A Covariance Estimator for Small Sample Size Classification Problems and Its Application to Feature Extraction

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Abstract

A key to successful classification of multivariate data is the defining of an accurate quantitative model of each class. This is especially the case when the dimensionality of the data is high, and the problem is exacerbated when the number of training samples is limited. For the commonly used quadratic maximum likelihood classifier the class mean vectors and covariance matrices are required and must be estimated from the available training samples. In high dimensional cases it has been found that feature extraction methods are especially useful, so as to transform the problem to a lower dimensional space without loss of information, however, here too, class statistics estimation error is significant. Finding a suitable regularized covariance estimator is a way to mitigate these estimation error effects. The main purpose of this work is to find an improved regularized covariance estimator of each class with the advantages of LOOC, and BLOOC. Besides, using the proposed covariance estimator to improve the linear feature extraction methods when the multivariate data is singular or nearly so is demonstrated. This work is specifically directed at analysis methods for hyperspectral remote sensing data.

1 The work described in this paper was sponsored in part by the National Imagery and Mapping Agency under grant NMA 201-01-C-0023

1 Introduction

As new sensor technology has emerged over the past few years, high dimensional multispectral data with hundreds of bands have become available. For example, the AVIRIS system[2] gathers image data in 210 spectral bands in the 0.4-2.4 μm range. Compared to the previous data of lower dimensionality (less than 20 bands), this hyperspectral data potentially provides a wealth of information. However, it also raises the need for more specific attention to the data analysis procedure if this potential is to be fully realized.

Among the ways to approach hyperspectral data analysis, a useful processing model that has evolved in the last several years [1] is shown schematically in Figure 1. Given the availability of data (box 1), the process begins by the analyst specifying what classes are desired, usually by labeling training samples for each class (box 2). New elements that have proven important in the case of high dimensional data are those indicated by boxes in the diagram marked 3 and 4. These are the focus of this work and will be discussed in more detail shortly, however the reason for their importance in this context is as follows. Classification techniques in pattern recognition typically assume that there are enough training samples available to obtain reasonably accurate class descriptions in quantitative form. Unfortunately, the number of training samples required to train a classifier for high dimensional data is much greater than that required for conventional data, and gathering these training samples can be difficult and expensive. Therefore, the assumption that enough training samples are available to accurately estimate the

2 Airborne Visible and Infrared Imaging Spectrometer system, built and operated by the NASA Jet Propulsion Center.

class quantitative description is frequently not satisfied for high dimensional data. There are many types of classification algorithms used on such data. Perhaps the most common is the quadratic maximum likelihood algorithm. For such a quadratic classifier, user classes must be modeled by a set of subclasses, and the mean vector and covariance matrix of each subclass are the parameters that must be estimated from training samples. Usually the ML estimator is used. When the dimensionality of data exceeds the number of training samples, the ML covariance estimate is singular and cannot be used, however even in cases where the number of training samples is only two or three times the number of dimensions, estimation error can be a significant problem.

