6	2.4	- 0.2	3.3	3.0	- 0.3
Franklin	1	6.8	11.8	5.0	6.7	11.6	4.9
Jefferson	H	7.4	6.9	- 0.5	7.8	7.3	- 0.5
Jennings	1	10.2	16.0	5.8	10.4	16.4	6.0
Ohio	H	0.6	0.6	0.0	2.7	2.7	0.0
Ripley	1	13.5	20.9	7 . 4	11.8	18.3	6.5
Scott	H	4.7	6.3	1.6	9.4	12.6	3.2
Switzerland		2.0	2.1	0.1	3.5	3.7	0.2
Total		53.8	73.3	19.5	7.5	10.2	2.7

<sup>\*</sup>Method of Estimation: H-historical; 1, 2, and 3 refer to the groups defined in Table 26.

## 7.0 SIGNIFICANT RESULTS AND CONCLUSIONS

The first sections of this report described the rationale and background of this research, defined the objectives and experimental approach, and presented the results. Many different phases of our investigation have produced results which we believe are significant in the development of remote sensing technology, particularly for crop surveys. New techniques for handling and analyzing multispectral scanner data were developed; crops were classified over larger areas than ever before. The results conclusively demonstrated the efficiency and applicability of computer-aided analysis techniques for estimating crop areas. The objectives and approach are briefly reviewed in this section; then the most significant results and conclusions are presented.

The overall objective of the investigation was to develop and test techniques utilizing Landsat MSS data to identify and determine the areal extent and distribution of crops over large geographic areas. The specific objectives were:

 Using Landsat data and computer-implemented pattern recognition, classify the major crops from regions encompassing different climates, soils and crops.

- Estimate crop areas for county and state size regions using identification data obtained from Landsat classifications.
- Evaluate the accuracy, precision and timeliness of crop estimates obtained from Landsat data.

The test areas and crops were Kansas, winter wheat, and Indiana, corn and soybeans. The major steps of the experimental approach used were:

- Use aerial photography as reference data for training the classifier.
- For counties without reference data, extend training statistics from adjacent counties having similar crops and soils.
- Classify and make area estimates from a systematic random sample of pixels distributed over an entire county.
- Adjust estimates for classification bias.
- Aggregate county estimates to district and state levels.
- Perform quantitative statistical evaluation of results using the area estimates made by USDA/SRS as a standard of comparison.

Landsat data acquired during March to June for the counties in seven crop districts of Kansas were classified; estimates of the area of wheat in each of the 80 counties were made and compared to the corresponding estimates made by the USDA/SRS. The correlation of the USDA/SRS county estimates of wheat area to the Landsat estimates was 0.80. The wheat proportion estimates of 49% of the Landsat county estimates were within ± 5% of the SRS estimates and 81% were within ± 10%. At the crop reporting district level there was a significant difference in

the Landsat and SRS estimates in only one of the seven districts. In that district the differences, although small, were all in one direction. For the state, the SRS estimate was 4,555,000 hectares compared to the Landsat estimate of 4,613,000 hectares, a relative difference of only 1.27%.

The coefficient of variation, a measure of the precision or sampling error, of the Landsat estimates was 0.06% compared to 4% for SRS estimates at the state level. The median coefficient of variation of the Landsat county estimates was 0.60%. At all levels, state, district, and county, the Landsat estimates were extremely precise compared to the corresponding USDA/SRS estimates.

Landsat data acquired during July, August, or September for 43 counties in four districts were classified for the Indiana portion of the study. The corn and soybean classification performances and area estimates were not as accurate as for wheat in Kansas. The correlation coefficients for Landsat and SRS county estimates of the areas of corn and soybeans were 0.67 and 0.56, respectively. The corn estimates were consistently high compared to SRS and the soybean estimates, although not biased as for corn, varied widely from SRS. There were also significant differences in the SRS and Landsat estimates at the district and state levels. As in Kansas, the Indiana Landsat estimates were very precise compared to the SRS estimates.

The generally lower level of performance in Indiana compared to Kansas is attributed to the greater number of crops and spectral classes to discriminate among; smaller, less homogeneous fields; less optimal timing of Landsat data acquisition; and less adequate reference or training data. A major difference between winter wheat identification in Kansas and corn and soybean identification in Indiana is that the crop calendar of winter wheat is different than most other cover types; whereas, corn and soybeans, both summer crops, have crop calendars similar to other cover types present, (i.e. are green at the same time) such as oats, hay, pasture, and trees. summary, the identification of corn and soybeans in Indiana is a much more difficult problem than winter wheat identification This fact was compounded by the lack of cloud-free in Kansas. Landsat data at critical times and inadequate reference data for optimal training of the classifier.