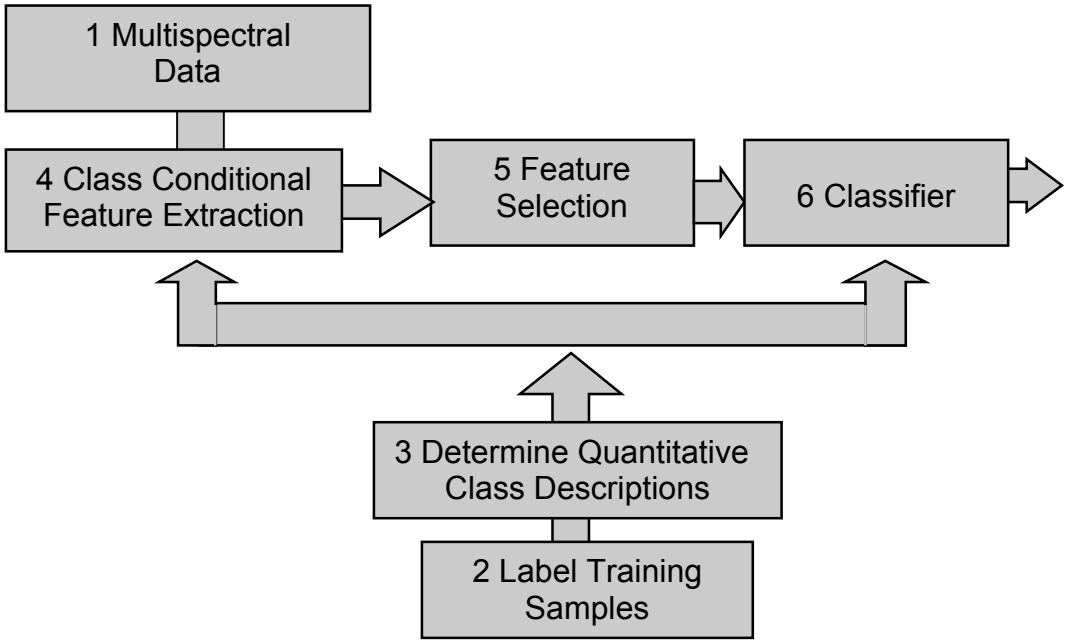

Figure 1. A schematic diagram for a hyperspectral data analysis procedure.

There are several ways to overcome this difficulty. In [2], these techniques are categorized into three groups:

a. Dimensionality reduction by feature extraction or feature selection.

b. Regularization of sample covariance matrix (e.g. [3], [4], [5], [6], [7]).

c. Structurization of a true covariance matrix described by a small number of parameters [2].

The purposes of this study are to find an improved regularized covariance estimator of each class that is invertible and with the advantages of LOOC [5], [6] and BLOOC [7] (box 3 of the figure), and show that linear feature extraction procedures (box 4) can be improved by using the proposed regularized covariance estimator.

2 Previous methods for regularization

Several methods for regularization have appeared in the literature. Regularized Discriminant Analysis (RDA) [3] is a two-dimensional optimization over four-way mixtures of the sample covariance, common covariance, the identity matrix times the average diagonal element of the common covariance, and the identity matrix times the average diagonal element of the sample covariance. The pair of mixing parameters is selected by cross-validating on the total number of misclassifications based on available training samples. Even though this procedure has the benefit of directly relating to the classification accuracy, it is computationally expensive, and the same mixing parameters must be used for all classes.

Leave-One-Out Covariance Estimator (LOOC; [5],[6]) uses the following mixture scheme.

$$\hat{\Sigma}_i(\lambda_i) = \begin{cases} (1-\lambda_i)\,diag(S_i) + \lambda_i S_i & 0 \le \lambda_i \le 1 \\ (2-\lambda_i)S_i + (\lambda_i - 1)S & 1 \le \lambda_i \le 2 \\ (3-\lambda_i)S + (\lambda_i - 2)\,diag(S) & 2 \le \lambda_i \le 3 \end{cases} \tag{1}$$

The mean of class *i*, without sample *k*, is $m_{i/k} = \frac{1}{N_i - 1}\sum_{\substack{j=1 \\ j \neq k}}^{N_i} x_{i,j}$, where the notation */k* indicates the quantity is computed without sample *k*. The sample covariance of class *i*, without sample *k*, is

$$\Sigma_{i/k} = \frac{1}{N_i - 2}\sum_{\substack{j=1 \\ j \neq k}}^{N_i} (x_{i,j} - m_{i/k})(x_{i,j} - m_{i/k})^T \tag{2}$$

and the common covariance, without sample *k* from class *i*, is

$$S_{i/k} = \left(\frac{1}{L}\sum_{\substack{j=1 \\ j \neq i}}^{L} \Sigma_j\right) + \frac{1}{L}\Sigma_{i/k} \tag{3}$$

The proposed estimate for class i, without sample k, can then be computed as follows:

$$C_{i/k}(\lambda_i) = \begin{cases} (1-\lambda_i)\,diag(\Sigma_{i/k}) + \lambda_i \Sigma_{i/k} & 0 \le \lambda_i \le 1 \\ (2-\lambda_i)\Sigma_{i/k} + (\lambda_i - 1)S_{i/k} & 1 < \lambda_i \le 2 \\ (3-\lambda_i)S_{i/k} + (\lambda_i - 2)\,diag(S_{i/k}) & 2 < \lambda_i \le 3 \end{cases} \tag{4}$$

The mixing parameter λ_i is determined by maximizing the average leave-one-out log likelihood of each class:

$$LOOL_i = \frac{1}{N_i}\sum_{k=1}^{N_i} \ln[f(x_k \mid m_{i/k}, C_{i/k}(\lambda_i)] \tag{5}$$

As aforementioned, in the process of selecting the mixing parameter values by maximizing the leave-one-out average log likelihood, the class covariance estimates can be determined independently, and then each class can have a mixing parameter that is optimal in terms of available training samples. Overall, classes with more training samples only need a small amount of bias, while classes with very few training samples need more bias.

BLOOC [7] is a modification of LOOC. LOOC was found to work well for well-trained classifiers, however, it is sensitive to outliers. In practice outliers frequently occur in cases where the class list is not exhaustive, such that the missing classes constitute outliers to the defined classes. Thus the following scheme was devised.

$$\hat{\Sigma}_i(\alpha_i) = \begin{cases} (1-\alpha_i)\dfrac{tr(S_i)}{p} I + \alpha_i S_i & 0 \le \alpha_i \le 1 \\ (2-\alpha_i)S_i + (\alpha_i - 1)S_p^*(t) & 1 \le \alpha_i < 2 \\ (3-\alpha_i)S + (\alpha_i - 2)\dfrac{tr(S)}{p} I & 2 < \alpha_i \le 3 \end{cases}$$

where t can be expressed as the function of α_i, $t = \dfrac{(\alpha_i - 1) f_i - \alpha_i (p+1)}{2 - \alpha_i}$, where p is the dimensionality and $f_i = N_i - 1$, which represents the degree of freedom in Wishart distributions, and the pooled covariance matrices are determined under a Bayesian context and can be represented as:

$$S_p^*(t) = \left[\sum_{i=1}^{L} \frac{f_i}{f_i + t - p - 1} \right]^{-1} \sum_{i=1}^{L} \frac{f_i S_i}{f_i + t - p - 1} \qquad (6)$$

The first difference between LOOC and BLOOC is that LOOC uses the diagonal entries of covariance matrices but BLOOC, like RDA, uses the trace of covariance matrices. Second, in LOOC, the maximum likelihood common covariance estimator is used, but, in BLOOC, the maximum a posterior common covariance estimator (S_p^*) is added. From [4], S_p^* tends to mitigate the outlier problem, and so does BLOOC. The criterion function of BLOOC is the same as that of LOOC.

A comparison of the performances of RDA, LOOC, and BLOOC is given in [8]. This comparison shows that LOOC performance is better than RDA in most situations, and BLOOC performs even better in special situations. In addition, computation time is decreasing in the order RDA, BLOOC, and LOOC. According to both accuracy and computation, LOOC is a better choice than the others. However, BLOOC has an advantage of being more resistant to outliers in the training set.

3 Mixed Leave-One-Out Covariance (Mixed-LOOC) Estimators

3.1 Mixed-LOOC1

LOOC and BLOOC are the linear combination of two of the three matrices, and in some situations, the difference between LOOC and BLOOC is in those matrices used to formulate the regularized covariance estimator. Only using some of the six matrices will not perform well in all situations. The basic idea of Mixed-LOOC is to use all six matrices to gain the advantages of both LOOC and BLOOC. Hence the first proposed regularized covariance estimator, Mixed-LOOC1, is

$$\hat{\Sigma}_i(a_i,b_i,c_i,d_i,e_i,f_i) = a_i\frac{tr(S_i)}{p}I + b_i diag(S_i) + c_i S_i + d_i\frac{tr(S)}{p}I + e_i diag(S) + f_i S$$

where $a_i + b_i + c_i + d_i + e_i + f_i = 1$ and $i = 1,2,...,L$

L : number of classes (7)

p : number of dimensions

S_i : covariance matrix of class i

S : common covariance matrix (pooled)