Results in both Kansas and Indiana could be improved by the following changes which can be recommended based on the results obtained in this investigation. In the area of stratification there are two recommendations: first, apply a more systematic, objective procedure for subdividing the scene into homogeneous areas, and second, use classification units smaller than a county when a county falls into more than one stratum. Two improvements in the area of data acquisition would be beneficial: first, coordinate aerial photography acquisition more

closely with the crop calendar and Landsat data acquisition; second, more timely delivery of Landsat data could be used to facilitate scheduling aerial photography missions. Finally, the computer costs for classification could be decreased by reducing the sampling fraction from 25% to either 6.25 or 4% without significantly affecting the accuracy or precision of the estimates.

The overall conclusions of the investigation are:

- Landsat MSS data was adequate to accurately identify wheat in Kansas; corn and soybean estimates for Indiana were less accurate.
- Computer-aided analysis techniques can be effectively used to extract crop identification information from Landsat data.
- Systematic sampling of entire counties made possible by computer classification methods resulted in very precise area estimates at county, district, and state levels.
- Training statistics can be successfully extended from one county to other counties having similar crops and soils if the training areas sampled the total variation of the area to be classified.

The synoptic view of Landsat provides the opportunity to obtain crop production information over very large areas, e.g. states and countries. By using computer processing techniques to classify pixels distributed over entire counties, it is also possible to make accurate and precise estimates for local areas, e.g. counties. These capabilities combining satellite, sensor, and computer make a worldwide, and at the same time, a local crop production information system possible. The procedures and

results of this investigation should be of particular interest to U.S. government "user" agencies including the Statistical Reporting Service, the Foreign Agricultural Service, and the Economic Research Service; international organizations such as the United Nations' Food and Agriculture Organization; and private firms such as grain exporting companies.

## 8.0 RECOMMENDATIONS

The experiences and results of this research with Landsat data have indicated a number of recommendations which should be considered in designing and implementing future satellite sensor/data processing systems. These are as follows:

Frequency of Data Collection: The 18 day collection sequence available with Landsat-2 proved to be inadequate for several phases of this study; although Landsat-1 data was used to fill in several gaps in the data, it was not readily available. An 8 to 10 day cycle would be much more satisfactory for crop surveys in the future. Because of frequent cloud cover problems, such an increase in frequency of coverage would assure a higher probability for collection of adequate quantity and quality of data during critical periods of the vegetative growing season. More frequent coverage than 18 days will also be required for monitoring crop conditions.

Wavelength Bands: Work with aircraft data and more recently with Skylab data has clearly shown the importance of the middle infrared and thermal infrared portions of the spectrum for crop identification. Because the Landsat scanner

does not obtain data in these wavelength regions, we believe that the classification accuracies achieved are not as high as would be possible. Addition of at least one wavelength band in the middle infrared portion of the spectrum (1.3-2.6µm) and at least one channel in the 8-13.5µm thermal infrared region in future satellite scanner systems will unquestionably allow significant improvements in many of the results obtained, and in the utility of this type of satellite data. Further, the narrower and more optimally placed visible and near infrared bands of the proposed thematic mapper sensor on Landsat D will be a substantial improvement [21].

Spatial Resolution: The 80 meter IFOV of the current Landsat MSS appears generally adequate for areas having relatively large fields, but it is definitely a limitation in working in areas with field sizes of 10 hectares or less. The 30 meter IFOV of the proposed thematic mapper sensor would be a major improvement in that it would greatly reduce the proportion of "mixed" field boundary pixels and facilitate locating field boundaries.

Time of Day: To maximize the signal/noise ratio and minimize the effect of shadows, Landsat overpasses near solar noon would be optimal. However, because of the normal mid-day build-up of cumulus clouds, it appears that the time of day utilized is nearly ideal, and a change in the time of data collection is not recommended for future systems.

Delays in Receipt of Data: Lengthy delays in receipt of data in either image or tape format precluded the possibility of a rapid analysis of the data and subsequent field checking. It is highly recommended that a system be developed to get an intermediate quality product into the hands of the investigators within 2-4 days after data collection. If cloud cover was minimal and overall data quality appeared promising, the investigator could then request tapes and final image product outputs and more intelligently schedule and utilize resources in collecting "ground truth."

Reference Data for Training: The importance of high quality, accurate reference data for training the classifier should be emphasized. A multistage sampling system combining coordinated ground observations; large scale aerial photography; small scale, high altitude photography; and Landsat data would be ideal and insure the greatest accuracy in the classification of Landsat data. However, in most instances one or two of the stages are sufficient and as additional knowledge and understanding of the multispectral responses of crops is gained, greater dependence can be placed on developing training statistics directly from the Landsat data. This approach is being utilized by LACIE for wheat and should be developed for other crops and regions.