The mixture parameters are determined by maximizing the average leave-one-out log likelihood of each class:

$$LOOL_i = \frac{1}{N_i}\sum_{k=1}^{N_i}\ln[f(x_k\,|m_{i/k},\hat{\Sigma}_{i/k}(\alpha_i))], \quad \text{where } \alpha_i = (a_i,b_i,c_i,d_i,e_i,f_i) \tag{8}$$

3.2 Mixed-LOOC2

Since using Mixed-LOOC1 is computationally intensive, finding a more simplified estimator will be more practical. It is shown in [8] that given two known matrices, the ML (not Leave-One-Out) estimate of mixture parameters in LOOC and BLOOC are at the end points (α_i =0, 1, 2, or 3). Figures 2, and 3 illustrate the relationship between LOOL and the mixture parameter, α_i. The first figure is generated from a simulated data set; Figure 3 is based on a real data set [8]. The detailed information about simulated and real data set is in the experiment design section (section 4). In the case of Figure 2, the sample size is greater than the dimensionality. For Figure 3, the sample size is less than the dimensionality. Figure 3 shows that when the ML covariance estimator is singular, the optimal choice of LOOC parameter under LOOL criteria is around the boundary points.

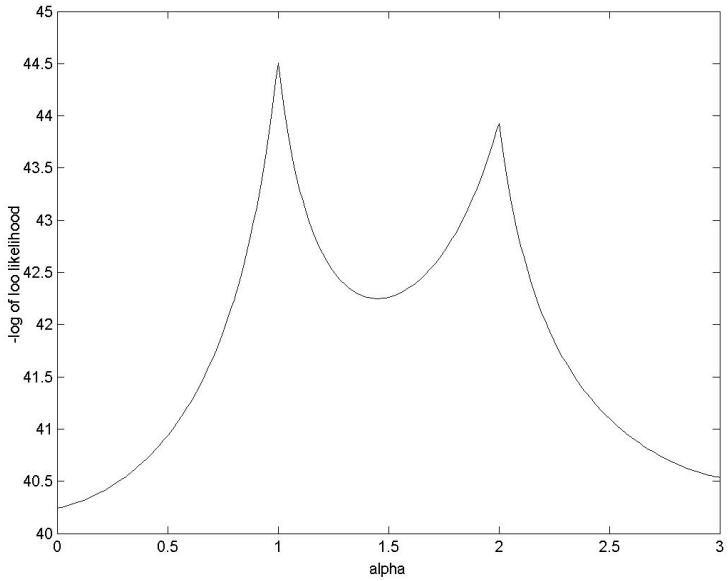

Figure 2 -LOOL of class1 in experiment 10 (p=10) and

the minimum of -LOOL occurs at alpha= 0

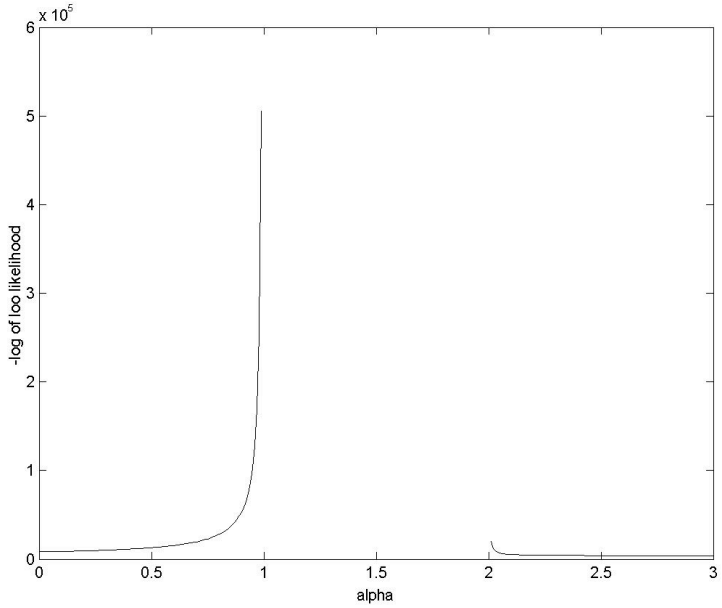

Figure 3 -LOOL of class 6 in the DC data set (p=191) and

the minimum of LOOL occurs at alpha= 2.99

Since a closed form solution for the parameter λ_i under the LOOL criteria is not available, and based on the above observations, one of the six support matrices is chosen to be the covariance estimator to reduce the computation time. The Mixed-LOOC2 is proposed as the following form:

$$\hat{\Sigma}_i(\alpha_i) = \alpha_i A + (1 - \alpha_i)B \tag{9}$$

where $A = \frac{tr(S_i)}{p} I,\ diag(S_i),\ S_i,\ \frac{tr(S)}{p} I,\ diag(S),$ or S, $B = S_i$, or $diag(S)$ and α_i is close to 1. $B = S_i$, or $diag(S)$ is chosen because if a class sample size is large, S_i will be a better choice. If total training sample size is less than the dimensionality, then the common (pooled) covariance S is singular but has much less estimation error than S_i. For reducing estimation error and avoiding singularity, $diag(S)$ will be a good choice. The selection criteria is the log leave-one-out likelihood function:

$$LOOL_i = \frac{1}{N_i} \sum_{k=1}^{N_i} \ln[f(x_k \mid m_{i/k}, \hat{\Sigma}_{i/k}(\alpha_i))] \tag{10}$$

The algorithm to decide the Mixed-LOOC2 of each class is to compute LOOL of the 12 covariance estimator combinations, then choose the maximal one. This method needs less computation time than the LOOC proposed in [5].

4 Experiment Design for Comparing LOOC, Mixed-LOOC1, and Mixed-LOOC2

In the following experiments, the grid method is used to estimate the mixture parameters of LOOC and Mixed-LOOC1. The range of the parameter λ in LOOC is from 0 to 3 and the grids

are γ = [0, 0.25, 0.5, …, 2.75, 3]. There are six parameters in Mixed-LOOC1 and the ranges of them are from 0 to 1. The grids of Mixed-LOOC1 are [0, 0.25, 0.5, 0.75, 1]. For Mixed-LOOC2, the parameter β is set to 0.05. In the simulation experiments, performances of all three covariance estimators are compared. Based on computational consideration, only the performances of LOOC and Mixed-LOOC2 are compared for the real data experiments.

Experiments 1 to 12 are based on simulated data sets. Experiments 1 to 6 and experiments 7 to 12 are generated from the same normal distributions respectively. The mean vectors and covariance matrices of experiments 1 to 6 (and 7 to 12) are the same as those six experiments in [3]. See also Appendix B of [8]. The only difference between these two sets of experiments is that experiment 1 to 6 are with equal training sample sizes in each class but experiments 7 to 12 are with different sample sizes in each class. Training and testing sample sizes of these experiments are in Table 1. There are three different dimensionalities, p=10, 30, 60, in every experiment. At each situation, 10 random training and testing data sets are generated for computing the accuracies of algorithms, and the standard deviations of the accuracies.

Table 1 The Design of Sample Size

	Experiments 1 ~ 6			Experiments 7 ~ 12		
Sample Size	Class 1	Class 2	Class 3	Class 1	Class 2	Class 3
Training	10	10	10	30	10	5
Testing	200	200	200	600	200	100

There are four different real data sets, the Cuprite site in western Nevada, an area of geologic interest; Jasper Ridge in California, an ecological site; Indian Pine in NW Indiana, an agricultural/forestry site; and the Washington, DC Mall, an urban site; in experiment 13 to 16 respectively. All real data sets have 191 bands. There are 8, 6, 6, and 7 classes used in the

Cuprite Site, Jasper Ridge Site, Indian Pine Site, and DC Mall, respectively. There are 20 training samples in each class. At each experiment, 10 training and testing data sets are selected for computing the accuracies of algorithms, and the standard deviations of the accuracies.

5 Experiment Results

The simulated data results are displayed in Table 2(a), 2(b), 2(c). The real data results are displayed in Table 2(d). They show the following.

1. In Table 2(a), (b), (c), the shadowed parts indicate that the differences of performances of LOOC and Mixed-LOOC2 are larger than the standard deviation of Mixed-LOOC2. If the difference is smaller than the standard deviation, we assume that the performances of these methods have no significant difference.

2. All the experiments with significant differences (the shaded parts) indicate that Mixed-LOOC outperformed LOOC.

3. The results of shaded parts show that the differences between Mixed-LOOC and LOOC increase as the number of dimensions increases.

4. When the training sample sizes of the classes are unbalanced, Mixed-LOOC performed better than LOOC in more situations.

5. Significant differences most often occurred in experiments 2, 7, and 8. Those are the situations in which BLOOC has better performances than LOOC. Since the Mixed-LOOCs are the union version of LOOC and BLOOC, based on these findings, we conclude that the Mixed-LOOCs have advantages over LOOC and BLOOC and can avoid their disadvantages.

6. In most of the experiments, the standard deviations of the Mixed-LOOCs are less than those of LOOC. This suggests Mixed-LOOCs are more stable than LOOC.

7. The results of experiment 13 (Cuprite Site) shows that Mixed-LOOC2 outperforms LOOC very much. The results of experiment 13 and 14 (Jasper Ridge Site) shows that the performances of Mixed-LOOC2 is more stable than those of LOOC

8. The computation time decreases in the order Mixed-LOOC1, LOOC, and Mixed-LOOC2.

Table 2(a) The Accuracy of Simulated Data Sets (p=10)

Experiment	LOOC	Mixed-LOOC1	Mixed-LOOC2
1	0.8630 (0.0425)	0.8632 (0.0243)	0.8602 (0.0466)
2	0.7753 (0.0481)	0.8373 (0.0180)	0.8450 (0.0224)
3	0.8948 (0.0241)	0.8915 (0.0251)	0.8992 (0.0265)
4	0.8875 (0.0309)	0.8893 (0.0263)	0.8837 (0.0386)
5	0.9860 (0.0283)	0.9822 (0.0361)	0.9858 (0.0282)
6	0.9885 (0.0033)	0.9833 (0.0085)	0.9885 (0.0036)
7	0.8500 (0.0286)	0.8622 (0.0252)	0.8641 (0.0249)
8	0.8433 (0.0410)	0.8750 (0.0289)	0.8792 (0.0250)
9	0.9021 (0.0230)	0.9041 (0.0183)	0.9041 (0.0203)
10	0.8928 (0.0247)	0.8948 (0.0204)	0.8940 (0.0245)
11	0.9883 (0.0064)	0.9920 (0.0041)	0.9872 (0.0065)
12	0.9841 (0.0076)	0.9830 (0.0075)	0.9827 (0.0116)

Table 2(b) The Accuracy of Simulated Data Sets (p=30)

Experiment	LOOC	Mixed-LOOC1	Mixed-LOOC2
1	0.8317 (0.0227)	0.8285 (0.0196)	0.8267 (0.0213)
2	0.7263 (0.0510)	0.8700 (0.0205)	0.8813 (0.0204)
3	0.8162 (0.0220)	0.8142 (0.0223)	0.8152 (0.0237)
4	0.7978 (0.0619)	0.7955 (0.0609)	0.7972 (0.0612)
5	0.9993 (0.0014)	0.9975 (0.0037)	0.9993 (0.0014)
6	0.9990 (0.0021)	0.9945 (0.0087)	0.9992 (0.0016)
7	0.8239 (0.0345)	0.8469 (0.0154)	0.8504 (0.0171)
8	0.8718 (0.0311)	0.9210 (0.0130)	0.9189 (0.0118)
9	0.8228 (0.0274)	0.8343 (0.0206)	0.8241 (0.0268)
10	0.8326 (0.0162)	0.8370 (0.0186)	0.8313 (0.0156)
11	0.9976 (0.0021)	0.9994 (0.0008)	0.9984 (0.0018)
12	0.9953 (0.0059)	0.9991 (0.0007)	0.9978 (0.0047)