Geometric Correction and Multitemporal Registration:
Although neither geometrically corrected or multitemporally

registered data were utilized in this investigation because of the current high cost of obtaining such data, both kinds of preprocessing should be made routinely available in order to increase the utility and performance of Landsat data. In this investigation geometrically corrected digital data would have considerably simplified the task of obtaining field and county coordinates. The ability to register multiple data sets is becoming increasingly important in that it allows the temporal dimension of the spectral measurements to be fully utilized, and will also allow satellite data to be effectively related to other maps. Future systems should provide a digital data format that has been geometrically corrected to a standard format base to facilitate data registration.

Data Analysis Techniques: Improvements in data analysis techniques are required to fully achieve the potential information content of multitemporal, spectral measurements acquired from space. The spatial dimension has been little used to date in computer-aided data analysis, although spatial characteristics are known to bear a great amount of information and are regularly used by photo interpreters. Still another aspect of satellite data analysis is the need to develop methods for effectively working over the large geographic areas for which Landsat data is obtained. The diversity of landscape patterns found over many areas of this size indicates that a logical first step in the classification of Landsat data is to stratify

or divide the scene into units which are internally similar. Such a stratification will be helpful in constructing sampling frames which minimize the variance among sample units and in determining the boundaries of areas over which training statistics can be satisfactorily extended.

Crop Yield Prediction: Although yield prediction or crop assessment was not an objective or within the scope of this investigation, there were indications as we analyzed the data that some of the observed variations in spectral response were due to factors which are related to yield such as amount of tillering, leaf area, and biomass. These relationships as well as the use of Landsat data to determine the extent and severity of catastrophic events such as drought should be explored in future studies.

In closing, we believe considerable progress toward an operational crop survey system was made as a result of this experiment. The results conclusively demonstrated the efficiency and applicability of computer-aided analysis techniques for estimating crop areas. Many of the techniques used in the investigation could be transferred to an operational system capable of producing accurate and precise crop area estimates for local areas such as counties, as well as for larger areas such as states or countries.

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## APPENDIX

Table Al. Summary of Landsat scenes and sources of training statistics used for classifications in Kansas.

County	Source of Training Statistics	Landsat <sup>.</sup> Scene	Date	
Northwest District				
Cheyenne Decatur Graham Norton Rawlins Sheridan Sherman Thomas	(local) Norton (local) (local) Cheyenne Trego (local) Sherman	2165-16450 2146-16392 2146-16395 2146-16392 2165-16450 2146-16395 2165-16453 2165-16453	June 17 June 17 June 17 July 6 June 17 June 7	1975 1975 1975 1975 1975 1975 1975 1975
North Central District				
Clay Cloud Jewell Mitchell Osborne Ottawa Phillips Republic Rooks Smith Washington	Ottawa (local) Smith Osborne (local) (local) Norton Cloud Graham (local) Cloud	2144-16282 2163-16334 2163-16340 2163-16340 2144-16282 2146-16392 2163-16334 2146-16395 2163-16334	July 4 July 4 July 4 July 4 July 4 June 15	1975 1975 1975 1975 1975 1975 1975
West Central District			•	
Gove Greeley Lane Logan Ness Scott Trego Wallace Wichita	Trego (local) Trego Wallace (local) Greeley (local) (local) Greeley	2146-16395 2165-16453 2146-16395 2165-16453 2146-16395 2165-16453 2165-16453 2165-16453	July 6 June 17 July 6 June 17 July 6 June 17 June 17 July 6	, 1975
Central District				
Barton Dickinson Ellis	(local) Saline Trego	2163-16340 2144-16282 2146-16395	June 15	, 1975 , 1975 , 1975

Table A1. (continued)