Table 2(c) The Accuracy of Simulated Data Sets (p=60)

Experiment	LOOC	Mixed-LOOC1	Mixed-LOOC2
1	0.7378 (0.0540)	0.7607 (0.0259)	0.7605 (0.0287)
2	0.6578 (0.0631)	0.8792 (0.0213)	0.8882 (0.0175)
3	0.7632 (0.0265)	0.7615 (0.0235)	0.7583 (0.0281)
4	0.7483 (0.0324)	0.7473 (0.0308)	0.7435 (0.0288)
5	1.0000 (0.0000)	0.9998 (0.0005)	1.0000 (0.0000)
6	1.0000 (0.0000)	1.0000 (0.0000)	1.0000 (0.0000)
7	0.7820 (0.0327)	0.8098 (0.0229)	0.8120 (0.0192)
8	0.8876 (0.0219)	0.9401 (0.0075)	0.9400 (0.0073)
9	0.7947 (0.0216)	0.8024 (0.0150)	0.7958 (0.0203)
10	0.7802 (0.0302)	0.7932 (0.0277)	0.7837 (0.0275)
11	0.9988 (0.0021)	0.9997 (0.0011)	0.9997 (0.0011)
12	1.0000 (0.0000)	1.0000 (0.0000)	1.0000 (0.0000)

Table 2(d) The Accuracy of Real Data Sets (p=191)

Real Data Set	LOOC	Mixed-LOOC2
Cuprite	0.7743 (0.1372)	0.9524 (0.0117)
Jasper Ridge	0.9864 (0.0042)	0.9849 (0.0019)
Indian Pine	0.7612 (0.0127)	0.7625 (0.0144)
DC Mall	0.7831 (0.0455)	0.7858 (0.0431)

6. Discriminate Analysis Feature Extraction Based on Mixed-LOOC

The purpose of DAFE (discriminate analysis feature extraction) is to find a transformation matrix A such that the class separability of transformed data $Y=A^TX$ is maximized. Usually within-class, between-class, and mixture scatter matrices are used to formulate the criteria of class separability. A within-class scatter matrix is expressed by [9]:

$$S_w = \sum_{i=1}^{L} P_i E\{(X - m_i)(X - m_i)^T \mid \Omega_i\} = \sum_{i=1}^{L} P_i \Sigma_i \tag{11}$$

where L is the number of classes and P_i and m_i are the prior probability and mean vector of the class i, respectively.

A between-class scatter matrix is expressed as

$$S_b = \sum_{i=1}^{L} P_i (m_i - m_0)(m_i - m_0)^T = \sum_{i=1}^{L-1} \sum_{j=i+1}^{L} P_i P_j (m_i - m_j)(m_i - m_j)^T \quad (12)$$

where m_0 represents the expected vector of the mixture distribution and is given by

$$m_0 = E\{X\} = \sum_{i=1}^{L} P_i m_i \quad (13)$$

Let $Y = A^T X$, then we have

$$S_{wY} = A^T S_{wX} A \text{ and } S_{bY} = A^T S_{bX} A \quad (14)$$

The optimal features are determined by optimizing the criterion given by

$$J_1 = tr(S_{wY}^{-1} S_{bY}) \quad (15)$$

The optimum A must satisfy

$$(S_{wX}^{-1} S_{bX}) A = A(S_{wY}^{-1} S_{bY}) \quad (16)$$

This is a generalized eigenvalue problem [10] and usually can be solved by the QZ algorithm. But if the covariance is singular, the result will have a poor and unstable performance on classification. In this section, the ML covariance estimate will be replaced by Mixed-LOOC when it is singular. Then the problem will become a simple eigenvalue problem.