Central District (cont.)  Ellsworth Lincoln McPherson Marion Rice Rush Russell	Russell Russell (local) McPherson Barton Trego (local)	.2163-16340 2163-16340 2144-16282 2144-16282 2163-16340 2146-16395 2163-16340	July July June June July June July	4, 15, 15,	1975 1975 1975
Saline	(local)	2144-16282	June	15,	1975
Clark Finney Ford Grant Gray Hamilton Haskell Hodgeman Kearney Meade Morton Seward Stanton Stevens	Ford (local) (local) Hamilton Haskell (local) (local) (local) Stanton (local) (local) Hamilton	5032-16310 5032-16310 5032-16310 2147-16460 5032-16310 2147-16460 5032-16310 2146-16395 2147-16460 5032-16310 2147-16460 5032-16310 2147-16460 2147-16460	May May May June May June June May June May June May June May June	21, 21, 21, 18, 21, 18, 21, 18, 21, 18, 21,	1975 1975 1975 1975 1975 1975 1975 1975
South Central District			-		
Barber Barber Comanche Comanche Edwards Edwards Harper Harper Harvey Kingman Kingman Kiowa Kiowa Pawnee Pratt Pratt Reno Sedgwick	(local) (local) Pratt Pratt Pratt Pratt Sumner Sumner (local) (local) Pratt Pratt Pratt Pratt Pratt Stafford (local) (local) Stafford Sumner	2073-16342 2109-16341 2073-16342 2109-16341 2073-16342 2109-16341 2072-16284 2144-16284 2072-16284 2144-16284 2073-16342 2109-16341 2073-16342 2109-16341 2073-16342 2109-16341 2073-16342 2073-16342 2073-16342 2073-16342 2073-16342	April May April May April June April June April May April May April May April May April April May April	11, 5, 11, 5, 11, 4, 15, 4, 15, 5, 11, 5, 11, 5,	1975 1975 1975 1975

Table A1. (continued)

South Central District	(cont.)			
Sedgwick Stafford Sumner Sumner	Sumner (local) (local) (local)	2144-16284 2073-16342 2072-16284 2144-16284		15, 1975 5, 1975 4, 1975 15, 1975
Southeast District				
Allen Allen Bourbon Butler Chautauqua Cherokee Cowley Crawford Elk Greenwood Labette Montgomery Neosho Wilson Woodson	(local) (local) Allen Harvey Allen Allen Sumner Allen	2142-16171 2107-16225 2142-16171 2144-16284 2107-16225 2142-16171 2144-16284 2142-16171 2107-16225 2107-16225 2142-16171 2107-16225 2142-16171 2107-16225 2142-16171 2107-16225 2142-16171	June May June May June June May May June May June May June May	13, 1975 9, 1975 13, 1975 15, 1975 9, 1975 13, 1975 13, 1975 9, 1975 9, 1975 13, 1975 9, 1975 9, 1975 9, 1975 9, 1975 9, 1975 9, 1975

Table A2. Summary of Landsat scenes and sources of training statistics used for classification in Indiana.

•			
. County	Source of Training Statistics	Landsat Scene	Date
Northwest District			
Benton Jasper Lake LaPorte Newton Porter Pulaski Starke White	Newton (local) (local) (local) Lake (local) (local) (local)	2228-15522 2228-15515 2228-15515 2228-15515 2228-15515 2228-15515 2228-15515 2228-15515 2228-15522	Sept. 7, 1975 Sept. 7, 1975
Clay Fountain Montgomery Owen Parke Putnam Tippecanoe Vermillion Vigo Warren	Vigo (local) (local) (local) Owen (local) Parke (local) (local)	2173-15480 2228-15522 2209-15464 2173-15480 2228-15522 2173-15480 2228-15522 2228-15522 2173-15480 2228-15522	July 14, 1975 Sept. 7, 1975 Aug. 19, 1975 July 14, 1975 Sept. 7, 1975 July 14, 1975 Sept. 7, 1975 Sept. 7, 1975 July 14, 1975 Sept. 7, 1975 July 14, 1975 Sept. 7, 1975
Bartholomew Boone Clinton Decatur Grant Hamilton Hancock Hendricks Howard Johnson Madison Marion Morgan Rush Shelby Tipton	Decatur Hamilton Tipton (local) (local) (local) Hamilton (local) (local) Hamilton (local) (local) Hamilton Owen Shelby (local) (local)	2208-15412 2209-15464 2209-15464 2208-15412 2209-15464 2209-15464 2209-15464 2209-15464 2208-15412 2208-15405 2209-15464 2173-15480 2208-15412 2208-15412 2208-15412	Aug. 18, 1975 Aug. 19, 1975 Aug. 19, 1975 Aug. 18, 1975 Aug. 19, 1975 Aug. 19, 1975 Aug. 18, 1975 Aug. 19, 1975 Aug. 19, 1975 Aug. 19, 1975 Aug. 18, 1975 Aug. 18, 1975 Aug. 19, 1975 Aug. 18, 1975

Table A2. (continued)

# East Central District

B1ackford	Jay	2208-15405	Aug.	18, 1975
Delaware	Randolph	2208-15405	Aug.	18, 1975
Fayette	(local)	2208-15412	Aug.	18, 1975
Henry	Wayne	2208-15405	Aug.	18, 1975
Jay	(local)	2208-15405	Aug.	18, 1975
Randolph	(local)	2208-15405	Aug.	18, 1975
Union ^	Fayette	2208-15412	Aug.	18, 1975
Wayne	(local)	2208-15405	Aug.	18, 1975
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