For convenience, denote DAFE based on ML estimators as DAFE, DAFE based on Mixed-LOOC2 as DAFE-Mix2, Gaussian classifier based on ML estimators as GC, and Gaussian classifier based on Mixed-LOOC2 estimators as GC-Mix2. Experiments 17 to 19 are for determining the performances of DAFE-Mix2. The classification process in experiment 17 is to

use DAFE then GC, in experiment 18 use DAFE-Mix2 then GC, and in experiment 19 use DAFE-Mix2 then GC-Mix2. The class sample sizes of experiment 18 and 19 are the same as those of experiments 13 to 16 (N_i=20). Since using those sample sizes in DAFE will cause very poor results, we increase the sample size of each class in Cuprite, Jasper Ridge, Indian Pine, and DC Mall data sets up to 40. The number of features extracted from the original space is set to the number of classes minus 1. The results of those experiments are shown in Table 3 and Figure 4.

Table 3 The Mean Accuracies and Standard Deviations of Experiments

Real Data Set	Exp17 DAFE+GC	Exp18 DAFE-Mix2+GC	Exp19 DAFE-Mix2+GC-Mix2
Cuprite	0.8943 (0.0205)	0.9474 (0.0194)	0.9627 (0.0196)
Jasper Ridge	0.9127 (0.0243)	0.9782 (0.0120)	0.9876 (0.0036)
Indian Pine	0.5727 (0.0156)	0.7547 (0.0316)	0.7562 (0.0191)
DC Mall	0.7392 (0.0530)	0.8691 (0.0282)	0.8600 (0.0345)

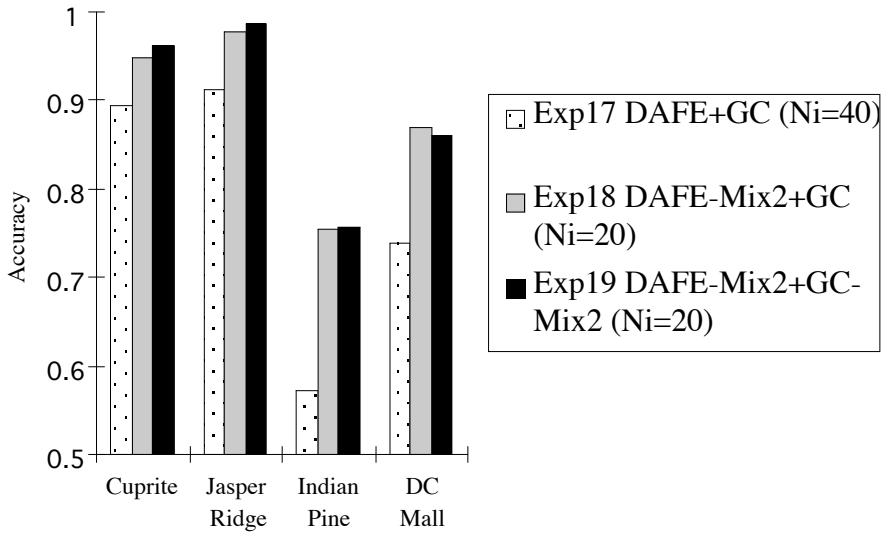

Figure 4 The Mean Accuracies of Three Classification Procedures

From above results we find the following.

1. Using DAFE-Mix2 provides higher accuracy and, in most cases, smaller standard deviation than using only DAFE.

2. Comparing Table 2(d) and Table 3, we find that in all data sets except the DC Mall sets, using DAFE-Mix2 then GC or GC-Mix2 have similar results with only using GC-Mix2. But the results for DC Mall show that using DAFE-Mix2 then GC or GC-Mix2 gave a significant improvement.

3. From Table 3 and Figure 4, DAFE-Mix2 -GC-Mix2 looks like the best choice.

7 Concluding Comments

The singularity or near-singularity problem often occurs in the case of high dimensional classification. From the above discussion, we know that finding a suitable regularized covariance estimator is a way to mitigate this problem. Further, Mixed-LOOC2 has advantages over LOOC and BLOOC and needs less computation than those two. The problems of class statistics estimation error resulting from training sets of finite size grows rapidly with dimensionality, thus making it desirable to use no larger feature space dimensionality than necessary for the problem at hand, and therefore the importance of an effective, case-specific feature extraction procedure. Usually DAFE cannot be used when the training sample size is less than dimensionality. The new procedure, DAFE-Mix2, overcomes this shortcoming, and can provide higher accuracy when the sample size is limited.

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